# CALCULATING ITEM DISCRIMINATION VALUES USING SAMPLES OF EXAMINEE SCORES AROUND REAL AND ANTICIPATED CUT SCORES: EFFECTS ON ITEM DISCRIMINATION, ITEM SELECTION, EXAMINATION RELIABILITY, AND CLASSIFICATION DECISION CONSISTENCY 

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#### Abstract

Darin S. Earnest: Calculating item discrimination values using samples of examinee scores around real and anticipated cut scores: Effects on item discrimination, item selection, examination reliability, and classification decision consistency (Under the direction of Gregory J. Cizek)


This study examined the degree to which limiting the calculation of item discrimination values to groups of examinee scores near real and anticipated cut scores affected item discrimination, item selection, examination reliability, and classification decision consistency. Three examinations used to credential individuals in health-related professions were used to answer the research questions. To replicate as closely as possible the context in which many credentialing examinations are developed, each of the examinations consisted of small samples of examinees and were analyzed using classical test theory procedures.

Item discrimination values, as expressed by the point-biserial statistic, were calculated for each examination item. Restricted item discrimination values were then calculated for each item using subsets of examinee scores. The restricted values were based on scores within $0.50 S D, 0.75 S D$, and $1.00 S D$ of five unique cut score locations. Differences between unrestricted and restricted item discrimination values were measured. Two 50-item test variants for each examination were created to evaluate the effect restricted item discrimination values had on item selection, examination reliability, and classification decision consistency. Form A variants included the 50 most discriminating items using
unrestricted discrimination values. Form B variants included the 50 most discriminating items using restricted discrimination values.

The results of the study indicated that (a) item discrimination values were lower when their calculation was limited to groups of scores near cut scores; (b) using restricted item discrimination values as the criterion by which items were selected for test variants resulted in the selection of items that were different than those selected when unrestricted values were used as the selection criterion; (c) differences in examination reliability between test variants were found to be statistically significant, with scores of variants based on restricted item discrimination values producing lower estimates; and (d) test variants based on restricted item discrimination values produced slightly lower observed classification decision consistency estimates than variants based on unrestricted item discrimination values. The results of the study were tied to several aspects of the test development process for credentialing examinations, including issues related to sample size, cut score location, and examination validity.

To my family.

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## CHAPTER 1

## INTRODUCTION

Examinations, and the roles they play in a variety of fields, have been the source of much debate in recent years. In the educational setting, for example, legislation like the 2001 No Child Left Behind Act (NCLB, 2002) shifted significant attention to student performance on mandatory end-of-grade examinations. The results of these examinations, depending on location, are often taken into consideration when important school-related decisions such as student retention and educator evaluation and compensation are made. In some areas, the results can even affect school and school district operating budgets.

Increased focus on the use of examinations in schools has led to greater scrutiny of the process by which these tests are developed. Ensuring that examinations are valid and reliable is in the interest of all who are affected by their results. To that end, the American Educational Research Association (AERA), the American Psychological Association (APA), and the National Council on Measurement in Education (NCME) jointly developed and published the Standards for Educational and Psychological Testing (1999; hereafter, Standards). According to the Standards, "The proper use of tests can result in wiser decisions about individual programs than would be the case without their use and also can provide a route to broader an more equitable access to education and employment" (p. 1). The intent of the Standards is to "promote the sound and ethical use of tests" and to provide a basis for "evaluating the quality of testing practices" (p.1).

Education is not the only field, however, in which the results of examinations can be significant and consequential. Government agencies and other professional organizations frequently require applicants for credentials to pass license- or certification-granting examinations. Lawyers, physicians, electricians, and barbers are all examples of professionals who are required to receive government-issued licenses before being authorized to practice in their respective fields. Likewise, non-governmental entities often use examinations as part of the process to certify persons to perform tasks or operations that require specific skill sets. An information technology company, for instance, may require technicians to pass an examination before authorizing them to work on certain software programs. Tests that are used to grant certifications or award professional recognitions are frequently referred to as credentialing examinations.

Like other types of tests, credentialing examinations need to be developed in a manner that ensures their results produce desired thresholds of validity and reliability. According to the Standards (AERA, APA, \& NCME, 1999), "Tests and testing programs should be developed on a strong scientific basis. Test developers and publishers should compile and document adequate evidence bearing on test development" (p. 43). In addition to the guidelines listed in the Standards, credentialing examinations may also be required to meet additional criteria. Depending on the nature of the organization using the examination, compliance with guidelines set forth by national and international standards organizations, such as the American National Standards Institute (ANSI; 2013) and the International Organization for Standardization (ISO; 2013) may also be desired or required. Standardization organizations such as these ensure that licensure and certification
requirements are consistent among relevant parties. Adhering to these standards may also provide legal defensibility for the developers and administrators of these examinations.

The process by which credentialing examinations are developed is similar to that which is used for other types of tests. According to Downing (2006), the typical test development process is comprised of 12 steps. These steps are included in Table 1.1. The process begins with gaining an understanding of the purpose of the examination, the desired inferences to be made by test scores, as well as the general format to be used. Additional steps include defining the content to be used, creating test specifications, developing examination items, designing and assembling the test, and test production. Following these procedures, items are frequently field-tested, scored, and analyzed to judge the appropriateness of their inclusion in final versions of examinations. If applicable, a standard setting process may be used to recommend a minimum passing score for the test. This is followed by the development of a test reporting protocol, the establishment of an examination item bank, and the creation of technical reports that document the development process. Each step in this process is as important as the next and frequently serves as evidence for claims of examination validity.

Table 1.1
Test Development Process

| Step | Examples of development tasks and conce |
| :--- | :--- |
| 1. Overall plan | Guidance for test development activities <br> Confirm desired test interpretations <br> Test format |
| 2. Content definition | Sampling plan <br> Content-related validity evidence |
| 3. Test specifications | Content domain sampling <br> Desired item characteristics |
| 4. Item development | Item writer training <br> Item review, editing |
| 5. Test design and assembly | Design/create test forms <br> Develop pretesting considerations |
| 6. Test production | Publishing/printing activities <br> Security/quality control |
| 7. Test administration | Standardization issues <br> Proctoring, security, timing issues |
| 8. Scoring test responses | Quality control <br> Item analysis |
| 9. Passing scores | Standard setting <br> Comparability of standards |
| 10. Reporting test results | Accuracy, quality control <br> Misuse/retake issues |
| 12. Test technical report banking | Security issues <br> Usefulness, flexibility <br> Decumentation of validity evidence |
| Recommendations |  |

Note. Test development steps adapted from Downing (2006).

The development process for credentialing examinations can be affected by the unique characteristics these tests frequently exhibit. Unlike large-scale standardized tests used in education, credentialing examinations are often developed and administered for organizations representing occupations or fields with relatively few potential members. As such, the resources these groups are able to devote to test development and administration may be relatively limited. Although not wholly unique to credentialing examinations, examinees seeking a license or certification must also typically reach a predetermined minimum score, or cut score, in order to pass the examination and, therefore, be eligible to receive the desired credential. Tests with cut scores are also sometimes referred to as competency or mastery examinations because obtaining a score at or higher than the cut score infers examinee mastery or competency over a specified set of content standards. These attributes make credentialing examinations different than many standardized tests used in education, such as those used to measure the aptitude of prospective first-year college students. Such examinations are administered to thousands of examinees each year, creating large sets of data by which the development process is significantly aided.

The focus of this study is on one step in the process used to develop credentialing examinations. This step, frequently referred to as item analysis, is used to assess the degree to which field-tested items are suitable for inclusion in final versions of examinations. Item analysis, which Crocker and Algina (2008) defined as "the computation and examination of any statistical property of examinees' responses to an individual test item" is included in Step 6 of Downing's (2006) development process (p.311). The statistical properties most commonly used to assess individual examination items are item difficulty which, in classical test theory (CTT) terms, is the proportion of examinees that respond to an item correctly; and
item discrimination, which measures the degree to which an item differentiates between examinees who possess more of some characteristic intended to be measured by a test (e.g., subject area mastery) and those who possess less of the characteristic. This differentiation is typically operationalized as the difference between those examinees who perform relatively well on an examination and those who perform relatively poorly.

The procedures used to calculate item discrimination values for credentialing examinations with relatively small samples of field-test data are the focus of this study. A number of statistics are currently used to gauge item discrimination. A common characteristic among these methods, however, is that in calculating the discrimination values they consider scores from all examinees. In this study, discrimination values calculated using these traditional methods are referred to as unrestricted, because they incorporate data from all examinees. This research studies the effects of limiting the data used to calculate item discrimination to that of examinees who score around the test cut score. These values are referred to as restricted because they consider only a limited subset of examinee scores. In addition to examining how restricting scores used in the calculation of discrimination values affects the values themselves, the study investigates the effects of restricted discrimination values on certain aspects of test development, including item selection, examination reliability, and classification decision consistency.

## Research Questions

The following research questions are addressed in this study:

1. What are the effects on item discrimination values when the values are calculated using restricted samples of examinee test scores within varying ability ranges around real or anticipated cut scores?
2. What are the effects of calculating item discrimination values based on varying ranges of examinees around cut scores on item selection, examination reliability, and classification decision consistency?

Need for the Study
The current study represents a unique contribution to the field of test development for credentialing examinations. Current procedures used to calculate item discrimination values, although appropriate and effective for many types of tests, may not be ideal for competency examinations. In addition, the study's emphasis on tests with small samples of examinees represents the realistic-and under-studied-conditions of many testing programs, particularly those used by credential-granting organizations. Using small sample sizes also necessitates the use of classical test theory procedures, which, despite the emergence of more sophisticated measurement models, remain popular among developers of credentialing examinations. Some of the important potential benefits of this study are described in the following paragraphs.

First, although a variety of procedures may be used to calculate item discrimination values, a common characteristic of these procedures is that they each use the entire population of previous examinees as the criterion group when calculating the discrimination
values. In this manner, they treat scores of examinees at both the extreme upper and lower ends of a distribution of test scores as they do scores from examinees near the examination cut score. The focus of competency examinations, however, is on candidates near the cut score. By limiting the basis for calculating discrimination values to scores of examinees near the actual or estimated cut score, greater emphasis may be applied to items that discriminate more effectively amongst examinees with ability levels closest to those for which the test was designed to distinguish.

If the sample of examinees on which discrimination values are calculated is restricted, the restriction is likely to affect the selection of items for competency examinations. It is expected that discrimination indices based on criterion groups having a narrower range of ability or performance would produce uniformly attenuated discrimination indices. However, if discrimination values based on responses within a restricted sample of examinees are significantly different than those calculated using all examinees, the items selected for an examination will be dependent on the method employed. In other words, restricting the range of test scores used to calculate discrimination values permits items that discriminate among examinees with ability levels closest to those the cut score operationalizes to be selected over those that discriminate in other areas within the range of test scores.

The degree to which limiting the calculation of item discrimination values to scores of examinees around cut scores affects other aspects of test development also warrants further research. This study specifically examines how calculating discrimination values in this manner affects examination reliability and classification decision consistency.

Second, an important aspect of this research is that it was conducted within the context of competency examinations with relatively small numbers of examinees. Limiting the research to small-sample examinations is of particular benefit to developers of tests used to credential individuals. Unlike many large-scale educational achievement examinations for which item analysis may rely on large numbers of student responses and test scores, credentialing tests, due to their very nature, are often limited to smaller pools of examinees. Much of the research related to item analysis has focused on tests with large numbers of examinees. Fewer, however, have examined these issues as they specifically relate to examinations with smaller samples of available test scores. Focusing the study in this manner represents a significant contribution to small-sample examination development.

Finally, whereas the research presented here focused on examinations with small samples of response data, classical test theory procedures were appropriate and were used throughout. These procedures, though less computationally complex than more recent measurement theories, are still widely used by those responsible for developing credentialing examinations. The results of classical test theory-based procedures are also frequently viewed as being easier to interpret by individuals without backgrounds in measurement theory or statistics than the more complex models. As such, the results of this study are generalizable to a large segment of the test development field.

## CHAPTER 2

## LITERATURE REVIEW

The subjects addressed in this study draw upon relevant literature from three major areas of research: (a) analyses regarding item discrimination and its role in the test development process, (b) studies pertaining to the development of mastery or competency examinations used to credential individuals, and (c) research related to test development when relatively small samples of examinee scores are available. Significant research from each of these three areas is described in the sections that follow.

## Item Discrimination and the Test Development Process

Assessing the degree to which items discriminate between examinees who possess more of some knowledge, skill, or ability and those who exhibit less is an important element in the process by which items are selected for inclusion in all types of examinations. This process, commonly referred to as item analysis, is used to compute the statistical properties of examinee responses to individual test items (Crocker \& Algina, 2008). The goal of item analysis is to ensure that items selected for examinations yield levels of reliability and validity that sufficiently support the test's intended purpose. Items that discriminate between high- and low-performing examinees are typically viewed as being desirable and, as such, worthy of being included in an examination; items that do not are frequently removed from consideration for inclusion in an examination.

Many item discrimination methods have been developed to assess the relationship between examinee responses to individual test items and test performance. Although the approaches used to calculate discrimination values according to these indices vary, they share a common purpose: to identify test items to which high-scoring examinees have a high probability of responding correctly, and to which low-scoring examinees have a low probability of responding correctly. A description of commonly used item discrimination indices is included in the sections that follow.

The index of discrimination, commonly referred to as the $D$-index, was one of the earliest methods developed to calculate item discrimination (Crocker \& Algina, 2008). $D$ is calculated by dividing examinees into upper- and lower-scoring groups of equal size. The criterion used to identify an examinee as belonging to either group is his or her observed test score. The proportion of examinees responding correctly to a particular item in the lowerscoring group ( $p_{\text {lower }}$ ) is subtracted from the proportion of examinees responding correctly in the upper-scoring group ( $p_{\text {upper }}$ ):

$$
\begin{equation*}
D=p_{\text {upper }}-p_{\text {lower }} \tag{2.1}
\end{equation*}
$$

Scores for this index range between -1.00 and 1.00, with negative values indicating negative discrimination, an undesirable situation in which a smaller proportion of higher-scoring examinees than lower-scoring examinees respond correctly to an item. For tests with dichotomously scored items, the proportion of correct responses to a particular item for a group of examinees also represents that item's average score for the group. Therefore, $D$ -
values also represent the difference in average item score between the high- and low-scoring groups (Ebel, 1967).

Although $D$-values are mathematically simple to compute, a number of drawbacks have limited their widespread use. With no known sampling distribution, it is not possible to test for statistical significance between $D$-values or to identify whether a particular $D$-value is significantly greater than zero (Crocker \& Algina, 2008). In addition, the index of discrimination can only be used for items that are scored dichotomously. The selection of the upper- and lower-scoring groups can also significantly impact the calculated values, which may be particularly problematic for examinations with a restricted range of scores or where only small numbers of candidates are available.

When item analysis is conducted, D-values may be used to help determine the appropriateness of including individual items in the final version of an examination. Ebel (1965) developed a guideline for interpreting $D$-values:

1. If $D$ is .40 or greater, the item is performing satisfactorily and no revision is required.
2. If $D$ is between .30 and .39 , little or no revision is required.
3. If $D$ is between .20 and .29 , the item needs revision.
4. If $D$ is .19 or lower, the item should not be used.

Items with large positive $D$-values, which represent large differences in the proportion of correct responses between the two groups, are viewed as suitable, while items with small or negative $D$-values are not. According to Ebel, items with small $D$-values, indicating small differences in scores between the lower- and upper-scoring groups, should be revised before
being considered for inclusion in an examination, as, among other reasons, a low $D$-value may simply indicate that the item contains problematic wording.

Several studies have examined the use of variations to the index of discrimination. A classic study by Kelley (1939), for example, explored varying the size of the groups upon which $D$-values are calculated. Instead of using all test scores to establish upper- and lowerscoring groups, Kelley found that utilizing the upper and lower $27 \%$ of test scores produced more sensitive and stable results. Beuchert and Mendoza (1979), however, found that when sample sizes were large enough, using the upper and lower $30 \%$ or $50 \%$ of test scores produced nearly identical results to those produced by the $27 \%$ recommended by Kelley. Although Kelley, as well as Beuchert and Mendoza, addressed issues related to the current study, neither focused the calculation of item discrimination values on contiguous groups of varying sizes around examination cut scores. In addition, the researchers emphasized using groups at the extreme ends of test score distributions, a position at odds with the research presented here.

In another important study, Brennan (1972) suggested that using groups of equal size was not necessary when calculating $D$. Creating groups of equal size, as was done in the research described previously, was a result, according to Brennan, of "the preoccupation of test theory with the normal distribution" (p. 291). Actual score distributions for most examinations, however, are not normal. Brennan called for the creation of a new index, referred to as $B$, to measure item discrimination. The index is represented by the following formula:

$$
\begin{equation*}
B=\frac{U}{n_{1}}-\frac{L}{n_{2}} \tag{2.2}
\end{equation*}
$$

where $U$ represents the number of examinees in the upper-scoring group responding correctly;
$L$ represents the number of examinees in the lower-scoring group responding correctly; and
$n_{1}$ and $n_{2}$ represent the total number of examinees in the upper- and lower-scoring groups, respectively.

According to Brennan, $B$ allows for an estimate of discrimination that does not require using groups of equal size. An important aspect of $B$, particularly as it relates to this research, is that it also allows evaluators to select the point along the distribution of test scores that most appropriately divides the upper and lower scoring groups:

Furthermore, regardless of the shape of the distribution of test scores, it seems reasonable to allow the test evaluator the freedom to choose the cut-off points between the upper and lower groups. Only he can determine the cut-off points that yield meaningful and interpretable upper and lower groups based upon his consideration of the test content, student population, and overall expectations for student performance on the test. When the test constructor is free to choose the cutoff points, there is, clearly, no reason to expect that the resulting groups will be of equal size. (p. 292)

Although the calculation of discrimination values used in this research does not utilize any adaptation of $D$ or $B$, Brennan's claim that the most appropriate method used to calculate item discrimination values may be examination-dependent is relevant. A major consideration in this study is that the focus of mastery examinations is the test cut score. It appears reasonable, therefore, to use the cut score as the central point in the distribution of test scores upon which discrimination values are estimated.

In addition to the index of discrimination, several methods utilizing variations of the Pearson product-moment correlation coefficient have been developed to measure item
discrimination. These methods are used to calculate the degree to which item performance and overall test performance are correlated. Two of the more commonly used correlational indices are the point-biserial correlation and the biserial correlation. Although both of these indices utilize correlation statistics to describe discriminating power, the results they produce are different. A brief description of each index is included in the paragraphs that follow.

The point-biserial correlation is the observed correlation between examinee performance on a dichotomously scored item and overall test score (Livingston, 2006). For dichotomously scored items, correct responses are scored 1 and incorrect responses are scored 0 . The observed correlation between item response and test performance forms the basis for the point-biserial correlation. Like all correlation coefficient values, the pointbiserial values range between -1.00 and 1.00. Negative values represent items that discriminate negatively, while positive values represent those that discriminate positively. Larger values represent items with greater levels of discriminating power.

The point-biserial statistic, $r_{p b i s}$, may be calculated using the following formula:

$$
\begin{equation*}
r_{p b i s}=\frac{\left(\mu_{+}-\mu_{x}\right)}{\sigma_{x}} \sqrt{p / q} \tag{2.3}
\end{equation*}
$$

where $\mu_{+}$is the mean total score for those who respond to the item correctly; $\mu_{x}$ is the mean total score for the entire group of examinees; $\sigma_{x}$ is the standard deviation for the entire group of examinees; $p$ is item difficulty; and
$q$ is equal to $(1-p)($ Crocker \& Algina, 2008).

A common criticism of the point-biserial statistic is that it may sometimes be spurious because the item score contributes to the total score for each examinee. This can result in inflated discrimination values. The effect is greatest for examinations with relatively few items, resulting in a curious situation in which shorter examinations, which typically produce lower levels of reliability, exhibit higher item discrimination values (Burton, 2001). For examinations with more than 25 items, such as those used in this study, however, the effect is rarely problematic and does not significantly affect discrimination values (Crocker \& Algina, 2008).

The biserial correlation index produces results similar to the point-biserial index, but is calculated in a slightly different manner. The biserial, which was first derived by Pearson (1909), treats scores on dichotomously scored items as indicators of an unobservable underlying proficiency. The biserial estimates the correlation between this latent underlying proficiency and total test score.

The biserial statistic, $r_{b i s}$, may be calculated using the following formula:

$$
\begin{equation*}
r_{b i s}=\frac{\left(\mu_{+}-\mu_{x}\right)}{\sigma_{x}}(p / Y) \tag{2.4}
\end{equation*}
$$

where $\mu_{+}$is the mean total score for those who respond to the item correctly;
$\mu_{x}$ is the mean total score for the entire group of examinees;
$\sigma_{x}$ is the standard deviation for the entire group of examinees;
$p$ is item difficulty; and
$Y$ is the $Y$ ordinate of the standard normal curve at the $z$-score associated with the $p$ value for the item (Crocker \& Algina, 2008).

In general, the biserial statistic produces larger discrimination values than those produced by the point-biserial. This is due to the fact that the $Y$ ordinate on the normal curve, which is used to calculate the biserial, will always be larger than $\sqrt{p q}$, which is used to calculate the point biserial (Lord \& Novick, 1968). The differences are more profound when item difficulty values are less than 0.25 or greater than 0.75 . Differences in item discrimination values, therefore, may be attributed not only to qualitative differences among examination items, but also to the statistic used to estimate the level of discrimination.

Item response theory, a general statistical theory that relates performance on test items to the abilities the test is intended to measure, may also be used to calculate item discrimination values (Hambleton \& Jones, 1993). At its core, item response theory estimates the probability that particular examinees will respond in certain ways to items with certain characteristics (Yen \& Fitzpatrick, 2006). Although the Rasch, or one-parameter logistic model, provides estimates for item location (i.e., item difficulty) only, the two- (and greater) parameter logistic models estimate difficulty and item discrimination. The discrimination estimate produced by item response theory models is analogous to the itemtotal correlation statistics (i.e., the biserial and point-biserial) used in classical test theory. Item response theory is also computationally more complex than the classical test theory discrimination indices mentioned earlier. The two-parameter logistic model uses two parameters to describe each item. These parameters include item difficulty, $b_{i}$, and item discrimination, $a_{i}$. The estimates may be calculated using the following equation:

$$
\begin{equation*}
P_{i}\left(X_{i}=1 \mid \theta\right)=\frac{1}{1+\exp \left[-D a_{i}\left(\theta-b_{i}\right)\right]} \tag{2.5}
\end{equation*}
$$

where $P_{i}$ represents the probability of a correct response $\left(X_{i}=1\right)$ given a particular ability level $(\theta)$; and
$D$ represents a multiplicative constant, typically set at either 1.7 or 1.702 (Yen \& Fitzpatrick, 2006).

When the parameters are plotted, they create what are commonly referred to as item characteristic curves (ICCs). The $a_{i}$, or discriminating parameter, specifies the slope of the ICC, with steeper slopes indicating greater levels of item discrimination (Luecht, 2006). An example of an ICC produced using a three-parameter logistic model, with the third parameter representing examinee noise or guessing, is shown in Figure 2.1. In the figure, the slope, labeled $a$, represents item discrimination.
0.5

Figure 2.1. Example of item characteristic curve (ICC)

Estimates produced using item response theory require larger sample sizes than those produced using classical test theory. Reise and Yu (1990), for example, found that at least 500 cases were needed to produce dependable item parameter estimates, including item discrimination, when using item response theory, with 1,000 to 2,000 cases required for more accurate estimates. Hambleton and Jones (1993) found that the number of cases required to effectively utilize item response theory depended on the particular model being used; however, in general, they recommended no less than 500 cases be used. Despite its advantages, therefore, when calculating discrimination values, developers of examinations for which relatively small samples of examinee responses are available must typically rely on classical test theory procedures, such as the biserial or point-biserial item-total correlation statistics.

Much of the research associated with item discrimination and its role in the test development process has focused on comparisons between the various indices. Beuchert and Mendoza (1979), for example, analyzed the results of eight studies that compared discrimination values produced by a number of indices. Four of the studies found the values to be virtually indistinguishable. The others found minor, but sufficiently significant, differences leading to a recommendation against using particular indices in certain situations. Using a Monte Carlo statistical simulation approach, Beuchert and Mendoza developed sixteen 100-item examinations and administered them to two pools of simulated examinees, resulting in 32 distinct testing scenarios. The pools of examinees were comprised of 60 and 200 examinees respectively. The researchers then calculated discrimination values for each examination item using ten different discrimination indices. When compared, the differences the various indices produced were, according to the researchers, "extremely small, or
nonexistent in situations intended to accentuate those differences" (p. 116). Based on these results, Beuchert and Mendoza recommended using the most computationally simple index.

In a related study, Oosterhof (1976) compared discrimination values produced by 19 different indices using exploratory factor analysis. His research found the loadings representing each of the discrimination indices to be "impressively high," with six indices exhibiting loadings greater than 0.98 and all but one with loadings greater than 0.85 when loaded against a single common factor (p. 149). Oosterhof summarized his findings in the following manner:

When any of the selected indices are used to evaluate the relative performance of an item, the preference of one index over another minimally affects the resulting analysis. Preference towards a particular index would more appropriately be based on convenience of calculation or intuitive preference. It is inappropriate to suggest that using any of the common indices included in the present study has an appreciable effect on the eventual outcome of an analysis. (p. 149)

A more recent study by Fan (1998) compared the results of item analysis using both item response theory and classical test theory for a 108-item examination given to over 190,000 high school students in Texas. Fan estimated item discrimination values for each item using a two- and three-parameter logistic item response theory model and the pointbiserial statistic. The majority of correlation coefficients for the discrimination values ranged between 0.60 and 0.90 . Although this relationship was somewhat weaker than that found for differences in item difficulty values, which was also assessed in the study, Fan indicated that the overall relationship between discrimination values calculated using item response theory and classical test theory to be "moderately high to high" (p. 378). According to Fan,

The findings here simply show that the two measurement frameworks produced very similar item and person statistics both in terms of the comparability of item and person statistics between the two frameworks and in terms of the degree of invariance of item statistics from the two competing measurement frameworks. (pp. 378-379)

Fan's findings are similar to conclusions reached by Thorndike (1982), who, in discussing the then relatively new use of item response theory in test development procedures, wrote:

For the large bulk of testing, both with locally developed and with standardized tests, I doubt there will be a great deal of change. The items that we will select for a test will not be much different from those we would have selected with earlier procedures, and the resulting tests will continue to have much the same properties. (p. 12)

Additional research associated with item discrimination has introduced new or modified versions of previously established indices. Harris and Subkoviak (1986), for example, developed a new index of discrimination, referred to simply as the agreement index. In developing the index, the authors hoped to create a procedure that incorporated certain aspects of item response theory, but which was computationally less complex. Designated $P\left(X_{c}\right)$, the agreement may be calculated using the following formula:

$$
\begin{equation*}
P\left(X_{c}\right)=\frac{a_{11}-a_{22}}{N} \tag{2.6}
\end{equation*}
$$

where $a_{11}$ represents the number of examinees responding to an item correctly; $a_{22}$ represents the number of examinees responding incorrectly; and $N$ represents the total number of examinees. $P\left(X_{c}\right)$ can be interpreted as the probability of agreement between performance on a single item and performance on the overall examination, with ideal items having values equal to 1.00 .

In their study, Harris and Subkoviak (1986) compared the selection of items for a set of examinations using both the agreement index and a two-parameter logistic item response theory model. The examinations were varied in terms of numbers of items, including lengths
of 30,50 , and 100 items, and numbers of examinees, ranging between 30,60 and 120. The results indicated that the average correlation between items selected using these two methods was 0.91 . According to the authors, the correlation was sufficiently strong as to recommend the use of the agreement index, as estimates are much easier to compute than when using the two-parameter logistic model.

## Credentialing Examinations

In many instances, examinations are developed for the purpose of classifying examinees into two or more groups. These types of tests, also frequently referred to as mastery or competency examinations, are used in a variety of fields. Competency examinations are used in education, for example, to identify students who may need remedial instruction, or to determine fitness for graduation. As such, they are not norm-referenced, as many achievement examinations used in education are, but rather are criterion-referenced; that is, examinees must meet specified standards, as operationalized by a pre-determined score, in order to pass. Government agencies and other professional organizations use mastery examinations to credential individuals in a variety of fields and occupations. Doctors, lawyers, and teachers, for example, must pass competency examinations before receiving the credentials they need to practice in their respective fields.

Buckendahl and Davis-Becker (2012) noted that individuals who take competency examinations are "candidates for a license, certification, or other credential" (p. 485).

Licenses represent a legal authority to practice in a particular field and are typically awarded by federal or state agencies. In order to begin practicing in fields requiring a governmentissued license, individuals must complete an associated licensure program. In most cases,
these programs require the candidates to pass a competency examination. In contrast to licensure programs, certification programs are not government-regulated, but rather are typically managed within an occupational field and are usually voluntary. A certification attests to the fact that the individual has met a credentialing organization's standards and is entitled to make the public aware of his or her professional competence.

A primary purpose behind using competency examinations as a requirement for granting credentials, both government-regulated licenses and certifications, is ensuring that individuals are properly qualified to practice in their respective fields. The requirement made by many states for certain occupations to obtain licensure is also driven by the desire to promote public safety. According to the Standards (AERA et al., 1999):

Tests used in credentialing are intended to provide the public, including employers and government agencies, with a dependable mechanism for identifying practitioners who have met particular standards. Credentialing also serves to protect the profession by excluding persons who are deemed to be not qualified to do the work of the occupation. Tests used in credentialing are designed to determine whether the essential knowledge and skills of a specified domain have been mastered by the candidate. (p. 156)

By requiring individuals in these occupations to obtain licensure, the public may be confident that those providing services will do so in a safe and effective manner. Those responsible for credentialing programs, however, must balance this consideration with the need to ensure credentialing requirements are not so stringent so as to prohibit those who have been trained and who may be qualified from practicing in the field (Clauser, Margolis, \& Case, 2006). In some situations, marginally qualified practitioners may be better than too few or no practitioners. In these cases, the public might actually be harmed by exceedingly high credentialing standards.

Government and professional organizations have used examinations to regulate a variety of occupations for hundreds of years. Chinese civil servants, for example, have been required to pass written examinations for nearly three millennia, with similar requirements for the medical and legal fields in place sometime before 500 B.C.E. (DuBois, 1970). Modern use of credentialing examinations originated, to a large degree, in the medical field. Garcia-Ballester, McVaugh, and Rubio-Vela (1989) listed several factors behind the rise of government-regulated standards in the medical field. Among these included: a concern for quality healthcare; a desire to restrict access to the field to those already practicing, in essence creating a monopoly for current practitioners; and political confrontations over the power to regulate certain occupations.

Today, government agencies continue to regulate an ever-growing number of fields. Atkinson (2012) listed the occupations in each state that required licensure as of 2010. California, at the top of the list, licensed 177 professions. Nine additional states licensed over 100 occupations each. Missouri, the state with the fewest number of licensed professions, required licenses for 41 occupations. Table 2.1 lists the ten states with the most and fewest licensed occupations.

Table 2.1
States with Most and Fewest Licensed Occupations

| Rank | State | Licensed <br> Occupations | Rank | State | Licensed <br> Occupations |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | California | 177 | 41 | Colorado | 69 |
| 2 | Connecticut | 155 | 42 | North Dakota | 69 |
| 3 | Maine | 134 | 43 | Mississippi | 68 |
| 4 | New Hampshire | 130 | 44 | Hawaii | 64 |
| 5 | Arkansas | 128 | 45 | Pennsylvania | 62 |
| 6 | Michigan | 116 | 46 | Idaho | 61 |
| 7 | Rhode Island | 116 | 47 | South Carolina | 60 |
| 8 | New Jersey | 114 | 48 | Kansas | 56 |
| 9 | Wisconsin | 111 | 49 | Washington | 53 |
| 10 | Tennessee | 110 | 50 | Missouri | 41 |

Note. Information derived from Atkinson (2012).

A common subject in the literature associated with credentialing examinations is the procedures by which these tests are developed. Credentialing examinations, not unlike other tests, must be developed in a manner that produces levels of reliability and validity that support the inferences the resulting tests score are intended to make. An important aspect in the development of these tests is the establishment of a cut score. The cut score represents the score examinees must obtain in order to pass the examination, and should, as Cizek (2012a) pointed out, be established using procedures that are as "defensible and reproducible
as possible" (p. 6). Appropriately, therefore, the Standards (AERA et al., 1999) recommend that those responsible for setting standards be "concerned that the process by which cut scores are determined be clearly documented and defensible" (p. 54).

The process used to develop cut scores is referred to as standard setting. Although a thorough review of the many standard-setting methodologies currently in use is beyond the scope of this study, a brief and general description of typical standard setting procedures is warranted. During standard setting conferences, subject matter experts, who are also frequently referred to as judges or participants, review definitions of the knowledge, skills, and attributes examinees must possess to be deemed minimally qualified for inclusion in a particular proficiency category. For many examinations, these categories may simply represent those who pass the test, and those who do not. Depending on the standard setting method used, the participants then make judgments about either individual examinees or individual test items. Through a variety of method-dependent procedures, the participants’ judgments are translated into a recommended cut score. Once approved by the examination's governing body, candidates must score at or above the cut score in order to pass the test.

The accuracy of classifications made when utilizing credentialing examinations with cut scores is, of course, critically important. Because of this, more focus is given to ensuring precision around the cut score. According to the Standards (AERA et al., 1999):

Tests for credentialing need to be precise in the vicinity of the passing, or cut, score. They may not need to be precise for those who clearly pass or clearly fail. Sometimes a test used in credentialing is designed to be precise only in the vicinity of the cut score. (p. 157).

The above quote is of particular relevance to the current study. As discussed previously, traditional methods used to calculate item discrimination values consider scores of all examinees, regardless of their proximity to the examination cut score. By restricting
the scores upon which discrimination values are calculated to those near the cut score, more precision is applied to those for whom the accuracy of the cut score is most relevant and consequential.

The knowledge and skills needed to practice in licensed fields changes periodically. In many instances advances in technology or methods of practice drive these changes. As such, the examinations used to credential individuals in these fields must also be altered to reflect the changes. When such changes occur, the examination cut score must also be reevaluated. Again, the Standards (AERA et al., 1999) describe the importance of this process:

Practice in professions and occupations often change over time. When change is substantial, it becomes necessary to revise the definition of the job, and the test content, to reflect changing circumstances. When major revisions are made in the test, the cut score that identifies required test performance is also reestablished. (p. 157)

In addition to research associated with the establishment and use of cut scores, the literature related to credentialing examinations has also emphasized issues related to examination validity and reliability. Researchers have focused on how these principles, critical to the development of any test, specifically relate to credentialing examinations.

According to the Standards (AERA et al., 1999), test validity is "the degree to which evidence and theory support the interpretation of test scores entailed by proposed uses" (p. 9). The interpretation of test scores produced by credentialing examinations is that examinees who pass the test are qualified to receive the associated credential and, therefore, are qualified to practice in their respective fields. According to Clauser et al. (2006):

Because the primary interpretation based on scores from licensing and certifying tests is that the examinee is (or is not) suitable for licensed or certified practice, it follows that a central issue of validity theory in this context is the question of whether the test scores properly classify examinees. (p. 716)

Obtaining the evidence necessary to support claims of examination validity is referred to as test validation. Cizek (2012b) summarized this process:

Validation is the ongoing process of gathering, summarizing, and evaluating relevant evidence concerning the degree to which that evidence supports the intended meaning of scores yielded by an instrument and inferences about standing on the characteristic it was designed to measure. (pp. 35-36)

As it specifically relates to credentialing examinations, gathering validity evidence can, at times, be somewhat challenging. Whereas the degree to which credentialing tests accurately classify examinees is the critical validity concern, it follows that a thoughtful analysis of this question might compare the performance of examinees who pass the examination with those who fail. Examinees who fail, however, are typically not allowed to practice in the field, and, therefore, such comparisons are normally not possible (Clauser, Margolis, \& Case, 2006).

A more realistic approach to gathering validity evidence for credentialing examinations may be one in which evidence supporting the appropriateness of the examination's interpretive argument is identified. According to Kane (1992):

A test-score interpretation always involves an interpretive argument, with the test score as a premise and the statements and decisions involved in the interpretation as conclusions. The inferences in the interpretive argument depend on various assumptions, which may be more-or-less credible. Because it is not possible to prove all of the assumptions in the interpretive argument, it is not possible to verify this interpretive argument in any absolute sense. The best that can be done is to show that the interpretive argument is highly plausible, given all available evidence. (p. 527)

According to Clauser et al. (2006), an area that is particularly important to the interpretive argument made by credentialing examinations is evidence that the test was constructed using rigorous development procedures. These procedures must ensure that the
examination content realistically reflects the knowledge and skills needed by those seeking licensure or certification.

Raymond and Neustel (2006) underscored the importance of ensuring that the content associated with credentialing examinations reflected requirements for safe and effective practice in the fields for which credentials are awarded. According to the authors, this can be accomplished through the use of practice analyses, which "identify the job responsibilities of those employed in the profession" (p. 181). After conducting these analyses, the knowledge, skills, and attributes of the associated responsibilities may be obtained. These, in turn, aid developers in establishing a test blueprint, or specification. Raymond and Neustel listed several useful tools to aid in the conduct of practice analyses, including task inventory questionnaires, task statements, and job responsibilities scales.

Although the majority of their study evaluated various methodologies used to ensure appropriate content, Raymond and Neustel (2006) also highlighted the importance of using empirical data, such as computed "statistical indices of item-domain congruence..." to inform the item selection process (p. 206). This process, inevitably, includes an analysis of the discriminating power of potential examination items.

Clauser et al. (2006) also examined methods used to identify appropriate content for credentialing examinations. Like Raymond and Neustel (2006), the authors emphasized the importance of generating job responsibility inventories. In order to limit the size and scope of the examination, however, Clauser et al. suggested restricting task inventories to those activities that ensured public safety:

The topic of task list should include only those elements that are necessary to protect the public; entries that might be necessary for success in the field but are not required for safe practice should be omitted. (p. 705)

Reliability is also the focus of considerable research related to credentialing examinations. Put simply, examination reliability is the "desired consistency (or reproducibility) of test scores" (Crocker \& Algina, 2008, p. 105). Over time, several methods have been developed to measure reliability. Early procedures relied on administering the same examination multiple times. Utilizing the test-retest method, for example, the developer administers an examination to a group of examinees, waits a predetermined amount of time, and then re-administers the examination. The correlation between examinee test scores, referred to in this context as the coefficient of stability, is then calculated (Crocker \& Algina, 2008). Similar methods require administering alternate test forms to examinees and calculating the correlation between scores on the forms.

Other approaches used to estimate reliability rely on single administrations of examinations. One such procedure is the split-half method, in which a single examination form is administered to a group of examinees. Before the test is scored, however, the examination is divided into two equivalent halves. The halves are scored as if they were separate examinations, and the correlation between test scores is calculated for each examinee. The method assumes that the halves are strictly parallel. In addition, because the split-half tests contain fewer items than the whole examination, the coefficient underestimates the reliability of the full-length test. The Spearman Brown correction was designed to overcome this problem (Crocker \& Algina, 2008).

Some of the most popular reliability estimates, however, rely on covariances between examination items. Possibly the most popular method, developed by Cronbach (1951), produces a unique estimate for the internal consistency of test scores. The method,
commonly referred to as Cronbach's alpha, or coefficient alpha, can be calculated using the following formula:

$$
\begin{equation*}
\hat{\alpha}=\frac{k}{k-1}\left(1-\frac{\sum \hat{\sigma}_{i}^{2}}{\hat{\sigma}_{x}^{2}}\right) \tag{2.7}
\end{equation*}
$$

where $k$ is the number of items on the examination;
$\hat{\sigma}_{i}^{2}$ is the variance of item $i$, and
$\hat{\sigma}_{x}^{2}$ is the total test variance (Crocker \& Algina, 2008).
Using coefficient alpha, it is possible to treat each test item as a subtest and, therefore, to estimate the degree of reliability between the subtests.

Although coefficient alpha is commonly used as an estimate of reliability for all types examinations, including those used to credential individuals, the literature suggests that other forms of reliability estimates may also be appropriate when an examination is used to make classification decisions. According to Haertel (2006):

When continuous scores are interpreted with respect to one or more cut scores, conventional indices of reliability may not be appropriate, and the standard error of measurement may not be directly informative concerning classification accuracy. Such cases arise when examinees above a cut score are classified as passing or proficient, for example. Instead of standard errors, users may be concerned with questions such as the following: What is the probability that an examinee with a true score above the cut score will have an observed score below the cut score, or conversely? What is the expected proportion of examinees who would be differently classified upon retesting? (p. 99)

Classification decision consistency indices have been developed to measure the degree to which the same decisions are made from two different sets of measurements. One of the earliest indices, referred to simply as $\hat{P}$, can be explained using a two-by-two table,
similar to that shown in Figure 2.2. The cells in the table represent the proportions of examinees who are classified as either masters or non-masters after taking different forms of the same examination. The cell labeled $\hat{P}_{11}$, for example, represents the proportion of examinees classified as masters by both forms. The cell labeled $\hat{P}_{10}$ represents the proportion of examinees classified as masters using the first form, but as non-masters using the second form.


Figure 2.2. Probabilities of consistent classifications using two forms (Crocker \& Algina, 2008)

The estimated probability of a consistent decision, therefore, can be calculated using the following formula:

$$
\begin{equation*}
\hat{P}=\hat{P}_{11+} \hat{P}_{00} \tag{2.8}
\end{equation*}
$$

Values for $\hat{P}$ can range between 0.00 and 1.00 , with 0.00 representing complete inconsistency and 1.00 representing total consistency.

Although $\hat{P}$ was recommended as a measure of classification decision consistency (Hambleton \& Novick, 1973), the index is not without flaw. For example, a value greater than 0.00 would be expected by chance, even if the measurements used were uncorrelated. In an effort to overcome this situation, Swaminathan, Hambleton, and Algina (1974) recommended using Cohen's (1960) $\kappa$ as a measure of classification decision consistency. The coefficient can be calculated using the following formula:

$$
\begin{equation*}
\kappa=\frac{P-P_{c}}{1-P_{c}} \tag{2.9}
\end{equation*}
$$

where $P_{c}$, also referred to as the chance consistency, is the probability of a consistent decision, and may be calculated using the following formula:

$$
\begin{equation*}
P_{c}=P_{1 .} . P_{.1}+P_{0 .} \cdot P_{.0} \tag{2.10}
\end{equation*}
$$

The four elements used to calculate $P_{c}$ represent the margin sums in the hypothetical table displayed in Figure 2.2. That is, $P_{1 .}$ represents the probability of a mastery classification on one form and $P_{.1}$ represents a similar probability on the other form. The same holds true for $P_{0}$ and $P_{.0}$, which represent misclassifications on the forms. The interpretation of $\kappa$ is somewhat different than that of $\hat{P}$, as it represents the increase in decision consistency over that expected by chance. The coefficient is 0.00 when there is no increase, and 1.00 when there is maximum increase (Crocker and Algina, 2008).

A limitation of the classification decision consistency indices discussed thus far is that they each require multiple administrations. Subkoviak (1976) and Huynh (1976)
developed procedures by which $P$ and $\kappa$ could be estimated from a single administration. The approaches produce estimates using a hypothetical form that is exchangeable with the examination from which data is gathered (Crocker \& Algina, 2008). Huynh's method has been shown to produce fairly accurate estimate of $P$ and $\kappa$ for parallel tests with as few as 10 items (Subkoviak, 1978).

Issues related to the validity and reliability of credentialing examinations are also significant when the legal defensibility of such tests are considered. According to Atkinson (2012), "as the number of regulated professions which use an examination as one criterion of eligibility increases so will the likelihood of a legal challenge" (p. 506). Although much of the attention competency examinations receive is on the score that defines passing and failing, Atkinson found that legal challenges rarely contest the cut scores themselves. Rather, legal challenges are focused on the entire test development process. According to Atkinson: "The basis for legally substantiating an examination program and its Pass/Fail determination discriminating between those recognized as establishing competence and those who have not, will necessitate an analysis of the entire examination development..." (p. 511).

Legal defensibility is an important consideration within the context of the current study because item analysis, including the calculation and evaluation of item discrimination values, is a critical step in the test development process. If calculating discrimination values using only restricted samples of examinee responses is more appropriate for credentialing examinations, the issue becomes relevant to the test's defensibility.

Very few studies have assessed the role item discrimination plays in the development of credentialing examinations. Although not specific to credentialing tests, Harris and Subkoviak (1986) discussed the importance of item discrimination in the development of
mastery examinations in general. They advocated developing tests that maximize score differences between groups who pass and fail, while simultaneously minimizing score differences within these groups:

For a mastery test, this means selecting items that discriminate between masters and non-masters, as opposed to within masters and within non-masters. The consensus appears to be that a good mastery item is one which masters answer correctly and non-masters answer incorrectly. (p. 496)

More closely related to the current study, Buckendahl and Davis-Becker (2012) conducted research regarding the establishment of passing standards for credentialing examinations. Although the majority of their work emphasized the processes used to develop recommended cut scores for credentialing tests, and the not methodologies used to conduct item analysis, the authors highlighted an important consideration related to the development of credentialing examinations. The organizations responsible for credentialing individuals often do not have the resources needed to support all aspects of a comprehensive test development process. Raymond and Neustel (2006) also underscored this point. According to them, "credentialing organizations often lack the resources required for the types of thorough experimentation and investigation required to support [their] claims..." (p. 205). This is an important consideration within the context of the current research because it may explain, at least in part, why item analysis for credentialing examinations frequently must rely on small samples of examinee responses. Credentialing organizations, in many cases, simply do not have resources available to collect the large numbers of responses necessary to conduct a more complete analysis of potential test items. In some cases, these constrained resources are not only financial in nature, but are also related to the fact that in many fields, the pool of potential examinees is relatively small.

## Test Development with Small Samples of Examinee Responses

An important aspect of the current study is that is utilizes tests for which relatively small numbers of examinee responses are available for item analysis. As discussed earlier, this is a realistic condition under which many credentialing examinations are developed. Jones, Smith, and Talley (2006) characterized this situation as one in which fewer than 200 examinee responses were available for analysis "either because the testing program is new or because the target population is inherently small" (p. 487).

A primary consideration in such situations is the process by which field-test data may be gathered for further analysis. According to the Standards (AERA et al., 1999), this process should be documented and should utilize examinees drawn from the population for which the examination was constructed:

When item tryouts or field tests are conducted, the procedures used to select the sample(s) of test takers for item tryouts and the resulting characteristics of the sample(s) should be documented. When appropriate, the sample(s) should be as representative as possible of the population(s) for which the test is intended. (p.44)

According to Jones et al. (2006), for examinations with relatively small numbers of possible test takers, this recommendation can be challenging because the developer must be in a position "to make sound statistical inferences while working within the constraints imposed by the testing system; namely, that there are fewer than 200 test takers available to participate in field testing - perhaps far fewer" (p. 493).

Millman and Greene (1989) suggested starting with a preliminary tryout of test items given to as few as five or six members of the target population or subject matter experts. The tryout would be followed by interviews aimed at ascertaining the examinees' thoughts regarding the test and individual test items. Jones et al. (2006) also provided recommendations for dealing with examinations for which relatively small pools of
examinees are available. They suggested recruiting a stratified sample of examinees that is distributed similarly to the projected population. Such a strategy can help ensure that the sample is diverse enough to allow for a meaningful evaluation of the items' discriminating properties.

Once field-test data is collected, a determination regarding the appropriate measurement model to use must be made. As discussed previously, in many cases, analysis of items may be limited to classical test theory, as other models, such as item response theory, require larger numbers of examinees. Jones et al. (2006) examined the potential use of various measurement models under three different conditions: (a) when there are no pretest data, (b) when a pretest sample up to $N=100$ is available, and (c) when a pretest sample of $N=100$ to 200 is available.

According to Jones et al. (2006), when no item response data is available, developers must rely on rigorous item review procedures that emphasize item appropriateness, alignment with test specifications, content domain representativeness, potential item bias, and the adequacy of instructions. The previously described recommendation by Millman and Greene (1989), that the items may be administered to a handful of subject matter experts, may also be beneficial. Thorndike (1982) suggested that item difficulty and discrimination parameters might be estimated using regression analysis. This approach requires previously used items with known item parameters as well as judges who estimate the difficulty of new items.

For examinations with sample sizes of $N=$ up to 100 , Jones et al. (2006) found itemlevel statistics, to include item discrimination values, to be stable using classical test theory procedures. Citing a study conducted by Farish (1984), the authors found that when utilizing a random sampling of examinee responses, item discrimination values converged with full
sample statistics when $N$ was as small as 40 . According to Jones et al., "In the end, if the item pool is small, sample sizes as low as $N=50$ may provide enough information to select desirable test items for inclusion in new test forms" (p. 506). The authors also discussed the use of item response theory for samples in this range. For tests being developed on $N \leq 100$ examinee responses, they found that the one-parameter logistic model could be effective in estimating item difficulty. The one-parameter model, as discussed earlier, however, holds all discrimination values as equal, and, therefore, is not appropriate for studies investigating the role of discrimination in item selection.

Finally, for sample sizes of $N=100$ to 200, Jones et al. (2006) found that classical test theory and item response theory procedures produced stable item parameters, which "facilitates making reliable item selection decisions within a larger item pool" (pp. 506-507).

In addition to the research conducted by Jones et al. (2006), other studies have compared the utility of classical test theory and item response theory in dealing with smallscale examinations. Not surprisingly, for examinations with 200 or fewer examinee responses, most suggest using classical test theory. Hambleton and Jones (1993), for example, found that whereas the number of cases required to use item response theory depended, to a certain extent, on the model being employed, at least 500 cases were desired.

## Summary

The preceding sections described the relevant literature in three areas: (a) research related to item discrimination and its role in the test development process, (b) the development of credentialing examinations, and (c) the development of examinations when relatively small samples of examinee data are available. As seen in the works presented, the
literature is both wide-ranging and relevant to the current study. None of these studies, however, have examined the calculation and use of item discrimination values as approached in the current study. That is, none have evaluated how restricting the calculation of item discrimination values to the scores of examinees near the cut scores of credentialing examinations with limited sample sizes affects item selection, examination reliability, and classification decision consistency. The current research, therefore, represents a unique and valuable contribution to the expansion of knowledge in this important field.

## CHAPTER 3

## METHOD

Three examinations were used to measure the effect restricting scores upon which item discrimination values were calculated to those near cut scores had on the discrimination values themselves, item selection, examination reliability, and classification decision consistency. Detailed information regarding participants, materials used, and data analysis procedures are included in the sections that follow.

## Participants

The participants in this study were examinees who took one of three tests used to credential individuals in health-related professions. As seen in Table 3.1, the number of participants varied according to examination. Utilizing examinations with various examinee population sizes allowed for a closer analysis of how the dependent variables were affected by sample size. The examinations used were also selected because the examinee population size for each is relatively small, reflecting realistic conditions under which many credentialing examinations are developed. In each case, the examinee population size is $N \leq$ 500, thus necessitating the use of classical test theory procedures, as opposed to other approaches, such as item response theory, that would ordinarily require larger sample sizes.

Table 3.1.
Summary of Examination Characteristics

| Examination | Type | Stakes | $N$ | Number of <br> Items | Scoring | Timing |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Examination 1 | C | M | 490 | 175 | D | 8 hours |
| Examination 2 | C | $\mathrm{M}-\mathrm{H}$ | 161 | 200 | D | 4 hours |
| Examination 3 | C | L | 76 | 175 | D | 4 hours |

Note. The following legend explains the symbols used in this table:
$N=$ sample size; Type: $\mathrm{C}=$ certification; Stakes: $\mathrm{L}=$ Low, $\mathrm{M}=$ Medium, $\mathrm{H}=\mathrm{High} ; N=$ number of examinees; Scoring: $\mathrm{D}=$ dichotomous; Timing: number of hours permitted.

## Materials

Three examinations were used in this study. Each examination was used to credential individuals in a health-field profession. Responses to test items were used to answer the research questions. A brief description of each examination used is included in the following sections.

## Examination 1

Examination 1, the largest data set used, included responses from 490 examinees to a 175 -item test used to credential individuals in the environmental health field. To be eligible to take the examination, candidates must hold a bachelor's degree or higher in engineering, chemistry, physics, or the biological or physical sciences. In addition, each candidate must have had at least four years of work experience in the environmental health field. Candidates
were given eight hours to complete the test. The examination is accredited by the American National Standards Institute (ANSI, 2013), which ensures it meets internationally recognized standards pertaining to certification of personnel. The examination, which is offered internationally, is considered to have low to medium stakes, with certification influencing some employment decisions. Descriptive statistics for Examination 1 (as well as for the other examinations used in the study) are included in Table 3.2. In addition, histograms representing total score distributions for Examinations 1, 2, and 3 are included in Figures 3.1, 3.2, and 3.3, respectively.

As seen in Table 3.2, scores for Examination 1 ranged between 46 and 161 with a mean score of 111.42. The distribution of scores was slightly negatively skewed, with a skewness value of -0.44 . The $S D$ was 19.55 and the $S E M$ was 5.79 . Examination reliability, expressed in terms of internal consistency using coefficient alpha, was 0.91 .


Figure 3.1. Histogram of total scores for Examination 1.

Table 3.2.
Descriptive Statistics for Examinations Used

| Examination | $N$ | No. Items | $M$ | $S D$ | SEM | Min | Max | Skewness | Kurtosis | $\alpha$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Examination 1 | 490 | 175 | 111.42 | 19.55 | 5.79 | 46 | 161 | -0.40 | 0.30 | 0.91 |
| Examination 2 | 161 | 200 | 134.27 | 18.23 | 5.84 | 82 | 174 | -0.20 | 0.02 | 0.90 |
| Examination 3 | 76 | 149 | 115.04 | 9.01 | 4.44 | 93 | 134 | -0.13 | -0.35 | 0.76 |

## Examination 2

Examination 2 is also used to credential individuals in a health-related field. The data set included responses from 161 examinees to 200 test items. To be eligible to take the examination, candidates must have at least an associate's degree and must be practicing in the field. Examinees are given four hours to complete the test. The examination is considered to have medium to high stakes, with scores influencing some employment and retention decisions.

Scores for Examination 2 ranged between 82 and 174, with a mean score of 134.27. The score distribution for Examination 2 was also slightly negatively skewed, as evidenced by its skewness value of -0.20 . The $S D$ was 18.23 and the $S E M$ was 5.84 . Examination reliability, expressed in terms of coefficient alpha, was estimated to be 0.90 .


Figure 3.2. Histogram of total scores for Examination 2.

## Examination 3

Examination 3, which yielded the smallest data set used, included responses from 76 examinees to 149 test items. The examination originally contained 150 items, but one item was eliminated from scoring after the examination was administered, resulting in 149 scored items. Examination 3 is used as a credentialing test for registered nurses. The intended purpose of the test is to measure nurses' understanding of diabetes. All examinees were practicing registered nurses. The examination was offered internationally through a network of computer-based testing centers. Candidates were given four hours to complete the examination. The test is considered to have low stakes, with results not impacting hiring or performance reviews.

Scores for Examination 3 ranged between 93 and 134. The mean score was 115.04. As was the case with Examinations 1 and 2, the distribution of test scores for Examination 3 was negatively skewed, with a skewness value of -0.35 . The reliability estimate for Examination 3, again as expressed using coefficient alpha, was 0.76.


Figure 3.3. Histogram of total scores for Examination 3.

## Data Analysis

Many of the procedures described in this study were conducted using jMetrik item analysis software (Version 3.0 for Mac; Meyer, 2013). jMetrik is an open source computer program used to conduct a variety of psychometric analyses, including item discrimination value calculation. It also generates descriptive statistics for examination data sets. In addition, R (version 2.15.2, GUI 1.53; R Core Team, 2012) was used to conduct several procedures and to generate graphical products. The R packages used to complete these operations are included in Table 3.3.

Table 3.3
R Packages Used to Complete Procedures

Package
Functions
base (R Core Team, 2012)
Descriptive statistics
car (Fox \& Weisberg, 2011)
cocron (Diedenhofen, 2013)
graphics (R Core Team, 2012)
psych (Revelle, 2013)
Descriptive statistics
Item analysis
psychometric (Fletcher, 2010)
stats (R Core Team, 2012)
Reliability estimates
Correlation analysis
ANOVA

The three data sets were first screened for missing or miscoded data. As indicated in Table 3.1, all items in each of the examinations used were scored dichotomously; correct responses were scored with a 1 , while incorrect responses were scored with a 0 . Descriptive statistics, as depicted in Table 3.2, were then calculated for each examination. The remainder of this section outlines the procedures used to answer the study's two research questions.

## Research Question 1

The first research question addressed the degree to which item discrimination values are affected when those values are calculated using scores from samples of examinees within ability ranges around real or anticipated cut scores. To begin, unrestricted discrimination values - those calculated in a traditional manner using all examinee responses available were calculated for all examination items. Calculating unrestricted discrimination values was necessary, as these values served as a baseline against which the restricted values were subsequently compared.

The point-biserial statistic, $r_{p b i s}$, was used to estimate item discrimination throughout this study. As described previously, the point-biserial is calculated as follows:

$$
\begin{equation*}
r_{p b i s}=\frac{\mu+-\mu_{x}}{\sigma_{x}} \sqrt{p-q} \tag{3.1}
\end{equation*}
$$

where $\mu_{+}$is the mean total score for those who respond to the item correctly; $\mu_{x}$ is the mean total score for the total group of examinees;
$\sigma_{x}$ is the standard deviation for the total group of examinees;
$p$ is the item's difficulty index; and
$q$ is equal to $(1-p)$ (Crocker \& Algina, 2008).
The point-biserial correlation, which is the observed correlation between item performance and test performance, was selected because it is one of the most commonly used estimates of item discrimination, thus facilitating replications of the procedures used in this study. In addition, previously described research (Oosterhof, 1976; Beuchert \& Mendoza, 1979) found differences between estimates produced by the various discrimination indices to
be insignificant. Use of the point-biserial, as opposed to any other discrimination statistic, therefore, was deemed to have not affected the outcome of the study.

Next, the actual examination cut score, $\mathrm{C}_{\mathrm{X} 1}$, was used as a center point for several groups of test scores upon which restricted point-biserial statistics were subsequently calculated. The actual cut score is defined as the score the examination's governing body or agency approved as the minimum score required in order to pass the test. The restricted point-biserials were calculated in the same manner as the unrestricted values, but were based on fewer test scores. Each group of test scores was centered on the original cut score, but varied in size according to the following increments:

- Original cut score $\pm 1.00 S D$
- Original cut score $\pm .75 S D$
- Original cut score $\pm .50 S D$

Thus, the three groups of test scores upon which the restricted point-biserials were calculated varied in size according to distance from the original cut score. The largest group included all examinee scores within $1.00 S D$ of the original cut score. The next largest group included those within $.75 S D$ of the original cut score. The smallest group included only those scores within $.50 S D$ of the original cut score.

Focusing on smaller groups of test scores around the examination cut score allowed for an increased focus on that portion of the distribution of scores for which classification accuracy is most important. According to Clauser et al. (2006), classification accuracy is the central issue of validity theory with regards to credentialing examinations. Likewise, the Standards (AREA et al., 1999) call for precision "in the vicinity of the passing, or cut, score" (p. 157). It seems important, therefore, to clarify the degree to which test items discriminate
among those examinees who obtain scores near the examination cut score. It was also important to assess how the size of the group of test scores considered affected the item discrimination values. By calculating restricted discrimination indices for the three groups of test scores described earlier, this analysis was made possible.

The process described thus far was repeated at four additional cut score locations. These locations were set at the actual cut score plus $1 \operatorname{SEM}\left(\mathrm{C}_{\mathrm{X} 2}\right)$ and minus $1 \operatorname{SEM}\left(\mathrm{C}_{\mathrm{X} 3}\right)$, and the actual cut score plus $2 \operatorname{SEM}\left(\mathrm{C}_{\mathrm{X} 4}\right)$ and minus $2 \operatorname{SEM}\left(\mathrm{C}_{\mathrm{X} 5}\right)$. These new cut score locations then also served as central points around which item discrimination values for three groups of test scores were calculated: scores $\pm 1.00 S D$, scores $\pm .75 S D$, and scores $\pm .50$ SD. Calculating discrimination values at these four additional cut score locations was important because in many instances the cut scores recommended by the results of a standard setting procedure vary from year to year. Adjustments to examination cut scores are not uncommon, and may be due to changes in the examination itself, or the composition of the standard setting panel.

When discrimination values are calculated using limited groups of examinee test scores around cut scores, the cut score represents a unique location in the distribution of all test scores. A change in the location of cut score represents a different point in that distribution. It is helpful to assess, therefore, the degree to which the location of the cut score affects item discrimination values when those values are calculated using limited samples of examinee test scores. Using the actual cut score, as well as the actual cut score, plus and minus 1 and $2 S E M$, as central points around which groups of scores are used to calculate discrimination values helped to determine how changes in the location of the cut score affected those values.

The procedures used to answer Research Question 1 resulted in the calculation of 16 unique discrimination values for each examination item: 1 unrestricted point-biserial calculated using all available test scores, and 15 restricted point-biserial calculated using three groups of test scores around each of the five cut scores. The location and size of the groups upon which the 15 restricted values were calculated are depicted in Table 3.4

After the 15 sets of restricted point-biserials were calculated for each examination item, the values were compared to their corresponding unrestricted point-biserials. The initial analysis of differences between the restricted and unrestricted values was based on visual comparisons. Each set of restricted values was jointly plotted with the corresponding unrestricted values, allowing for a better understanding of differences and general trends. Next, a series of boxplots were produced. Each boxplot represented one of the 15 conditions under which the restricted point-biserials were calculated. These boxplots were displayed alongside a boxplot representing the unrestricted values. The procedure allowed for a visual comparison of means and the variation of values between the 15 sets of restricted pointbiserials and the unrestricted set.

Actual differences between the restricted and unrestricted values were also calculated. For each condition, the item-level differences were characterized with regard to their direction and magnitude. The mean item discrimination value for the examination was also calculated for each condition. These mean values were then compared to the mean value of the unrestricted point-biserials. These procedures provided further insights into general trends regarding changes in magnitude and direction of the item discrimination values as the conditions were applied.

Table 3.4.
Description of Item Discrimination Values Calculated

Cut score $\quad$ Cut score location $\quad$ Size of group used to calculate point-biserial

| $\mathrm{C}_{\mathrm{X} 1}$ | Actual cut score | $\begin{aligned} & \mathrm{C}_{\mathrm{X} 1} \pm 1.00 S D \\ & \mathrm{C}_{\mathrm{X} 1} \pm .75 S D \\ & \mathrm{C}_{\mathrm{X} 1} \pm .50 S D \end{aligned}$ |
| :---: | :---: | :---: |
| $\mathrm{C}_{\mathrm{X} 2}$ | $\mathrm{C}_{\mathrm{X} 1}+1$ SEM | $\begin{aligned} & \mathrm{C}_{\mathrm{X} 2} \pm 1.00 S D \\ & \mathrm{C}_{\mathrm{X} 2} \pm .75 S D \\ & \mathrm{C}_{\mathrm{X} 2} \pm .50 S D \end{aligned}$ |
| $\mathrm{C}_{\mathrm{X} 3}$ | $\mathrm{C}_{\mathrm{X} 1}-1$ SEM | $\begin{aligned} & \mathrm{C}_{\mathrm{X} 3} \pm 1.00 S D \\ & \mathrm{C}_{\mathrm{X} 3} \pm .75 S D \\ & \mathrm{C}_{\mathrm{X} 3} \pm .50 S D \end{aligned}$ |
| $\mathrm{C}_{\mathrm{X} 4}$ | $\mathrm{C}_{\mathrm{X} 1}+2$ SEM | $\begin{aligned} & \mathrm{C}_{\mathrm{X} 4} \pm 1.00 S D \\ & \mathrm{C}_{\mathrm{X} 4} \pm .75 S D \\ & \mathrm{C}_{\mathrm{X} 4} \pm .50 S D \end{aligned}$ |
| $\mathrm{C}_{\mathrm{X} 5}$ | $\mathrm{C}_{\mathrm{X} 1}-2$ SEM | $\begin{aligned} & \mathrm{C}_{\mathrm{X} 5} \pm 1.00 S D \\ & \mathrm{C}_{\mathrm{X} 5} \pm .75 S D \\ & \mathrm{C}_{\mathrm{X} 5} \pm .50 S D \end{aligned}$ |

Note. Point-biserial statistics were used to estimate item discrimination values.

The procedures described thus far helped to establish a broader understanding of how limiting examinee scores upon which discrimination values are calculated to those within a specified distance of an examination cut score affects those values. To examine the degree to which the differences between restricted an unrestricted values were or were not statistically significant, however, two addition procedures were conducted. First, a correlation analysis of the 16 sets of point-biserials was conducted. The correlation coefficients generated were tested for statistical significance at the $\alpha=.05$ level. A correlation matrix was produced and
is included in the results. The correlation analysis helped to determine where the relationships between the sets of point-biserials were strongest and weakest. This was important as it helped to identify those conditions that resulted in the greatest differences between restricted and unrestricted values.

Next, in order to ascertain the presence of statistically significant differences among the sets of point-biserials, a one-way analysis of variance (ANOVA) procedure was conducted. ANOVA procedures are used to test for differences among the means of two or more samples (Huck, 2008). In this study, item discrimination value was the dependent variable and the conditions under which the values were calculated served as the independent variables. The analysis was used to test for differences among the means of the 15 sets of restricted point-biserials and the unrestricted set. The null hypothesis tested, therefore, was:

$$
\mathrm{H}_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4}=\ldots \mu_{16}
$$

Four assumptions are associated with ANOVA procedures. These include independence, randomness, normality, and homogeneity of variance (Huck, 2008). Because the procedure involved the analysis of the same sets of items under various conditions, the assumption of independence could not be assumed. As a result, a one-way repeated measures ANOVA technique was used to analyze the data. Repeated measures ANOVAs are frequently used when participants or other entities are measured according to some factor over repeated occasions (Huck, 2008).

The purpose of ANOVA using repeated measures is identical to that of those without repeated measures: to test the null hypothesis that the means among the groups of data are
equal. Repeated measures ANOVAs, however, are also subject to the assumption of sphericity. This assumption is satisfied when the population variances, as well as all of the bivariate correlations, are identical (Huck, 2008). When the assumption is not met, the $F$ value produced by the ANOVA will be too large. Mauchly's (1940) test, a procedure frequently used to assess sphericity, was used in this study to evaluate compliance with this assumption. The null hypothesis for the Mauchly test is that the differences in variances between the groups from which the samples were drawn are all equal. This null hypothesis can be expressed in the following terms:

$$
\mathrm{H}_{0}: \sigma_{y 1-y 2}^{2}=\sigma_{y 1-y 3}^{2}=\sigma_{y 2-y 3}^{2} \cdots
$$

Violations of the sphericity assumption were corrected using an approach developed by Greenhouse and Geisser (1959). This Greenhouse-Geisser technique bases the critical $F$ value on the degrees of freedom that would have been appropriate if only two levels of the repeated measure had been used. By doing so, the approach assumes maximum violation of the sphericity assumption, and produces a conservative $F$ statistic (Huck, 2008). Finally, analysis of differences was conducted using Tukey's honest significant difference (HSD) test. All of the procedures associated with the repeated measures ANOVA used in this study were tested at the $\alpha=.05$ level of significance.

In sum, then, the procedures used to characterize the differences between the unrestricted point-biserials and the restricted values included visual comparisons, analysis of differences in discrimination values at the individual item and examination levels, a
correlation analysis, and ANOVA procedures. Collectively, these procedures led to a better understanding of Research Question 1.

## Research Question 2

The second research question addressed the degree to which restricted item discrimination values affects certain key test specifications, including item selection, examination reliability, and classification decision consistency. This section outlines the procedures followed to answer this question.

In order to evaluate how restricted point-biserials affect the aforementioned psychometric aspects of credentialing examinations, each examination used in this study was treated as a test bank from which items were drawn to create two forms of a new 50-item examination. The bank's 50 most discriminating items, as determined by the items' unrestricted point-biserial values, were used to create a pseudo form, Form A. A second pseudo form, Form B, was also created, comprised of the bank's 50 most discriminating items as determined by the items' restricted point-biserial values. The restricted pointbiserials used to select items for Form A were based on examinee scores within $1.00 S D$ of the actual examination cut score $\left(\mathrm{C}_{\mathrm{X} 1}\right)$. The forms were created for each of the examinations used in this study.

After the two forms were created, item selection consistency across the forms was evaluated. Because the criteria by which items were selected for the two forms were different - Form A utilized unrestricted discrimination values and Form B utilized restricted values - it was anticipated that they would likely include different items. Any observed differences in items between the two forms would indicate that the use of restricted point-
biserials resulted in the selection of items that were different than those selected using unrestricted, or traditionally calculated, point-biserials.

Form A and Form B were also used to measure how restricted item discrimination values affected examination reliability. The estimate of examination reliability used in this study was coefficient alpha (Cronbach, 1951). Coefficient alpha was selected because it is a versatile and widely used measure of examination reliability, which requires, unlike other reliability estimates, only a single test administration. Coefficient alpha was calculated for the test variants using the following equation:

$$
\begin{equation*}
\hat{\alpha}=\frac{k}{k-1}\left(1-\frac{\Sigma \hat{\sigma}_{i}^{2}}{\hat{\sigma}_{x}^{2}}\right) \tag{3.2}
\end{equation*}
$$

where $k$ is the number of items on the examination;
$\hat{\sigma}_{i}^{2}$ is the variance of item $i$, and
$\hat{\sigma}_{x}^{2}$ is the total test variance (Crocker \& Algina, 2008).
The coefficients were tested for significant differences at the $\alpha=.05$ level of significance using the cocron package (Diedenhofen, 2013) in R (R Core Team, 2012). The package incorporates earlier work by Alsawalmeh and Feldt (1994) who developed a model by which the null hypothesis of equal coefficient alphas among dependent samples can be tested. This null hypothesis can be stated in the following terms:

$$
\mathrm{H}_{0}=\alpha_{1}-\alpha_{2}=0
$$

Alsawalmeh and Feldt's model can be expressed in the following terms:

$$
\begin{equation*}
W \dot{\sim} \frac{1-\hat{\alpha}_{2}}{1-\hat{\alpha}_{1}}\left(F_{v, v}\right), \quad v=\frac{N-1-7 r_{1,2}^{2}}{1-r_{1,2}^{2}} \tag{3.3}
\end{equation*}
$$

where $\alpha_{1}$ and $\alpha_{2}$ are the coefficient alpha values for the examinations; and $r_{1,2}^{2}$ is the squared correlation coefficient between the examination total scores. It is important to note, however, that the reliability estimates for both forms were calculated using all examinee test scores. Although the items selected for Form B were those that were most discriminating among examinees with total scores within $1.00 S D$ of the original cut score only, using this subset of scores to calculate examination reliability for Form B would have resulted in it being less reliable than Form A in all cases. This is because the number of cases, or $N$ size, directly affects reliability, with, all other things being equal, examinations with more cases generally producing larger reliability coefficients than examinations with fewer cases (Crocker \& Algina, 2008). The selection of items for Form B was based on a substantially smaller number of cases. Because those cases were all within 1.00 SD of the cut score, they were also much more homogenous than the group of scores used to select items for Form A (i.e. all cases). The homogeneity of a group of examinees is an important factor that affects the magnitude of reliability estimates. In general, higher levels of group homogeneity result in lower estimates of reliability because in such situations, total test variance is diminished (Crocker \& Algina, 2008). As documented in Equation 3.2, total test variance is an element in the calculation of coefficient alpha. It was necessary to base all estimates of reliability, therefore, on all test scores available. In
summary, then, although the items found on Form A and Form B were selected using divergent methods for calculating item discrimination, estimates of reliability for both forms were based on all test scores available.

Similar procedures were used to measure the degree to which using restricted item discrimination values as the item selection criterion affected examination classification decision consistency. In this study, Cohen's (1960) $\kappa$ was used to estimate classification decision consistency. As described in Chapter 2, the coefficient can be calculated using the following formula:

$$
\begin{equation*}
\kappa=\frac{P-P_{c}}{1-P_{c}} \tag{3.4}
\end{equation*}
$$

where $P_{c}$, also referred to as the chance consistency, is the probability of a consistent decision, and may be calculated using the following formula:

$$
\begin{equation*}
P_{c}=P_{1 .} P_{.1}=P_{0 .} P_{.0} \tag{3.5}
\end{equation*}
$$

where $P_{1 \text {. }}$ represents the probability of a mastery classification on one form of an examination;
$P_{.1}$ represents a similar probability on the another equivalent form; and $P_{0 \text {. }}$ and $P_{.0}$, which represent misclassifications on the forms.

Coefficient $\kappa$ represents the increase in decision consistency over that expected by chance. The coefficient is equal to 0.00 when there is no increase, and 1.00 when there is maximum increase (Crocker \& Algina, 2008).

The discussion of test forms in the preceding paragraph may be somewhat confusing within the context of the current study. Although the tests created to answer Research Question 2 are referred to as Form A and Form B, they are treated as distinct non-equivalent examinations. Their purpose was to make possible an assessment of how using two distinct methods for calculating item discrimination values affected item selection, reliability, and classification decision consistency. The use of forms within the calculation of Cohen's (1960) $\kappa$ refers to equivalent or parallel forms of the same examination. This does, however, raise the question of how classification decision consistency can be calculated using only a single examination, in this case, Form A and Form B. Modifications to $\kappa$ made by Huynh (1976) allowed for its calculation when only a single examination is available. Using this modification, estimates for classification decision consistency were calculated for Form A and Form B.

The difference in classification decision consistency between the forms was also tested for significant differences. Using a formula developed by Donner, Shoukri, Klar, and Bartfay (2000), coefficient $\kappa$ values for two dependent groups can be tested for significant differences. To conduct this test, the following formula is used:

$$
\begin{equation*}
Z_{V D}=\frac{\hat{\kappa}_{1}-\hat{\kappa}_{2}}{\left[\operatorname{var}\left(\hat{\kappa}_{1}\right)+\operatorname{var}\left(\hat{\kappa}_{2}\right)-2 \operatorname{cov}\left(\hat{\kappa}_{1}, \hat{\kappa}_{2}\right)\right]} \tag{3.6}
\end{equation*}
$$

where $\kappa_{1}$ and $\kappa_{2}$ represent the classification decision consistency ratings for the two examinations.

The formula tests the null hypothesis that the difference between the coefficients is zero:

$$
\mathrm{H}_{0}=\kappa_{1}-\kappa_{2}=0
$$

The test for differences between $\kappa$ coefficients was also tested at the $\alpha=.05$ level of significance.

The procedures described in this section were used to answer Research Question 2. That is, they led to a determination of the degree to which using restricted item discrimination values as a criterion for selecting examination items affected item selection, examination reliability, and classification decision consistency. The procedures were conducted for each of the three examination used in this study. By doing so, greater insight into the role sample size plays in these considerations was possible.

The three variables examined in Research Question 2, namely item selection, examination reliability, and classification decision consistency, are critical elements in the gathering of validity evidence for credentialing examinations. Understanding how examinations with items that were selected using restricted discrimination values affects these specifications, therefore, becomes a valuable endeavor. Again, according to the Standards (AERA et al., 1999), examinations used for credentialing individuals "may not need to be precise for those who clearly pass or clearly fail," as "sometimes a test used in credentialing is designed to be precise only in the vicinity of the cut score" (p. 157).

## CHAPTER 4

## RESULTS

Utilizing examinee scores from three examinations used to credential individuals in health-related professions, the study examined the degree to which limiting scores upon which item discrimination values are calculated to those near actual or anticipated cut scores affected the item discrimination values, item selection, examination reliability, and classification decision consistency. An initial analysis of the data used found that there were no missing or miscoded responses. Procedures were then followed to answer the study's two research questions. Research Question 1 addressed the effect on item discrimination values. Research Question 2 examined the effect on item selection, examination reliability, and classification decision consistency. The results for each research question are addressed in the sections that follow.

## Research Question 1

Research Question 1 examined the effect limiting scores upon which item discrimination values are calculated to those near actual or anticipated cut scores had on item discrimination values themselves. To accomplish this, the unrestricted point-biserial statistic, $r_{p b i s}$, was calculated for each examination item. Restricted point-biserials were then calculated for all items under each of 15 conditions. These conditions, as well as the cut scores and examinee scores that were considered under each condition, are listed in Table 4.1. Differences between the restricted and unrestricted values were then analyzed. The results for each examination are presented individually in the sections that follow.

Table 4.1
Conditions for the Calculation of Restricted Point-biserials

| Stat/Cut Score | Group | Examination 1 | Examination 2 | Examination 3 |
| :---: | :---: | :---: | :---: | :---: |
| No. Items |  | 175 | 200 | 149 |
| $S D$ |  | 19.55 | 18.23 | 9.01 |
| SEM |  | 5.79 | 5.84 | 4.44 |
| $\mathrm{C}_{\mathrm{X} 1}$ |  | 115 | 128 | 106 |
| (Actual $\mathrm{C}_{\mathrm{X}}$ ) | +/-1.00 SD | 96-134 | 110-146 | 97-115 |
|  | +/-0.75SD | 101-129 | 115-141 | 100-112 |
|  | +/-0.50 SD | 106-124 | 119-137 | 102-110 |
| $\mathrm{C}_{\mathrm{X} 2}$ |  | 121 | 134 | 111 |
| (+ 1 SEM) | +/-1.00 SD | 102-140 | 116-152 | 102-119 |
|  | +/-0.75SD | 107-135 | 121-147 | 104-117 |
|  | +/- 0.50 SD | 112-130 | 125-142 | 106-114 |
| $\mathrm{C}_{\mathrm{X} 3}$ |  | 110 | 123 | 102 |
| (-1 SEM) | +/-1.00 SD | 90-128 | 104-140 | 93-110 |
|  | +/-0.75SD | 95-123 | 109-135 | 95-108 |
|  | +/-0.50 SD | 100-118 | 114-131 | 98-106 |
| $\mathrm{C}_{\mathrm{X} 4}$ |  | 127 | 140 | 115 |
| (+2SEM) | +/-1.00 SD | 108-146 | 122-157 | 106-123 |
|  | +/-0.75SD | 112-141 | 127-153 | 109-121 |
|  | +/- 0.50 SD | 117-136 | 131-148 | 111-119 |
| $\mathrm{C}_{\mathrm{x} 5}$ |  | 104 | 117 | 98 |
| (-2 SEM) | +/-1.00 SD | 84-122 | 99-134 | 89-106 |
|  | +/-0.75SD | 89-118 | 103-129 | 91-103 |
|  | +/-0.50 SD | 94-113 | 108-125 | 93-101 |

Note. Ranges indicate examinee scores considered in the calculation of item discrimination values at each cut score location.

## Examination 1

Examination 1, which in terms of sample size was the largest test used in the study, consisted of 490 examinee responses to 175 items. The initial analyses of differences between restricted and unrestricted discrimination values were based on visual comparisons. First, the unrestricted and restricted discrimination values for each of the five cut scores used were plotted. Figure 4.1 displays the unrestricted values plotted along side the restricted values calculated using scores within $1.00 S D, .75 S D$, and $.50 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ (the actual cut score). Figures 4.2, 4.3, 4.4, and 4.5 show the same information for $\mathrm{C}_{\mathrm{X} 2}, \mathrm{C}_{\mathrm{X} 3}, \mathrm{C}_{\mathrm{X} 4}$, and $\mathrm{C}_{\mathrm{X} 5}$, respectively.

The plots displayed in Figures 4.1 through 4.5 suggest that unrestricted item discrimination values are, in general, larger than their corresponding restricted values. This appeared to be the case at each of the five cut score locations examined. As the size of the groups of test scores used to calculate the restricted discrimination values became smaller, the discrimination values themselves were generally smaller. For example, item discrimination values calculated using only scores within $.50 S D$ of the cut scores appeared to be smaller than values calculated using scores within $.75 S D$ and $1.00 S D$ of the cut scores.

To further understand general trends associated with the differences in the discrimination values calculated, a boxplot was also produced. As seen in Figure 4.6, the distribution of unrestricted point-biserials included larger discrimination values than any other group of values. As suggested by the plots in Figures 4.1 through 4.5, the boxplot also indicates that the values decrease as the size of the group used to calculate them becomes smaller.


Figure 4.1. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ for Examination 1.


Figure 4.2. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ for Examination 1.


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Figure 4.3. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 3}$ for Examination 1.

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Figure 4.4. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 4}$ for Examination 1.


Figure 4.5. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}$ for Examination 1.


Figure 4.6. Boxplot highlighting distribution of item discrimination values under each condition for Examination 1.

To gain additional insights into the discrimination values associated with each of the 15 conditions, several descriptive statistics were calculated. These statistics are included in Table 4.2. The descriptive statistics support the trends observed in both the data plots (Figures 1.4 through 4.5) and the boxplot (Figure 4.6). The mean item discrimination value for the unrestricted group was $M=0.25(S D=0.09)$, which was larger than the mean value for any other group. In addition, at each cut score location, the mean item discrimination value decreased as the size of the group used to calculate the value became smaller. For example, when scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ were used, the mean item discrimination value was $M=0.13(S D=0.08)$. When scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ were used, the mean discrimination value was $M=0.10(S D=0.07)$. The mean value was $M=0.07(S D=0.08)$ when scores within $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ were considered. As displayed in the table, this was consistent for each cut score location. The table also indicates the minimum and maximum
value for each condition, as well as the range of values. In addition, it includes the number of test scores that were considered for each of the conditions examined. The set of restricted values that included the most examinee scores was that which considered scores within 1.00 $S D$ of $\mathrm{C}_{\mathrm{X} 3}$. Calculations for this set included 344 scores. The set with the fewest scores included those within $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}$. For this set, only 152 test scores were used to calculate the point-biserials.

To further understand the nature of the relationships between the sets of discrimination values, a correlation table was produced. The correlation table is included in Table 4.3. As seen in the table, 102 of the 120 correlation coefficients calculated were significantly different than zero at the $\alpha=.05$ level of significance. Of particular interest were the relationships between the unrestricted discrimination values and each of the 15 restricted sets of values. In each case, the correlation coefficient between these sets of values was found to be significantly different than zero. The strongest correlation observed was between discrimination values calculated using examinee scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ and values calculated using scores between $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 1}, r(173)=.84, p<.001$. The weakest correlation, .02 , was found between four sets of values. In each of these instances, the correlation was found to be not significantly different than zero. The majority of correlation coefficients expressed positive relationships between the sets of values, with only 12 of the 120 coefficients expressing negative relationships.

Table 4.2
Descriptive Statistics of Discrimination Values Calculated for Examination 1

| Cut Score | Group | M | $S D$ | Min | Max | Range | Skew | $n$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Unrestricted |  | 0.25 | 0.09 | -0.05 | 0.49 | 0.54 | 0.11 | 490 |
| $\mathrm{C}_{\mathrm{X} 1}$ |  |  |  |  |  |  |  |  |
|  | +/- 1.00 SD | 0.13 | 0.08 | -0.05 | 0.35 | 0.40 | 0.35 | 339 |
|  | +/- 0.75SD | 0.10 | 0.07 | -0.07 | 0.31 | 0.38 | 0.33 | 287 |
|  | +/- 0.50 SD | 0.07 | 0.08 | -0.12 | 0.25 | 0.37 | 0.14 | 198 |
| $\mathrm{C}_{\mathrm{X} 2}$ |  |  |  |  |  |  |  |  |
|  | +/- 1.00 SD | 0.12 | 0.07 | -0.04 | 0.34 | 0.38 | 0.37 | 317 |
|  | +/- 0.75SD | 0.10 | 0.07 | -0.11 | 0.27 | 0.39 | 0.03 | 267 |
|  | +/- 0.50 SD | 0.07 | 0.08 | -0.14 | 0.32 | 0.46 | 0.12 | 200 |
| $\mathrm{C}_{\mathrm{X} 3}$ |  |  |  |  |  |  |  |  |
|  | +/- 1.00 SD | 0.14 | 0.08 | -0.06 | 0.31 | 0.37 | 0.11 | 344 |
|  | +/- 0.75SD | 0.10 | 0.08 | -0.10 | 0.30 | 0.40 | 0.02 | 255 |
|  | +/- 0.50 SD | 0.07 | 0.09 | -0.21 | 0.32 | 0.53 | -0.05 | 171 |
| $\mathrm{C}_{\mathrm{X} 4}$ |  |  |  |  |  |  |  |  |
|  | +/- 1.00 SD | 0.13 | 0.07 | -0.05 | 0.31 | 0.36 | 0.17 | 292 |
|  | +/- 0.75SD | 0.10 | 0.08 | -0.08 | 0.31 | 0.39 | 0.22 | 240 |
|  | +/- 0.50 SD | 0.07 | 0.09 | -0.15 | 0.30 | 0.45 | -0.13 | 169 |
| $\begin{array}{llllllll}\mathrm{CX5}^{2} & +/-1.00 S D & 0.14 & 0.08 & -0.07 & 0.37 & 0.44 & 0.13\end{array}$ |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
|  | +/- 0.75SD | 0.11 | 0.08 | -0.12 | 0.32 | 0.44 | 0.06 | 239 |
|  | +/- 0.50 SD | 0.07 | 0.10 | -0.22 | 0.33 | 0.55 | -0.11 | 152 |

Table 4.3
Correlation Matrix of Item Discrimination Values Calculated for Examination 1


Note. * $p=<.05$.

The final procedure performed for each examination with respect to Research Question 1 was the one-way, repeated measures ANOVA. This was used to test for differences among the means of the unrestricted discrimination values and the 15 sets of restricted values. In the ANOVA conducted, the discrimination values served as the dependent variable and the conditions under which those values were calculated served as the independent variables, or groups. The purpose of the analysis was to test the null hypothesis that the group mean discrimination values were equal.

The analysis began with an evaluation of ANOVA assumptions. The discrimination values, which served as the dependent variable, were not independent. Each group in the ANOVA consisted of the same subjects, or in this case, examination items, tested under different conditions. Because of this, a repeated measures ANOVA approach was taken. Normality was assessed using two procedures. First, the Shapiro-Wilk normality test (R Core Team, 2012) was performed. The procedure tests the null hypothesis that the data are normally distributed. The results indicated that all but one of the sets of data was normally distributed. The set for which the null hypothesis was rejected included values calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 1}, \mathrm{~W}=0.98, p=.04$. Assessing skewness values was the second test for normality. As seen in Table 4.2, the skewness value for each of the sets of data fell between -1.00 and 1.00. According to Huck (2008), skewness values that fall within this range are typically considered to approximate a normal distribution. Based on this criterion, the data were deemed to have sufficiently satisfied the assumption of normality.

Because the procedure involved a repeated measures approach, the data also needed to satisfy the assumption of sphericity. In a traditional ANOVA procedure, that is, one that does not involve repeated measures, sphericity is not required, but rather the assumption of
homogeneity of variance is tested. Sphericity is satisfied when the variances and bivariate correlations among the sets of data are equal (Huck, 2008). This assumption was tested using Mauchly's (1940) procedure, which tests the hull hypothesis that the variances and bivariate correlations among the groups are equal. The results of Mauchly's test, $W<.001, p$ $<.001$, indicated that the assumption of sphericity was not met. When the assumption is not met, the $F$-statistic will be positively biased and, therefore, the risk of Type I error increases. The Greenhouse-Geisser (1959) correction, estimated to be $\varepsilon=0.28$, was applied to the degrees of freedom in order to obtain a valid critical $F$-value. The results of the ANOVA, to include the Greenhouse-Geisser correction for sphericity, are included in Table 4.4.

## Table 4.4

Results of the One-way Repeated Measures ANOVA for Examination 1

| Source | Model | $S S$ | $d f$ | $M S$ | $F$ | $p$ |
| :--- | :--- | :---: | :---: | :---: | :---: | :---: |
| Condition | Sphericity Assumed | 5.265 | 15.000 | 0.351 | 94.539 | $<.001^{*}$ |
|  | Greenhouse-Geisser <br> Corrected | 5.265 | 4.200 | 1.254 | 94.539 | $<.001^{*}$ |
| Error | Sphericity Assumed | 9.690 | 2610.000 | 0.004 |  |  |
|  | Greenhouse-Geisser <br> Corrected | 9.690 | 730.800 | 0.013 |  |  |
|  |  |  |  |  |  |  |

Note. The Greenhouse-Geisser correction, estimated at $\varepsilon=0.28$, was used to correct for the violation of the sphericity assumption. $*=$ Significant at $\alpha=.05$ level of significance.

As observed in Table 4.4, the results of the corrected ANOVA indicated that differences between the sets of values were significantly greater than zero, $F(4.2,730.8)=$ $94.539, p<.001$. As such, the null hypothesis of equal means among the groups of discrimination values was rejected. To assess where differences existed, Tukey's honest significant difference (HSD) test was conducted. Although Tukey's HSD was utilized to identify significant differences among all groups, the primary objective was to identify differences between each of the restricted groups and the unrestricted group. The results of this analysis are included in Table 4.5. As seen in Table 4.5, a significant mean difference was observed between the unrestricted set of discrimination values and each of the restricted sets when tested at the $\alpha=.05$ level of significance. In each instance, the mean unrestricted value was larger than the mean restricted value.

Table 4.5.
Results of Tukey's HSD Test for Examination 1-Unrestricted vs. Restricted Values

| Condition | Condition |  | Mean difference | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | Cut score | Group size |  |  |
| Unrestricted | $\mathrm{C}_{\mathrm{X} 1}$ | +/-1.00 SD | 0.119 | <.001* |
|  |  | +/- 0.75SD | 0.143 | <.001* |
|  |  | +/- 0.50 SD | 0.177 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 2}$ | +/-1.00 SD | 0.123 | <.001* |
|  |  | +/- 0.75SD | 0.149 | <.001* |
|  |  | +/- 0.50 SD | 0.177 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 3}$ | +/-1.00 SD | 0.101 | $<.001 *$ |
|  |  | +/- 0.75SD | 0.149 | <.001* |
|  |  | +/- 0.50 SD | 0.182 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 4}$ | +/- 1.00 SD | 0.120 | <.001* |
|  |  | +/- 0.75SD | 0.148 | <.001* |
|  |  | +/- 0.50 SD | 0.178 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 5}$ | +/-1.00 SD | 0.113 | <.001* |
|  |  | +/- 0.75SD | 0.140 | <.001* |
|  |  | +/- 0.50 SD | 0.180 | <.001* |

Note. ${ }^{*}=$ Significant at the $\alpha=.05$ level of significance.

## Examination 2

Examination 2, according to sample size, was the second largest test used in the study. Responses from 161 examinees to the 200 -item examination were included in the analysis. The procedures conducted with respect to Examination 1 were also performed for Examination 2. In Figures 4.7 through 4.11, the unrestricted discrimination values are plotted with their corresponding restricted point-biserials for each of the five cut score locations. As was the case with Examination 1, the plots revealed that the unrestricted discrimination values were, in general, larger than the restricted values. The differences between the restricted and unrestricted values, however, appeared to be, at least visually, less stark than was the case with the differences observed in Examination 1. Results of the boxplot, included in Figure 4.12, also suggested that the unrestricted values were larger than the restricted values. In addition, at each cut score location, the mean discrimination value decreased in magnitude as the group of scores used to calculate the values grew smaller in size.

Descriptive statistics for the sets of discrimination values are included in Table 4.6. As seen in the table, the largest mean discrimination value was associated with the unrestricted set, $M=0.22(S D=0.12)$. As was the case with Examination 1, the mean discrimination value decreased as the size of the group of scores considered decreased in size. The smallest mean discrimination value observed was associated with the set of values calculated using scores within $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}, M=0.05(S D=0.18)$. This set included 37 examinee scores, which was fewer than for any other set. Values calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ considered 115 scores, which was more than any other set of restricted values.


Figure 4.7. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ for Examination 2.


Figure 4.8. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ for Examination 2.


Figure 4.9. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 3}$ for Examination 2.


Figure 4.10. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 4}$ for Examination 2.


Figure 4.11. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}$ for Examination 2.


Figure 4.12. Boxplot highlighting distribution of item discrimination values under each condition for Examination 2.

Table 4.6
Descriptive Statistics of Discrimination Values Calculated for Examination 2

| Cut Score | Group | M | $S D$ | Min | Max | Range | Skew | $n$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Unrestricted |  | 0.22 | 0.12 | -0.21 | 0.50 | 0.71 | -0.67 | 161 |
| $\mathrm{C}_{\mathrm{X} 1}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.12 | 0.12 | -0.20 | 0.36 | 0.56 | -0.17 | 111 |
|  | +/- 0.75SD | 0.09 | 0.12 | -0.27 | 0.44 | 0.71 | 0.01 | 85 |
|  | +/-0.50 SD | 0.06 | 0.13 | -0.30 | 0.40 | 0.70 | 0.06 | 67 |
| $\mathrm{C}_{\mathrm{X} 2}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.12 | 0.11 | -0.20 | 0.43 | 0.63 | -0.18 | 115 |
|  | +/- 0.75SD | 0.10 | 0.12 | -0.25 | 0.36 | 0.61 | -0.25 | 96 |
|  | +/- 0.50 SD | 0.06 | 0.13 | -0.35 | 0.33 | 0.69 | -0.29 | 64 |
| $\mathrm{C}_{\mathrm{X} 3}$ |  |  |  |  |  |  |  |  |
|  | +/- 1.00 SD | 0.11 | 0.12 | -0.29 | 0.44 | 0.73 | -0.21 | 93 |
|  | +/- 0.75SD | 0.08 | 0.13 | -0.42 | 0.42 | 0.84 | -0.06 | 76 |
|  | +/- 0.50 SD | 0.06 | 0.14 | -0.34 | 0.39 | 0.73 | 0.05 | 57 |
| $\mathrm{C}_{\mathrm{X} 4}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.12 | 0.11 | -0.19 | 0.39 | 0.58 | -0.19 | 109 |
|  | +/- 0.75SD | 0.09 | 0.11 | -0.32 | 0.41 | 0.72 | -0.19 | 89 |
|  | +/- 0.50 SD | 0.07 | 0.13 | -0.30 | 0.45 | 0.74 | -0.03 | 57 |
| $\mathrm{C}_{\mathrm{X} 5}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.10 | 0.12 | -0.26 | 0.44 | 0.71 | -0.11 | 82 |
|  | +/- 0.75SD | 0.09 | 0.13 | -0.27 | 0.41 | 0.68 | -0.12 | 60 |
|  | +/-0.50 SD | 0.05 | 0.18 | -0.38 | 0.62 | 1.00 | -0.02 | 37 |

The correlation matrix developed for Examination 2 is included in Table 4.7. Among the 120 correlation coefficients calculated, 99 were found to be significantly different than zero at the $\alpha=0.05$ level of significance. The strongest correlation observed was between values calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ and values calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 4}, r(198)=0.83, p<.001$. The weakest correlation observed was between values calculated using scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ and values calculated using scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 5}$. The correlation between these sets of values was not significantly different than zero. The majority of coefficients were positive, with only 24 of the 120 expressing negative relationships.

One-way repeated measures ANOVA was also conducted for the sets of discrimination values associated with Examination 2. Tests for normality were similar to those conducted for Examination 1. The Shapiro-Wilk normality test (R Core Team, 2012) results indicated that the assumption of normality was not satisfied for one of the sets of values. The null hypothesis of normality was rejected for the set of unrestricted discrimination values, $\mathrm{W}=0.97, p<.001$. As seen in Table 4.6, however, the skewness values for each set of point-biserials ranged between -1.00 and 1.00, indicating the data approximated normality. The data were, therefore, deemed suitable for further analysis.

Mauchly's (1940) test indicated that the assumption of sphericity had not been satisfied, $W<0.001, p<.001$. The Greenhouse-Geisser (1959) correction, $\varepsilon=0.28$, was once again applied to the degrees of freedom in order to combat the effects associated with the violation of the sphericity assumption. The results of the ANOVA are included in Table 4.8.

Table 4.7
Correlation Matrix of Item Discrimination Values Calculated for Examination 2


Note. $* p=<.05$.

Table 4.8
Results of the One-way Repeated Measures ANOVA for Examination 2

| Source | Model | $S S$ | $d f$ | $M S$ | $F$ | $p$ |
| :--- | :--- | :---: | :---: | :---: | :---: | :---: |
| Condition | Sphericity Assumed | 4.571 | 15.000 | 0.305 | 27.046 | $<.001^{*}$ |
|  | Greenhouse-Geisser <br> Corrected | 4.571 | 4.200 | 1.088 | 27.046 | $<.001^{*}$ |
| Error | Sphericity Assumed | 33.636 | 2985.000 | 0.011 |  |  |
|  | Greenhouse-Geisser <br> Corrected | 33.636 | 835.800 | 0.040 |  |  |
|  |  |  |  |  |  |  |

Note. The Greenhouse-Geisser correction, estimated at $\varepsilon=0.28$, was used to correct for the violation of the sphericity assumption. $*=$ Significant at $\alpha=.05$ level of significance.

As seen in the table, the sphericity-corrected ANOVA produced a significant result, $F(4.20,835.80)=27.046, p<.001$, indicating that differences among the group means existed. Tukey's HSD was used to identify those differences. The results of this test are included in Table 4.9.

Table 4.9
Results of Tukey's HSD Test for Examination 2 - Unrestricted vs. Restricted Values

| Condition | Condition |  | Mean difference | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | Cut score | Group size |  |  |
| Unrestricted | $\mathrm{C}_{\mathrm{X} 1}$ | +/-1.00 SD | 0.096 | $<.001 *$ |
|  |  | +/- 0.75SD | 0.128 | <.001* |
|  |  | +/-0.50 SD | 0.153 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 2}$ | +/-1.00 SD |  | $<.001 *$ |
|  |  | +/- 0.75SD | $0.116$ | <.001* |
|  |  | +/- 0.50 SD | 0.152 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 3}$ | +/-1.00 SD | 0.108 | $<.001 *$ |
|  |  | +/-0.75SD | 0.133 | <.001* |
|  |  | +/-0.50 SD | 0.155 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 4}$ | +/-1.00 SD | 0.098 | $<.001 *$ |
|  |  | $+/-0.75 S D$ | 0.121 | <.001* |
|  |  | +/-0.50 SD | 0.148 | < .001* |
|  | $\mathrm{C}_{\mathrm{X} 5}$ | +/-1.00 SD | 0.111 | $<.001 *$ |
|  |  | +/-0.75SD | 0.130 | <.001* |
|  |  | +/- 0.50 SD | 0.165 | <.001* |

Note. ${ }^{*}=$ Significant at the $\alpha=.05$ level of significance.

As was the case with Examination 1, significant differences were identified between the unrestricted set of values and each of the restricted sets. In each instance, the unrestricted mean discrimination value was greater than the unrestricted mean values.

## Examination 3

Examination 3, which was the smallest test used in this study, consisted of 149 items. Responses from 76 examinees were included in the analysis. The same procedures described for Examination 1 and Examination 2 were conducted for Examination 3. The plots depicting the location of unrestricted and restricted discrimination values for each cut score considered are included in Figures 4.13 through 4.17. Unlike the plots highlighting the same information for Examinations 1 and 2, the plots for Examination 3 indicated no discernable visual relationship between the unrestricted and restricted point-biserials. The restricted discrimination values, in particular, appeared to vary greatly when compared to those observed for the previous examinations.

The boxplot, included in Figure 18, indicated a somewhat more familiar pattern, with the mean discrimination value for the unrestricted set being larger than the mean values for the restricted sets. As was the case with the previous examinations, the mean discrimination value for each set appeared to decrease as the size of the group of scores considered decreased.

Descriptive statistics for the values calculated are included in Table 4.10. As the table indicates, the largest mean discrimination value was that associated with the unrestricted set, $M=0.16(S D=0.13)$. This was the lowest unrestricted mean discrimination value among all of the examinations considered in the study. The largest mean discrimination value among the restricted sets represented point-biserials calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 4}$. This set considered 52 examinee scores, which was more than any other restricted set. The set with the fewest number of scores included those within 0.50 $S D$ of $\mathrm{C}_{\mathrm{X} 5}$. Only six scores were included in the calculations for this set.


Figure 4.13. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ for Examination 3.


Figure 4.14. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ for Examination 3.


Figure 4.15. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 3}$ for Examination 3.


Figure 4.16. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 4}$ for Examination 3.


Figure 4.17. Plots of unrestricted and restricted item discrimination values based on scores within $1.00 S D, 0.75 S D$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}$ for Examination 3.


Figure 4.18. Boxplot highlighting distribution of item discrimination values under each condition for Examination 3.

Table 4.10
Descriptive Statistics of Discrimination Values Calculated for Examination 3

| Cut Score | Group | M | $S D$ | Min | Max | Range | Skew | $n$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Unrestricte |  | 0.16 | 0.13 | -0.11 | 0.46 | 0.58 | -0.02 | 76 |
| $\mathrm{C}_{\mathrm{X} 1}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.08 | 0.17 | -0.29 | 0.57 | 0.86 | 0.34 | 37 |
|  | +/-0.75SD | 0.06 | 0.19 | -0.42 | 0.52 | 0.94 | 0.09 | 25 |
|  | +/- 0.50 SD | 0.04 | 0.23 | -0.44 | 0.70 | 1.14 | 0.30 | 17 |
| $\mathrm{C}_{\mathrm{X} 2}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.08 | 0.15 | -0.36 | 0.49 | 0.84 | 0.18 | 43 |
|  | +/- 0.75SD | 0.06 | 0.17 | -0.42 | 0.49 | 0.90 | 0.04 | 36 |
|  | +/- 0.50 SD | 0.05 | 0.21 | -0.52 | 0.54 | 1.06 | -0.12 | 25 |
| $\mathrm{C}_{\mathrm{X} 3}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.08 | 0.19 | -0.37 | 0.69 | 1.06 | 0.36 | 23 |
|  | +/- 0.75SD | 0.06 | 0.20 | -0.39 | 0.55 | 0.94 | 0.37 | 20 |
|  | +/-0.50 SD | 0.04 | 0.29 | -0.62 | 0.86 | 1.47 | 0.24 | 11 |
| $\mathrm{C}_{\mathrm{X} 4}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.09 | 0.13 | -0.21 | 0.37 | 0.58 | -0.10 | 52 |
|  | +/- 0.75SD | 0.06 | 0.16 | -0.35 | 0.50 | 0.85 | -0.02 | 33 |
|  | +/- 0.50 SD | 0.05 | 0.19 | -0.41 | 0.55 | 0.96 | -0.08 | 26 |
| $\mathrm{C}_{\mathrm{X} 5}$ |  |  |  |  |  |  |  |  |
|  | +/-1.00 SD | 0.06 | 0.28 | -0.62 | 0.86 | 1.48 | 0.25 | 13 |
|  | +/- 0.75SD | 0.05 | 0.34 | -0.77 | 0.83 | 1.60 | 0.16 | 8 |
|  | +/- 0.50 SD | 0.04 | 0.42 | -0.88 | 0.88 | 1.76 | 0.29 | 6 |

The correlation matrix for Examination 3 is included in Table 4.11. Unlike the previous examinations considered, far fewer correlations were significant for Examination 3. Among the 120 correlation coefficients calculated, only 78 were found to be significantly different than zero at the $\alpha=0.05$ level of significance. In addition, 42 coefficients expressed negative relationships, which was more than both Examination 1 (12) and Examination 2 (24). The strongest correlation existed between the set of values calculated using scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 2}$ and the set that included scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 2}, r(147)=0.79$, $p<.001$.

One-way repeated measures ANOVA was also conducted with respect to Examination 3. Unlike the previous examinations, however, the Shapiro-Wilk normality test (R Core Team, 2012) identified several groups of data for which the null hypothesis of normality was rejected. The problematic groups were those calculated using scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 3}, 0.50 S D$ of $\mathrm{C}_{\mathrm{X} 3}, 0.75 S D$ of $\mathrm{C}_{\mathrm{X} 5}$, and $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 5}$. An analysis of group skewness values, however, indicated that each fell within -1.00 to 1.00 . These values are included in Table 4.10. Once again, therefore, the decision was made to proceed with the ANOVA.

The assumption of sphericity was also violated, with Mauchly's (1940) test producing a significant result, $W<0.001, p<.001$. The Greenhouse-Geisser (1959) correction, $\varepsilon=$ 0.28 , was applied to the degrees of freedom to correct the positively biased $F$-statistic. As seen in Table 4.12, the results of the ANOVA were significant, $F(4.20,621.60)=2.989, p=$ .016. The results indicated the presence of a significant difference between at least two of the group means.

Table 4.11
Correlation Matrix of Item Discrimination Values Calculated for Examination 3


Note. $* p=<.05$.

Table 4.12
Results of the One-way Repeated Measures ANOVA for Examination 3

| Source | Model | $S S$ | $d f$ | $M S$ | $F$ | $p$ |
| :--- | :--- | :--- | :---: | :--- | :---: | :---: |
| Condition | Sphericity Assumed | 2.104 | 15.000 | 0.140 | 2.989 | $<.001^{*}$ |
|  | Greenhouse-Geisser <br> Corrected | 2.104 | 4.200 | 0.501 | 2.989 | $.016^{*}$ |
| Error | Sphericity Assumed | 104.188 | 2220.00 | 0.047 |  |  |
|  | Greenhouse-Geisser <br> Corrected | 104.188 | 621.600 | 0.168 |  |  |
|  |  |  |  |  |  |  |

Note. The Greenhouse-Geisser correction, estimated at $\varepsilon=0.28$, was used to correct for the violation of the sphericity assumption. * $=$ Significant at $\alpha=.05$ level of significance.

The results of Tukey's HSD test are included in Table 4.13. Unlike Examinations 1 and 2 , significant differences were not observed between the unrestricted set of values and each of the restricted groups. Significant differences were not observed between the unrestricted group and four of the restricted groups. The groups between which a significant difference with the unrestricted set was not observed included values calculated using scores with 1.00 $S D$ of $\mathrm{C}_{\mathrm{X} 1}, 1.00 S D$ of $\mathrm{C}_{\mathrm{X} 2}, 1.00 S D$ of $\mathrm{C}_{\mathrm{X} 3}$, and $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 4}$. Significant differences between the unrestricted group and each of the other sets, however, were identified, with the mean unrestricted discrimination value being greater than the mean values of the restricted groups.

Table 4.13
Results of Tukey's HSD Test for Examination 3 - Unrestricted vs. Restricted Values

| Condition | Condition |  | Mean difference | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | Cut score | Group size |  |  |
| Unrestricted | $\mathrm{C}_{\mathrm{X} 1}$ | +/-1.00 SD | 0.087 | . 081 |
|  |  | +/-0.75SD | 0.105 | .008* |
|  |  | +/- 0.50 SD | 0.128 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 2}$ | +/- 1.00 SD | 0.081 | . 148 |
|  |  | +/-0.75SD | 0.100 | .015* |
|  |  | +/- 0.50 SD | 0.119 | <.001* |
|  | $\mathrm{C}_{\mathrm{X} 3}$ | +/- 1.00 SD | 0.088 | . 068 |
|  |  | +/- 0.75SD | 0.102 | .011* |
|  |  | +/-0.50 SD | 0.125 | $<.001 *$ |
|  | $\mathrm{C}_{\mathrm{X} 4}$ | +/- 1.00 SD | 0.072 | . 326 |
|  |  | +/- 0.75SD | 0.103 | .010* |
|  |  | +/- 0.50 SD | 0.111 | .003* |
|  | $\mathrm{C}_{\mathrm{X} 5}$ | +/-1.00 SD | 0.104 | .009* |
|  |  | +/- 0.75SD | 0.113 | .002* |
|  |  | +/- 0.50 SD | 0.124 | <.001* |

Note. ${ }^{*}=$ Significant at the $\alpha=.05$ level of significance.

## Summary - Research Question 1

For Research Question 1, results of the analysis based on visual comparisons suggested that at each cut score location, the unrestricted set of discrimination values was, on average, greater in value than the restricted sets. This was evident in both the plots and boxplots. It was particularly true for Examinations 1 and 2, but less apparent for Examination 3, which, according to sample size, was the smallest examination used in the study. The results of the ANOVA supported this initial assessment, indicating that for Examinations 1 and 2, the difference between the unrestricted set of discrimination values and the sets of restricted values was statistically significant. In the case of Examination 3, the difference between the unrestricted set of discrimination values and 11 of the 15 restricted sets was significant. Where significant differences were found, the mean unrestricted point-biserial was larger than the mean restricted item discrimination value.

## Research Question 2

Research Question 2 examined the degree to which using restricted item discrimination values affected item selection, examination reliability, and classification decision consistency. The results of the procedures conducted to answer this research question for each examination are included in the sections that follow.

## Examination 1

To examine the effect restricted item discrimination values had on item selection, two 50-item forms were created. Form A included the 50 most discriminating items from Examination 1, using unrestricted point-biserials as the criterion for selection. Form B included the 50 most discriminating items, using restricted point-biserials as the selection criterion. The set of restricted values used to create Form B was based on scores within 1.00 $S D$ of $\mathrm{C}_{\mathrm{X} 1}$. Descriptive statistics for the Examination 1 forms, as well as for those associated with the other examinations used in the study, are included in Table 4.14.

The items selected for Form A and Form B are included in Table 4.15. As indicated in the table, selecting items based on their restricted point-biserial value, as opposed to their unrestricted value, resulted in Form B including 19 items that were not included in Form A.

The test variants for Examination 1, as well as for Examinations 2 and 3, therefore, included both similar and dissimilar items. Because each form included items that were also included on its corresponding test variant, dependent samples tests were used when evaluating differences in reliability and decision consistency. For each test conducted, however, a similar test using independent sample procedures was also performed. In each case, the result of the independent samples test was identical to the dependent samples test.

Descriptive statistics were calculated for the discrimination values associated with both forms. As seen in Table 4.16, the mean discrimination value for Form A, $M=0.39$ ( $S D$ $=0.08)$ was slightly larger than the mean value for Form $\mathrm{B}, M=.37(S D=0.09)$. To further investigate the difference in mean discrimination values, however, a dependent samples $t$-test was conducted.

Table 4.14
Descriptive Statistics of Test Forms $A$ and $B-$ All Examinations

| Exam | Form | $M$ | $S D$ | Min | Max | Range | Skew | $n$ | $\alpha$ | $S E M$ | $\kappa^{*}$ | $\kappa 95 \%$ CI |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

Note. * Numbers in parentheses indicate $S E$ for $\kappa$.

Table 4.15
Items and Discrimination Values for Form A and Form B - Examination 1

| Form A |  |  |  |  |  |  |  |  |  | Form B |  |  |  |
| ---: | ---: | ---: | :--- | ---: | :--- | ---: | :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Item | $r_{p b i s}$ | Item | $r_{p b i s}$ | Item | $r_{p b i s}$ | Item | $r_{p b i s}$ |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  | 6785 | 0.573 | 16725 | 0.369 |  |  |  |  |  |  |
| 6785 | 0.573 | 16375 | 0.370 | 8120 | 0.553 | 17240 | 0.368 |  |  |  |  |  |  |
| 16720 | 0.556 | 15600 | 0.369 | 16210 | 0.533 | 2830 | 0.362 |  |  |  |  |  |  |
| 8120 | 0.548 | 15635 | 0.366 | 16720 | 0.532 | 15810 | 0.350 |  |  |  |  |  |  |
| 16845 | 0.521 | 2830 | 0.364 | 16845 | 0.513 | $* 3570$ | 0.323 |  |  |  |  |  |  |
| 16210 | 0.510 | 3535 | 0.363 | 16710 | 0.492 | 16910 | 0.312 |  |  |  |  |  |  |
| 16710 | 0.489 | 16840 | 0.362 | 10645 | 0.473 | 2900 | 0.312 |  |  |  |  |  |  |
| 7240 | 0.480 | 16725 | 0.356 | 7240 | 0.452 | $* 16615$ | 0.309 |  |  |  |  |  |  |
| 10645 | 0.473 | 16760 | 0.355 | 14185 | 0.447 | $* 16940$ | 0.306 |  |  |  |  |  |  |
| 15450 | 0.461 | 12895 | 0.353 | 15450 | 0.443 | $* 16135$ | 0.306 |  |  |  |  |  |  |
| 14185 | 0.446 | 17240 | 0.351 | 17430 | 0.442 | $* 15355$ | 0.300 |  |  |  |  |  |  |
| 17430 | 0.442 | 7185 | 0.340 | 16005 | 0.438 | $* 15430$ | 0.295 |  |  |  |  |  |  |
| 15905 | 0.437 | 16910 | 0.337 | 10005 | 0.437 | $* 17595$ | 0.291 |  |  |  |  |  |  |
| 16740 | 0.436 | 15585 | 0.332 | 16740 | 0.429 | $* 15395$ | 0.288 |  |  |  |  |  |  |
| 16005 | 0.436 | 15810 | 0.329 | 10560 | 0.407 | $* 16060$ | 0.288 |  |  |  |  |  |  |
| 2545 | 0.432 | 15705 | 0.326 | 15890 | 0.407 | $* 6620$ | 0.281 |  |  |  |  |  |  |
| 2175 | 0.427 | 12960 | 0.317 | 2175 | 0.401 | $* 15755$ | 0.280 |  |  |  |  |  |  |
| 5925 | 0.425 | 17370 | 0.317 | 16110 | 0.395 | $* 16635$ | 0.273 |  |  |  |  |  |  |
| 10005 | 0.417 | 16690 | 0.312 | 2545 | 0.391 | $* 13695$ | 0.265 |  |  |  |  |  |  |
| 16765 | 0.415 | 2900 | 0.311 | 16765 | 0.388 | $* 3520$ | 0.260 |  |  |  |  |  |  |
| 15890 | 0.413 | 2385 | 0.300 | 15635 | 0.385 | $* 16745$ | 0.257 |  |  |  |  |  |  |
| 16110 | 0.410 | 17500 | 0.294 | 3535 | 0.384 | $* 17010$ | 0.254 |  |  |  |  |  |  |
| 6710 | 0.407 | 1870 | 0.292 | 2250 | 0.381 | $* 17385$ | 0.249 |  |  |  |  |  |  |
| 13245 | 0.392 | 17580 | 0.278 | 13245 | 0.380 | $* 12965$ | 0.249 |  |  |  |  |  |  |
| 10560 | 0.379 | 16445 | 0.277 | 15600 | 0.370 | $* 16305$ | 0.170 |  |  |  |  |  |  |
| 2250 | 0.374 | 14275 | 0.264 |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |

Note. * = Item exclusive to Form B.

## Table 4.16

Descriptive Statistics of Form A and Form B Discrimination Values - Examination 1

| Form | $M$ | $S D$ | Min | Max | Range | Skew |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Form A | 0.39 | 0.08 | 0.26 | 0.57 | 0.31 | 0.51 |
| Form B | 0.37 | 0.09 | 0.17 | 0.57 | 0.40 | 0.31 |

The data were first tested for compliance with the assumption of normality. Results of the Shapiro-Wilk normality test (R Core Team, 2012) revealed that the data were normally distributed, Form A: $\mathrm{W}=0.97, p=.14$, Form $\mathrm{B}: \mathrm{W}=0.97, p=.21$. Results of the $t$-test revealed that the differences between discrimination values for Form A and Form B were not significant, $t(49)=1.66, p=.10$.

In terms of examination reliability, the estimate for Form A, $\alpha=0.89$, was slightly higher than the estimate for Form $B, \alpha=0.87$. These estimates, as well as other descriptive statistics, are included in Table 4.14. The result of the test for significant differences among coefficient alphas was significant, $t(488)=5.93, p<.001$, and, consequently, the null hypothesis of equal reliability estimates was rejected. The difference in examination reliability, therefore, was significant.

The classification decision consistency coefficient, $\kappa$, was also calculated for Form A and Form B. Huynh's (1976) modification to $\kappa$, which allows for an estimate based on a single test administration, was used for this purpose. As discussed in Chapter 3, the classification decision coefficients were to be compared using a method proposed by Donner
et al. (2000). Upon further investigation, however, the method required the $\kappa$ coefficients to be based on multiple test administrations. This was not possible for the data used in this study. Additional studies and models were considered, but each required the classification decision consistency coefficients to be based on multiple test administrations (Barnhart \& Williamson, 2002; McKenzie et al., 1996; Williamson, Lipsitz, \& Manatunga, 2000) or to utilize independent samples (Fleiss, 1981; Lipsitz, Williamson, Klar, Ibrahim, \& Parzen, 2001). Whereas no identified model for testing significant differences in $\kappa$ coefficients fit the data and context used in this study ( $\kappa$ based on dependent samples from a single test administration), a test was not possible. The comparison of classification decision consistency coefficients, therefore, was limited to an analysis of their associated 95\% confidence intervals. These intervals are included in Table 4.14.

A degree of caution should be used when characterizing differences between statistics using confidence intervals. Whereas one may conclude that a significant difference at the $\alpha=.05$ level of significance exists when $95 \%$ confidence intervals do not overlap, it may be misleading to suggest that a significant difference does not exist when the confidence intervals do overlap. Previous research has shown that statistics with overlapping confidence intervals may, in fact, be significantly different (Odeuyungbo, Thabane, \& Markle-Reid, 2009). As seen in Table 4.14, the classification decision consistency estimate for Form A, к $=0.69(S E=0.01,95 \%$ CI $[0.67,0.71])$ was slightly higher than that of Form B, $\kappa=0.67(S E$ $=0.01,95 \% \mathrm{CI}[0.65,0.69]$. Had the $95 \%$ confidence intervals not overlapped, the null hypothesis of no significant difference between $\kappa$ coefficients could have been rejected. In this case, however, the confidence intervals did overlap. It is plausible, although not a
certainty, therefore, that there is no significant difference between the classification decision consistency coefficients associated with Form A and Form B.

## Examination 2

The procedures used to create Form A and Form B for Examination 1 were also conducted for Examination 2. As seen in Table 4.17, using restricted point-biserials resulted in Form B including 22 items that were not included in Form A.

Descriptive statistics for the discrimination values associated with the Examination 2 forms are included in Table 4.18. Once again, the mean item discrimination value was slightly larger for Form A, $M=0.37(S D=0.05)$, than it was for Form B, $M=0.35(S D=$ 0.08). A dependent samples $t$-test was conducted to test for differences among the group means. The Shapiro-Wilk normality test (R Core Team, 2012) indicated that the data approximated normality, $\mathrm{W}=0.98, p=.41$. Results of the $t$-test indicated that the difference between mean discrimination values was not significant, $t(49)=1.90, p=.06$.

Table 4.17
Items and Discrimination Values for Form A and Form B - Examination 2

| Item | Form A |  | $r_{p b i s}$ | Item | Form B |  | $r_{p b i s}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $r_{p b i s}$ | Item |  |  | $r_{p b i s}$ | Item |  |
| 108153 | 0.524 | 108155 | 0.363 | 108153 | 0.545 | 108162 | 0.338 |
| 108486 | 0.454 | 108507 | 0.363 | 108143 | 0.502 | *60242 | 0.337 |
| 90390 | 0.450 | 108146 | 0.361 | 108486 | 0.487 | 108494 | 0.336 |
| 108492 | 0.448 | 59231 | 0.360 | 40885 | 0.443 | *40944 | 0.333 |
| 108178 | 0.443 | 94586 | 0.358 | 90390 | 0.437 | *51781 | 0.331 |
| 108143 | 0.436 | 108162 | 0.357 | 108487 | 0.435 | *108478 | 0.328 |
| 108138 | 0.428 | 108189 | 0.356 | 108178 | 0.434 | *108479 | 0.326 |
| 40885 | 0.423 | 108494 | 0.354 | 51787 | 0.432 | *108144 | 0.323 |
| 51787 | 0.417 | 51793 | 0.348 | 62140 | 0.423 | *41645 | 0.321 |
| 62140 | 0.413 | 108487 | 0.346 | 108154 | 0.423 | *108198 | 0.320 |
| 40655 | 0.413 | 108182 | 0.346 | 40969 | 0.418 | *62132 | 0.319 |
| 94582 | 0.407 | 94566 | 0.338 | 40992 | 0.418 | 108189 | 0.318 |
| 40969 | 0.402 | 108156 | 0.338 | 41181 | 0.415 | 62136 | 0.316 |
| 108509 | 0.397 | 108515 | 0.330 | 108176 | 0.409 | *108193 | 0.307 |
| 108489 | 0.394 | 108166 | 0.327 | 94582 | 0.395 | *108197 | 0.297 |
| 108477 | 0.394 | 84442 | 0.326 | 40655 | 0.387 | *90398 | 0.265 |
| 108154 | 0.390 | 108201 | 0.324 | 84442 | 0.376 | 108515 | 0.264 |
| 40992 | 0.389 | 59149 | 0.319 | *108187 | 0.372 | *108481 | 0.255 |
| 108177 | 0.389 | 108506 | 0.317 | 94566 | 0.369 | *108512 | 0.252 |
| 41181 | 0.385 | 108180 | 0.303 | 108492 | 0.366 | *108488 | 0.242 |
| 51817 | 0.380 | 40668 | 0.303 | 108138 | 0.364 | *90400 | 0.240 |
| 40657 | 0.379 | 108152 | 0.298 | *59128 | 0.350 | *108158 | 0.228 |
| 108176 | 0.375 | 108483 | 0.295 | 60258 | 0.349 | *62135 | 0.225 |
| 108204 | 0.373 | 60258 | 0.292 | 108180 | 0.343 | *108502 | 0.170 |
| 62136 | 0.368 | 108167 | 0.284 | 108477 | 0.340 | *59138 | 0.157 |

Note. * = Item exclusive to Form B.

Table 4.18
Descriptive Statistics of Form A and Form B Discrimination Values - Examination 2

| Form | $M$ | $S D$ | Min | Max | Range | Skew |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Form A | 0.37 | 0.05 | 0.28 | 0.52 | 0.24 | 0.47 |
| Form B | 0.35 | 0.08 | 0.16 | 0.55 | 0.39 | -0.07 |

As indicated in Table 4.14, the examination reliability estimate for Form A, $\alpha=0.87$, was slightly larger than the estimate for Form B, $\alpha=0.85$. Further analysis also revealed that the difference between examination reliability estimates was significant, $t(159)=2.20, p=$ .03. Thus, the null hypothesis of equal coefficient alphas was rejected.

In terms of classification decision consistency, the $\kappa$ coefficient associated with Form $\mathrm{A}, \kappa=0.66(S E=0.02,95 \%$ CI $[0.62,0.70])$ was slightly greater than the estimate for Form B, $\kappa=0.63(S E=0.02,95 \%$ CI $[0.59,0.68])$. As was the case with Examination 1, the confidence intervals for the forms associated with Examination 2 overlapped. A similar conclusion, therefore, may be made: It is plausible that the observed difference in classification decision consistency between Form A and Form B is not statistically significant.

## Examination 3

For Examination 3, the use of restricted point-biserials resulted in Form B including 22 items that were not included in Form A. These items are annotated in Table 4.19. Descriptive statistics for the discrimination values for Form A and Form B are included in Table 4.20. The mean discrimination value for Form A, $M=0.33$ ( $S D=0.08$ ), was slightly higher than that of Form $\mathrm{B}, M=0.29(S D=0.13)$. Prior to conducting the $t$-test, the values were also tested for normality. The Shapiro-Wilk test (R Core Team, 2012) revealed that the data were normally distributed, Form $\mathrm{A}: \mathrm{W}=0.98, p=.74$, Form $\mathrm{B}: \mathrm{W}=0.97, p=.18$. Results of the test indicated that the difference between mean discrimination values was not significant, $t(49)=1.85, p=.07$.

Examination reliability estimates were also calculated for the Examination 3 forms. As indicated in Table 4.14, the reliability estimate for Form A, $\alpha=0.83$, was again larger than the estimate for Form $\mathrm{B}, \alpha=0.78$. Once again, further analysis revealed that the difference between examination reliability estimates was significant, $t(74)=2.10, p=.04$.

Finally, Form A and Form B were compared with regard to classification decision consistency. As seen in Table 4.14, the estimate for Form A, $\kappa=0.59(S E=0.04,95 \%$ CI [ $0.51,0.67])$ was larger than for Form $\mathrm{B}, \kappa=0.52(S E=0.05,95 \%$ CI $[0.43,0.62])$.

Table 4.19
Items and Discrimination Values for Form A and Form B - Examination 3

| Item | Form A |  | $r_{\text {pbis }}$ | Item | Form B |  | $r_{p b i s}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $r_{p b i s}$ | Item |  |  | $r_{p b i s}$ | Item |  |
| 160005 | 0.540 | 160048 | 0.313 | 160084 | 0.518 | 160050 | 0.310 |
| 160041 | 0.467 | 160549 | 0.311 | 160552 | 0.516 | *160021 | 0.309 |
| 160557 | 0.455 | 160055 | 0.311 | 160057 | 0.483 | 160048 | 0.307 |
| 160034 | 0.449 | 160056 | 0.309 | 160030 | 0.464 | 160064 | 0.265 |
| 160057 | 0.443 | 160548 | 0.308 | 160074 | 0.464 | 160053 | 0.261 |
| 160032 | 0.443 | 160011 | 0.306 | 160005 | 0.458 | *160013 | 0.260 |
| 160051 | 0.419 | 160558 | 0.306 | 159998 | 0.444 | 160028 | 0.240 |
| 160541 | 0.417 | 160559 | 0.305 | 160032 | 0.443 | *160047 | 0.228 |
| 160550 | 0.407 | 160573 | 0.298 | 160557 | 0.442 | *160026 | 0.214 |
| 160029 | 0.402 | 160569 | 0.295 | 160041 | 0.422 | *160020 | 0.208 |
| 160030 | 0.400 | 159999 | 0.293 | 160541 | 0.404 | *160578 | 0.207 |
| 159998 | 0.399 | 160526 | 0.290 | 160051 | 0.402 | *160536 | 0.200 |
| 160074 | 0.396 | 160535 | 0.285 | 160079 | 0.398 | *160577 | 0.197 |
| 160555 | 0.396 | 160018 | 0.276 | 160550 | 0.396 | *160025 | 0.190 |
| 160079 | 0.392 | 160050 | 0.263 | 160055 | 0.392 | *159996 | 0.176 |
| 159997 | 0.378 | 160028 | 0.261 | 160014 | 0.365 | *160527 | 0.167 |
| 160552 | 0.367 | 160035 | 0.253 | 160526 | 0.355 | *160072 | 0.156 |
| 160006 | 0.366 | 160570 | 0.247 | *160031 | 0.352 | *160039 | 0.140 |
| 160084 | 0.353 | 160572 | 0.245 | 160006 | 0.346 | *160554 | 0.129 |
| 160547 | 0.343 | 160064 | 0.242 | *160540 | 0.332 | *160533 | 0.125 |
| 160564 | 0.336 | 160059 | 0.233 | 160035 | 0.331 | *160000 | 0.110 |
| 160014 | 0.335 | 160046 | 0.218 | 160034 | 0.330 | *160575 | 0.077 |
| 160556 | 0.335 | 160566 | 0.196 | 160564 | 0.324 | *160090 | 0.074 |
| 160537 | 0.333 | 160017 | 0.178 | 160535 | 0.312 | *160545 | 0.057 |
| 160053 | 0.315 | 160089 | 0.175 | 159999 | 0.312 | *160568 | 0.053 |

Note. * = Item exclusive to Form B.

Table 4.20
Descriptive Statistics of Form A and Form B Discrimination Values - Examination 3

| Form | $M$ | $S D$ | Min | Max | Range | Skew |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Form A | 0.33 | 0.08 | 0.18 | 0.54 | 0.37 | 0.19 |
| Form B | 0.29 | 0.13 | 0.05 | 0.52 | 0.46 | -0.14 |

Once again, the $95 \%$ confidence intervals overlapped, suggesting that a conclusion similar to those drawn regarding the other examinations may be made: It is plausible that the differences in $\kappa$ between Form A and Form B are not statistically significant.

## Summary - Research Question 2

The procedures conducted to answer Research Question 2 suggested that using restricted point-biserials as the criterion for selection significantly affected the items selected. For Examination 1, 19 of the 50 items selected for Form B were not selected for Form A. For Examinations 2 and 3, 22 of the 50 items selected for Form B were not selected for Form A. Analysis of the item discrimination values indicated that the difference in mean pointbiserial between forms for each examination was not significant. The results also revealed, however, a significant difference in examination reliability. For each examination, the observed reliability estimate for Form A was slightly larger than that of Form B. In each case, further analysis found the difference in reliability between forms to be statistically significant. Finally, classification decision consistency was calculated for each form.

Observed estimates revealed that the $\kappa$ coefficient associated with Form A was larger than that of Form B for each examination. Analysis of the $95 \%$ confidence intervals, however, prevented a rejection of the null hypothesis of equal classification decision consistency estimates. It is plausible, therefore, that the observed difference between forms was not statistically significant. Further discussion and analysis of the results presented here are included in Chapter 5.

## CHAPTER 5

## DISCUSSION

Three areas of test development were emphasized in this study. First, the role item discrimination plays in the test development process was a key aspect of the research. Second, competency examinations used to credential individuals served as the context. Third, the study replicated the realistic conditions that many developers of credentialing examinations face by using relatively small samples of examinees, which, consequently, necessitated the use of classical test theory procedures. Previous research has been devoted to each of these areas individually. Prior to this study, however, no published research has examined the ways in which these areas interact. As such, the research represents a unique contribution to the field of test development.

No study, however, is without limitations. This chapter begins with an outline of limitations that affected the research. Next, a discussion of the study's key findings, specifically as they relate to the research questions, is provided. This is followed by recommendations for future research, practical implications for developers of examinations that incorporate cut scores to foster categorical decisions about examinees (e.g., credentialing examinations), and conclusions.

## Limitations

The study was affected by two general limitations. These limitations are presented below and are discussed in greater detail in the sections that follow.

1. The items used in the analysis were included in the final versions of their respective examinations. As such, they potentially may have been more vetted than the types of items typically under consideration during the item analysis phase of test development, which this study sought to replicate.
2. The criteria by which items were selected for the Form A and Form B test variants potentially did not wholly correspond to methods used by those who develop credentialing examinations.

## Refinement of Items Used in the Study

In many ways, this study attempted to replicate certain aspects of the process by which credentialing examinations are developed. Like other types of tests, the process used to develop credentialing examinations includes well-established steps. These steps were highlighted in Table 1.1, which was adapted from research conducted by Downing (2006). As seen in the table, item analysis typically occurs during the eighth step in this process, a step labeled "scoring test responses." During this stage of development, analysis of fieldtested items frequently occurs. One component of this analysis is the calculation and evaluation of item discrimination values.

The procedures conducted in this study, specifically as they relate to Research Question 2, sought to imitate item analysis. Item discrimination values were calculated and used, for example, to select items for the Form A and Form B test variants just as they might be used by test developers to select items for a credentialing examination. The items included in these forms were then used to answer Research Question 2. For the items used in
this study, however, the process of item analysis had already occurred. These items were included in the final version of each of their respective examinations, and, therefore, were presumably selected over other less qualified items. The process of item analysis also may have resulted in some form of modification to the items, such as changes in wording or order of response options. A more accurate replication of the item analysis phase of test development might include items that had not yet been selected for inclusion in the final version of an examination, but which, along with many other items, were under consideration for inclusion in the final version.

Although this limitation represents a slight deviation from the typical conditions under which credentialing examination are normally developed, it is unlikely that it affected the study's overall findings. The research conducted was focused on the degree to which using restricted point-biserials affected the discrimination values themselves, item selection, examination reliability, and classification decision consistency. In each instance, the results of the research with respect to the dependent variables just named would likely not have been different had items in earlier phases of development been used. Exploring the effect less refined items might have, however, may be valuable and is discussed in greater detail in the section outlining recommendations for future research.

## Item Selection Criteria

Another limitation associated with the current research is that in creating the forms used to answer Research Question 2, item discrimination served as the sole criterion for item selection. Although discrimination is typically an important consideration when selecting
items for examinations, other factors, which were not considered here, may also determine the degree to which an item is appropriate for inclusion.

One such consideration, for example, may be item difficulty, which, for items scored dichotomously, represents the proportion of examinees who have answered the item correctly (Crocker \& Algina, 2008). Item difficulty may be of particular importance for examinations for which examinees are classified into multiple categories, such as basic, proficient, and advanced. For these types of examinations, developers may want to include several items with difficulty levels that correspond to the various performance categories. Item discrimination, in such instances, may be a secondary consideration.

In most cases, however, item statistics, to include both discrimination and difficulty, are not the only factors that determine suitability for inclusion in final-version examinations. According to Livingston (2006), "Statistics alone cannot determine which items on a test are good and which are bad, but statistics can be used to identify items that are worth a particularly close look" (p. 423). Other factors that influence the selection of items are based on test specifications. Such specifications are sometimes referred to as test blueprints, because they specify how the test or form is to be constructed (Schmeiser \& Welch, 2006). A test blueprint may specify, for example, that a certain number of items should be associated with particular content standards. Using a third-grade mathematics examination as an example, a test blueprint may direct that $20 \%$ of the items should be devoted to each of the following: addition, subtraction, simple multiplication, geometry, and fractions. As such, regardless of the statistics associated with items related to geometry, $20 \%$ of the items must cover that content standard. Consequently, it is possible that a geometry-related item with a relatively low item discrimination value may be selected, whereas an addition-related
item with a higher discrimination value may not. Test blueprints may also stipulate other examination characteristics, such as the types of items to be used (i.e., constructed-response versus selected-response formats), the ordering of items (i.e., based on difficulty or content domain), test length, item scoring, and delivery specifications.

An important aspect of this study was the creation of the Form A and Form B test variants used to answer Research Question 2. Because test blueprints were not available for the examinations used in this research, item discrimination was the only criterion considered when selecting items for inclusion in the pseudo-forms. Form A included the 50 most discriminating items using unrestricted discrimination values as the selection criterion. Form B included the 50 most discriminating items using restricted discrimination values as the selection criterion. Had test blueprints been available, those specifications could have also been considered. Examination 3, for example, measured nurses' understanding of diabetes. It is possible that the test blueprint required a certain number of items to cover risk factors for diabetes, others to cover treatment of diabetes, and yet others to cover differences between types of diabetes. Because test specifications were not available, considering these types of issues in the item selection process was not possible.

It is important that the items selected for an examination match the requirements outlined in the test specifications because the items represent portions of the content standards. If the items do not sufficiently cover the content standards as specified by a test blueprint, a case for the validity of the examination is more difficult to make. According to Schmeiser and Welch (2006), "The domains to which test-score inferences are to be made serve as examples of the sources of validity evidence that can be used" (p.315).

Although the process used to select items for the Form A and Form B test variants did not consider these additional factors, the results are still important with respect to the relationship between restricted item discrimination values and those aspects of test development examined in this study, namely, item selection, examination reliability, and classification decision consistency. Item discrimination may not be the only factor considered when selecting items for an examination, but it almost always is a factor of consideration. Despite this limitation, therefore, the research findings presented here are still important and relevant.

## Key Findings

The results of the study produced four key findings: one related to Research Question 1 and three related to Research Question 2. The key findings are summarized below. Each key finding is discussed in greater detail in the sections that follow.

1. Restricting the calculation of item discrimination values to scores of examinees at or near anticipated cut scores resulted in lower discrimination values than those calculated using all examinee scores. (Research Question 1)
2. Using restricted item discrimination values as the primary criterion for item selection resulted in the selection of items that were different than those selected using unrestricted item discrimination values. (Research Question 2)
3. Examinations comprised of items selected using restricted item discrimination values as the primary selection criterion produced slightly lower reliability estimates than examinations comprised of items selected using unrestricted item
discrimination values. Differences in reliability estimates between these examinations were significantly greater than zero. (Research Question 2)
4. Examinations comprised of items selected using restricted item discrimination values as the primary selection criterion produced slightly lower observed classification decision consistency estimates than examinations comprised of items selected using unrestricted item discrimination values. The degree to which these differences were statistically significant, however, was uncertain. (Research Question 2)

## Effect on Item Discrimination Values

The purpose of Research Question 1 was to determine the effect limiting discrimination values to scores of examinees near real or anticipated cut scores had on item discrimination values. To answer this question, procedures were followed that resulted in the creation of 15 sets of restricted item discrimination values. The sets of restricted values were based on scores within 1.00 SD, 0.75SD, and 0.50 SD of five distinct cut score locations. As observed in Table 5.1, in all cases, mean restricted item discrimination values were smaller than corresponding mean unrestricted item discrimination values. For Examinations 1 and 2, the results of one-way ANOVA analysis indicated that the difference between the unrestricted set of discrimination values and each of the 15 sets of restricted discrimination values was significantly greater than zero. This was also the case for all but four of the sets of restricted discrimination values in Examination 3.

Table 5.1
Descriptive Statistics of Discrimination Values - All Examinations

| Cut Score | Group | Examination 1 |  | Examination 2 |  | Examination 3 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $M(S D)$ | $n$ | $M(S D)$ | $n$ | $M(S D)$ | $n$ |
| Unrestricted |  | 0.25 (.09) | 490 | 0.22 (.12) | 161 | 0.16 (.13) | 76 |
| $\mathrm{C}_{\mathrm{X} 1}$ | +/-1.00 SD | *0.13 (.08) | 339 | *0.12 (.12) | 111 | 0.08 (.17) | 37 |
|  | +/- 0.75SD | *0.10 (.07) | 287 | *0.09 (.12) | 85 | *0.06 (.19) | 25 |
|  | +/- 0.50 SD | *0.07 (.08) | 198 | *0.06 (.13) | 67 | *0.04 (.23) | 17 |
| $\mathrm{C}_{\mathrm{X} 2}$ | +/-1.00 SD | *0.12 (.07) | 317 | *0.12 (.11) | 115 | 0.08 (.15) | 43 |
|  | +/- 0.75SD | *0.10 (.07) | 267 | *0.10 (.12) | 96 | *0.06 (.17) | 36 |
|  | +/- 0.50 SD | *0.07 (.08) | 200 | *0.06 (.13) | 64 | *0.05 (.21) | 25 |
| $\mathrm{C}_{\mathrm{X} 3}$ | +/-1.00 SD | *0.14 (.08) | 344 | *0.11 (.12) | 93 | 0.08 (.19) | 23 |
|  | +/- 0.75SD | *0.10 (.08) | 255 | *0.08 (.13) | 76 | *0.06 (.20) | 20 |
|  | +/-0.50 SD | *0.07 (.09) | 171 | *0.06 (.14) | 57 | *0.04 (.29) | 11 |
| $\mathrm{C}_{\mathrm{X} 4}$ | +/-1.00 SD | *0.13 (.07) | 292 | *0.12 (.11) | 109 | 0.09 (.13) | 52 |
|  | +/- 0.75SD | *0.10 (.08) | 240 | *0.09 (.11) | 89 | *0.06 (.16) | 33 |
|  | +/-0.50 SD | *0.07 (.09) | 169 | *0.07 (.13) | 57 | *0.05 (.19) | 26 |
| $\mathrm{C}_{\mathrm{X} 5}$ | +/-1.00 SD | *0.14 (.08) | 307 | *0.10 (.12) | 82 | *0.06 (.28) | 13 |
|  | +/- 0.75SD | *0.11 (.08) | 239 | *0.09 (.13) | 60 | *0.05 (.34) | 8 |
|  | +/- 0.50 SD | *0.07 (.10) | 152 | *0.05 (.18) | 37 | *0.04 (.42) | 6 |

Note. * = Analysis indicated a significant difference between these sets of restricted values and their corresponding set of unrestricted values at the $\alpha=.05$ level of significance.

One of the reasons the study incorporated three examinations, five distinct cut score locations, and three different bands of examinee scores around those cut score locations was to evaluate how each of those variables affected the results. For Examination 1, which was the largest sample used in study $(n=490)$, and Examination 2, which was the second largest sample used $(n=161)$, the location of the cut score appeared to have little influence on the outcome of the results. Regardless of either the location of the cut score or the size of the group around that cut score location considered, the difference between the resulting sets of restricted discrimination values and the unrestricted set of discrimination values was found to be significantly greater than zero. In all cases, the mean restricted discrimination values were smaller than the mean unrestricted discrimination value.

Examination 3, which was the smallest sample used in the study ( $n=76$ ), produced results similar to those identified for Examinations 1 and 2, with four exceptions. Although the observed mean discrimination value for each of the sets of restricted values was smaller than the unrestricted set, the difference between the unrestricted set and four of the restricted sets was found to be not significantly greater than zero. Those four sets included values based on scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 1}, \mathrm{C}_{\mathrm{X} 2}, \mathrm{C}_{\mathrm{X} 3}$, and $\mathrm{C}_{\mathrm{X} 4}$.

After completing the initial set of procedures used to answer Research Question 1, it was determined that additional two-way ANOVA would be useful in determining the effect cut score location and examinee score group size, as well as any possible interaction between these two factors, had on restricted item discrimination values. The analysis was developed using five levels for the cut score location factor ( $\mathrm{C}_{\mathrm{X} 1}, \mathrm{C}_{\mathrm{X} 2}, \mathrm{C}_{\mathrm{X} 3}, \mathrm{C}_{\mathrm{X} 4}$, and $\left.\mathrm{C}_{\mathrm{X} 5}\right)$ and three levels for the examinee score group size factor (1.00 SD, 0.75SD, and $0.50 S D)$. The procedure, like the one-way ANOVA conducted earlier, incorporated repeated measures.

Violations of the sphericity assumption, which affected each of the examinations, were resolved using the Greenhouse-Geisser (1959) correction.

For each of the examinations used, there was no significant main effect for the cut score location factor or for the interaction of cut score location and examinee score group size. There was, however, a significant main effect for the factor representing the size of the group of scores considered for each examination, Examination 1: $F(1.20,208.80)=492.35, p$ $<.001$; Examination 2: $F(1.22,243.97)=161.84, p<.001$; Examination 3: $F(1.75,259.59)=$ $14.62, p<.001$. The analysis confirmed that cut score location did not significantly affect the magnitude of the restricted item discrimination values. The examinee score group size considered when calculating the discrimination values, however, did significantly affect the restricted point-biserials. For each examination, as the size of the group of scores considered decreased in size, so too did its associated mean item discrimination value.

After the procedures described in Chapter 3 were conducted, an evaluation of the differences between unrestricted discrimination values and restricted values was conducted at the item level. That is, for each item in each examination, differences between the unrestricted discrimination value and the 15 sets of restricted values were examined. The results of this analysis are included in Table 5.2. The table indicates the number of item discrimination values for each examination that either increased or decreased when the restricted conditions were considered.

Table 5.2
Change in Direction of Discrimination Values - All Examinations

| Cut Score | Group | Examination 1 |  | Examination 2 |  | Examination 3 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | - $\Delta$ (\%) | + $\Delta(\%)$ | - $\Delta$ (\%) | $+\Delta(\%)$ | - $\Delta$ (\%) | + $\Delta(\%)$ |
| $\mathrm{C}_{\mathrm{X} 1}$ | +/-1.00 SD | $\begin{gathered} 169 \\ (96.6 \%) \end{gathered}$ | $\begin{gathered} 6 \\ (3.4 \%) \end{gathered}$ | $\begin{gathered} 165 \\ (82.5 \%) \end{gathered}$ | $\begin{gathered} 35 \\ (17.5 \%) \end{gathered}$ | $\begin{gathered} 106 \\ (71.1 \%) \end{gathered}$ | $\begin{gathered} 43 \\ (28.9 \%) \end{gathered}$ |
|  | +/- 0.75SD | $\begin{gathered} 172 \\ (98.3 \%) \end{gathered}$ | $\begin{gathered} 3 \\ (1.7 \%) \end{gathered}$ | $\begin{gathered} 167 \\ (83.5 \%) \end{gathered}$ | $\begin{gathered} 33 \\ (16.5 \%) \end{gathered}$ | $\begin{gathered} 106 \\ (71.1 \%) \end{gathered}$ | $\begin{gathered} 43 \\ (28.9 \%) \end{gathered}$ |
|  | +/- 0.50 SD | $\begin{gathered} 168 \\ (96.0 \%) \end{gathered}$ | $\begin{gathered} 7 \\ (4.0 \%) \end{gathered}$ | $\begin{gathered} 168 \\ (84.0 \%) \end{gathered}$ | $\begin{gathered} 32 \\ (16.0 \%) \end{gathered}$ | $\begin{gathered} 102 \\ (68.5 \%) \end{gathered}$ | $\begin{gathered} 47 \\ (31.5 \%) \end{gathered}$ |
| $\mathrm{C}_{\mathrm{X} 2}$ | +/-1.00 SD | $\begin{gathered} 170 \\ (97.1 \%) \end{gathered}$ | $\begin{gathered} 5 \\ (2.9 \%) \end{gathered}$ | $\begin{gathered} 175 \\ (87.5 \%) \end{gathered}$ | $\begin{gathered} 25 \\ (12.5 \%) \end{gathered}$ | $\begin{gathered} 109 \\ (73.2 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (26.8 \%) \end{gathered}$ |
|  | +/- 0.75SD | $\begin{gathered} 174 \\ (99.4 \%) \end{gathered}$ | $\begin{gathered} 1 \\ (0.6 \%) \end{gathered}$ | $\begin{gathered} 175 \\ (87.5 \%) \end{gathered}$ | $\begin{gathered} 25 \\ (12.5) \end{gathered}$ | $\begin{gathered} 109 \\ (73.2 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (26.8 \%) \end{gathered}$ |
|  | +/- 0.50 SD | $\begin{gathered} 169 \\ (96.6 \%) \end{gathered}$ | $\begin{gathered} 6 \\ (3.4 \%) \end{gathered}$ | $\begin{gathered} 172 \\ (86.0 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (14.0 \%) \end{gathered}$ | $\begin{gathered} 102 \\ (68.5 \%) \end{gathered}$ | $\begin{gathered} 47 \\ (31.5 \%) \end{gathered}$ |
| $\mathrm{C}_{\mathrm{X} 3}$ | +/-1.00 SD | $\begin{gathered} 170 \\ (97.1 \%) \end{gathered}$ | $\begin{gathered} 5 \\ (2.9 \%) \end{gathered}$ | $\begin{gathered} 160 \\ (80.0 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (20.0 \%) \end{gathered}$ | $\begin{gathered} 110 \\ (73.8 \%) \end{gathered}$ | $\begin{gathered} 39 \\ (26.2 \%) \end{gathered}$ |
|  | +/- 0.75SD | $\begin{gathered} 171 \\ (97.7 \%) \end{gathered}$ | $\begin{gathered} 4 \\ (2.3 \%) \end{gathered}$ | $\begin{gathered} 160 \\ (80.0 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (20.0 \%) \end{gathered}$ | $\begin{gathered} 108 \\ (72.5 \%) \end{gathered}$ | $\begin{gathered} 41 \\ (27.5 \%) \end{gathered}$ |
|  | +/- 0.50 SD | $\begin{gathered} 172 \\ (98.3 \%) \end{gathered}$ | $\begin{gathered} 3 \\ (1.7 \%) \end{gathered}$ | $\begin{gathered} 166 \\ (83.0 \%) \end{gathered}$ | $\begin{gathered} 34 \\ (17.0 \%) \end{gathered}$ | $\begin{gathered} 101 \\ (67.8 \%) \end{gathered}$ | $\begin{gathered} 48 \\ (32.3 \%) \end{gathered}$ |
| $\mathrm{C}_{\mathrm{X} 4}$ | +/-1.00 SD | $\begin{gathered} 162 \\ (92.6 \%) \end{gathered}$ | $\begin{gathered} 13 \\ (7.4 \%) \end{gathered}$ | $\begin{gathered} 164 \\ (82.0 \%) \end{gathered}$ | $\begin{gathered} 36 \\ (18.0 \%) \end{gathered}$ | $\begin{gathered} 107 \\ (71.8 \%) \end{gathered}$ | $\begin{gathered} 42 \\ (28.2 \%) \end{gathered}$ |
|  | +/- 0.75SD | $\begin{gathered} 169 \\ (96.6 \%) \end{gathered}$ | $\begin{gathered} 6 \\ (3.4 \%) \end{gathered}$ | $\begin{gathered} 172 \\ (86.0 \%) \end{gathered}$ | $\begin{gathered} 28 \\ (14.0 \%) \end{gathered}$ | $\begin{gathered} 110 \\ (73.8 \%) \end{gathered}$ | $\begin{gathered} 39 \\ (26.2 \%) \end{gathered}$ |
|  | +/- 0.50 SD | $\begin{gathered} 169 \\ (96.6 \%) \end{gathered}$ | $\begin{gathered} 6 \\ (3.4 \%) \end{gathered}$ | $\begin{gathered} 167 \\ (83.5 \%) \end{gathered}$ | $\begin{gathered} 33 \\ (16.5 \%) \end{gathered}$ | $\begin{gathered} 112 \\ (75.2 \%) \end{gathered}$ | $\begin{gathered} 37 \\ (24.8 \%) \end{gathered}$ |
| $\mathrm{C}_{\mathrm{X} 5}$ | +/-1.00 SD | $\begin{gathered} 171 \\ (97.7 \%) \end{gathered}$ | $\begin{gathered} 4 \\ (2.3 \%) \end{gathered}$ | $\begin{gathered} 163 \\ (81.5 \%) \end{gathered}$ | $\begin{gathered} 37 \\ (18.5 \%) \end{gathered}$ | $\begin{gathered} 106 \\ (71.1 \%) \end{gathered}$ | $\begin{gathered} 43 \\ (28.9 \%) \end{gathered}$ |
|  | +/- 0.75SD | $\begin{gathered} 167 \\ (95.4 \%) \end{gathered}$ | $\begin{gathered} 8 \\ (4.6 \%) \end{gathered}$ | $\begin{gathered} 160 \\ (80.0 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (20.0 \%) \end{gathered}$ | $\begin{gathered} 102 \\ (68.5 \%) \end{gathered}$ | $\begin{gathered} 47 \\ (31.5 \%) \end{gathered}$ |
|  | +/- 0.50 SD | $\begin{gathered} 166 \\ (94.9 \%) \end{gathered}$ | $\begin{gathered} 9 \\ (5.1 \%) \end{gathered}$ | $\begin{gathered} 160 \\ (80.0 \%) \end{gathered}$ | $\begin{gathered} 40 \\ (20.0 \%) \end{gathered}$ | $\begin{gathered} 93 \\ (62.4 \%) \end{gathered}$ | $\begin{gathered} 56 \\ (37.6 \%) \end{gathered}$ |

As seen in the table, the discrimination value associated with the vast majority of examination items decreased when a subset of examinee scores was used to calculate the point-biserials. This was particularly true for Examination 1, where the discrimination values of 174 of the 175 items, or $99.4 \%$, decreased in size when their calculation was limited to examinee scores within $0.75 S D$ of $\mathrm{C}_{\mathrm{X} 2}$. This was the largest percentage of change in either direction for any group of restricted values across all examinations. For Examination 2, as many as $87.5 \%$ of unrestricted point-biserials decreased in value when restricted conditions were applied. The largest group of change in Examination 3 was associated with values calculated using scores within $0.50 S D$ of $\mathrm{C}_{\mathrm{X} 4}$. For this group, $75.2 \%$ of values decreased.

The relationship between the percentages of item discrimination values that decreased under restricted conditions and the examination sample size may also be observed. The most dramatic changes appear in the restricted sets associated with Examination 1, which included the largest sample size. Those changes were less dramatic in Examinations 2 and 3, for which the sample sizes were much smaller.

The observed decrease in discrimination values under restricted conditions was likely due in part to the fact that the samples upon which restricted values were calculated were more homogeneous than the samples upon which unrestricted values were calculated. As observed earlier, a key component in the calculation of the point-biserial statistic is the standard deviation of scores considered, which serves as the denominator in the formula:

$$
\begin{equation*}
\rho_{p b i s}=\frac{\mu+-\mu_{x}}{\sigma_{x}} \sqrt{p / q} \tag{5.1}
\end{equation*}
$$

where $\mu_{+}$is the mean total score for those who respond to the item correctly; $\mu_{x}$ is the mean total score for the entire group of examinees; $\sigma_{x}$ is the standard deviation for the entire group of examinees; $p$ is item difficulty; and $q$ is equal to $(1-p)($ Crocker \& Algina, 2008).

All else being equal, therefore, smaller standard deviations for the group of scores used in its calculation will result in smaller point-biserial statistics. In some cases, the smaller, more homogeneous groups of scores, along with their smaller standard deviations, resulted in positive point-biserials becoming negative. The standard deviations of the groups of scores used to calculate the unrestricted discrimination values and the restricted values based on scores within $1.00 S D$ of $\mathrm{C}_{\mathrm{X} 1}$ for each examination are included in Table 5.3. As seen in the table, the standard deviation of scores used to calculate restricted discrimination values are smaller than their corresponding unrestricted values. Although not included in Table 5.3, the standard deviations for each group of restricted values becomes smaller in size as the group used to calculate the values decreases in size. The standard deviations of groups of scores within 1.00 SD of each cut score were closer to the standard deviation of all scores, upon which the unrestricted values were calculated than were any of the other sets of restricted values. This may explain why the only four sets of restricted values that were not found to be significantly different than their corresponding unrestricted discrimination values, as seen in Table 5.1, were each based on scores within $1.00 S D$ of their associated cut score. The standard deviations of these groups of scores were closer to the standard deviation of the unrestricted set than were any other sets of restricted values in Examination 3.

Table 5.3
Descriptive Statistics of Selected Groups of Scores

| Exam | Cut score | Group | $n$ | $M$ | $S D$ | Min | Max | Range |
| :--- | :--- | :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Unrestricted |  | 490 | 111.42 | 19.55 | 46 | 161 | 115 |
|  | C $_{\mathrm{X} 1}$ | $+/-1.00 S D$ | 339 | 115.23 | 10.06 | 96 | 134 | 38 |
| 2 | Unrestricted |  | 161 | 134.27 | 18.30 | 82 | 174 | 92 |
|  | C $_{\mathrm{X} 1}$ | $+/-1.00 S D$ | 111 | 131.43 | 9.92 | 112 | 148 | 36 |
| 3 | Unrestricted |  | 76 | 115.04 | 9.01 | 93 | 134 | 41 |
|  | C $_{\mathrm{X} 1}$ | $+/-1.00 S D$ | 37 | 108.57 | 4.69 | 98 | 115 | 17 |

## Effect on Item Selection

Using restricted item discrimination values as the criterion by which items were selected resulted in the Form B test variant of each examination including numerous items that were not included in the Form A variant, which utilized unrestricted values as the selection criterion. The Form B variant of Examination 1, for instance, included 19 items that were not included in the Form A variant. The Form B variant of Examinations 2 and 3 each contained 22 items that were not included in the Form A variants. These values, as well as other descriptive statistics for the test variants used to answer Research Question 2, are included in Table 5.4. They are also represented in Figure 5.1.

Table 5.4
Descriptive Statistics of Forms A and B Test Variants - All Examinations


Note. * Numbers in parentheses indicate $S E$ for $\kappa$.


Figure 5.1. Pie charts representing breakdown of Form B test variant items for each examination.

This finding (i.e., that using restricted item discrimination values when evaluating items for inclusion in the test variants resulted in the selection of items that were different than selected using unrestricted discrimination values) is important because, in many ways, it speaks to the validity of such examinations. The items unique to the Form B variants were more discriminating for those examinees with total scores closest to the cut score, than were the items unique to Form A. Again, according to the Standards (AERA et al., 1999), test validity is "the degree to which evidence and theory support the interpretation of test scores entailed by proposed uses" (p. 9). The interpretation of scores associated with credentialing examinations is, of course, that those who pass are qualified to receive certification or licensure. Selecting items that better discriminate among examinees with scores close to the test cut score helps to support this interpretation.

This position is also supported by the previously mentioned work of Harris and Subkoviak (1986), who proposed item selection criteria for mastery tests, of which credentialing examinations are a form. Their words seem particularly relevant within the context of the current discussion:

For a mastery test, this means selecting items that discriminate between masters and non-masters, as opposed to within masters and within non-masters. The consensus appears to be that a good mastery item is one which masters answer correctly and non-masters answer incorrectly. (p. 496)

Likewise, The Standards (AERA et al., 1999) make similar recommendations regarding the importance of emphasizing examinees with total scores near cut scores:

Tests for credentialing need to be precise in the vicinity of the passing, or cut, score. They may not need to be precise for those who clearly pass or clearly fail. Sometimes a test used in credentialing is designed to be precise only in the vicinity of the cut score. (p. 157).

Although the use of restricted discrimination values may result in the selection of items that are less discriminating for examinees who clearly pass or fail the examination, their use in the item analysis phase of test development, as documented in this research, resulted in the selection of items that better discriminate among examinees with scores in the vicinity of the cut score. As such, their use supports the suggestions and recommendations outlined in the preceding quotations.

It is also important to understand how the inclusion of items based on restricted discrimination values affects the consistency with which candidates pass or fail the examination. To gain a better understanding of this effect, a series of tables highlighting the degree to which item selection affected pass/fail consistency were created. Figures 5.2, 5.3, and 5.4 display pass/fail consistency comparisons between the full test and both test variants for each of the examinations used in the study. In each of the figures, the cut score used for
the full examination was the actual cut score $\left(\mathrm{C}_{\mathrm{X} 1}\right)$ and a proportionally comparable cut score for the test variants.

Form A
(Examination 1)

|  |  | 212 |
| :---: | :---: | :---: |
| Pass | 47 |  |
| Fail | 22 | 209 |

Form B
(Examination 1)


Form B
(Examination 1)
Form A (Examination 1)
Pass Fail

| Pass | 223 | 9 |
| :---: | :---: | :---: |
| Fail | 36 | 222 |

Figure 5.2. Pass/fail consistency tables for Examination 1 and associated test variants.
Examination 2 (Full) Pass Fail
Form A
(Examination 2)

| Pass | 91 | 7 |
| :---: | :---: | :---: |
| Fail | 11 | 52 |

Form B
(Examination 2)

| Pass | 97 | 5 |
| :---: | :---: | :---: |
| Fail | 5 | 54 |

Form B (Examination 2)


Figure 5.3. Pass/fail consistency tables for Examination 2 and associated test variants.
Examination 3 (Full)
Pass Fail
Form A
(Examination 3)

|  |  |
| :---: | :---: |
| Pass | 54 |
| Fail | 10 |

Form B
(Examination 3)

| Pass | 59 | 2 |
| :---: | :---: | :---: |
| Fail | 5 | 10 |

Form B
(Examination 3)


Figure 5.4. Pass/fail consistency tables for Examination 3 and associated test variants.

The effect item selection method and, consequently, examination composition had on pass/fail rates may be observed in the figures. As seen in Figure 5.2, among those who passed the full version of Examination 1, 22 examinees would have failed the Form A variant and 31 would have failed the Form B variant.

The figure also illustrates the difference in pass/fail rates between the two test variants. Among those who passed the Examination 1 Form A variant, 36 examinees would have failed the Form B variant, which used restricted discrimination values to select items. Among the candidates who passed the Form B variant, nine would have failed the Form A variant, which used unrestricted discrimination values as the selection criterion. As displayed in Figures 5.3 and 5.4, this was not the case for Examinations 2 and 3. For each of those examinations, there were more examinees who passed the Form B variant (based on restricted discrimination values) but failed the Form A variant (based on unrestricted discrimination values) than examinees who passed the Form A but failed the Form B.

The method by which items were selected for inclusion in the test variants, therefore, played an important and consequential role in determining who passed and who failed the tests. For two of the examinations used in this study, the number of examinees who failed the variant using unrestricted discrimination values but passed the variant using restricted values was greater than those who passed the variant using unrestricted values but failed the variant using restricted values.

## Effect on Examination Reliability

For each of the examinations used in this study, scores associated with the Form B test variant, which utilized restricted item discrimination values as the item selection criterion, produced lower estimates of reliability than did the Form A variants, which used unrestricted discrimination values to select items. The reliability estimates, expressed in terms of coefficient alpha (Cronbach, 1951), are included in Table 5.4. Further analysis
found the differences in reliability estimates between the test variants for each examination to be significantly greater than zero.

A primary consideration when explaining the lower reliability estimates associated with the Form B variants may be their inclusion of unique items that were not included in the Form A variants. As mentioned earlier in this chapter, the Examination 1 Form B variant included 19 unique items. The Form B variant for Examinations 2 and 3 each included 22 unique items. The presence of these unique items led to lower total score variances for the Form B test variants when compared to the Form A versions. As seen in Table 5.4, the standard deviation of each Form B variant is lower than its corresponding Form A. Like the point-biserial statistic, total test score variance is a factor in the calculations of coefficient alpha:

$$
\begin{equation*}
\hat{\alpha}-\frac{k}{k-1}\left(1-\frac{\sum \hat{\sigma}_{i}^{2}}{\hat{\sigma}_{x}^{2}}\right) \tag{5.2}
\end{equation*}
$$

where $k$ is the number of items on the examination;
$\hat{\sigma}_{i}^{2}$ is the variance of item $i$; and
$\hat{\sigma}_{x}^{2}$ is the total test variance (Crocker \& Algina, 2008).

Examinations with lower total test variance, consequently, generally produce lower estimates of reliability.

A comparison of the unique items included in the Form B variants with the items that would have been selected had unrestricted discrimination values been used to select items, as was done with the Form A variants, is helpful in illustrating this point. These comparisons
are included in Table 5.5. As seen in the table, the items unique to Form B variants in each case led to scores that produced lower standard deviations and lower estimates of reliability when compared to the unique items included in their Form A counterparts. With the exception of these items, the remaining items among the examination-specific test variants were identical. The inclusion of the unique Form B items resulted in lower test score variance, and, therefore, lower reliability estimates.

Table 5.5
Descriptive Statistics for Unique Form A and Form B Test Variant Items - All Examinations

| Exam | Form | No. Items | $n$ | $M$ | $S D$ | Min | Max | $\alpha$ | SEM |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | Form A | 19 | 490 | 13.69 | 3.09 | 1 | 19 | 0.68 | 1.74 |
|  | Form B | 19 | 490 | 12.16 | 2.93 | 3 | 18 | 0.57 | 1.93 |
| 2 | Form A | 22 | 161 | 14.06 | 3.75 | 2 | 21 | 0.73 | 1.96 |
|  | Form B | 22 | 161 | 15.05 | 3.11 | 7 | 21 | 0.60 | 1.96 |
| 3 |  |  |  |  |  |  |  |  |  |
|  | Form A | 22 | 76 | 16.79 | 2.94 | 5 | 22 | 0.66 | 1.72 |
|  | Form B | 22 | 76 | 17.58 | 2.02 | 13 | 21 | 0.31 | 1.67 |

The inclusion of items that better discriminate among examinees with total scores nearest the examination cut score, therefore, appears to come at a cost. Whereas the use of restricted discrimination values allowed items to be included that emphasized the region of
the cut score, it resulted in lower levels of test score variance and lower examination reliability estimates.

A more important question in terms of this finding may be the degree to which developers of credentialing examinations are willing to accept the lower estimate of reliability that appear to be associated with the use of restricted discrimination values. Reliability, not unlike validity, should be interpreted within the framework of the examination's purpose. According to Haertel (2006), "test score reliability must be conceived relative to particular testing purposes and contexts" (p. 65). With the purpose of credentialing examinations being placed squarely on the qualification of examinees to receive certification or licensure, it may be possible that the validity-based benefits of using restricted discrimination values outweigh the lower estimates of reliability.

Another important consideration regarding this balance may be the degree to which the use of unrestricted discrimination values results in higher examination reliability estimates. Although the differences in reliability estimates between the test variants were found to be significantly greater than zero, the observed difference between variants was relatively minor. The observed difference in the Examination 1 and 2 variants, for example, was 0.02 . The difference between Examination 3 variants was 0.05 . Although there is no official threshold for the acceptability of examination reliability estimates, it is seems unlikely that such minor differences would be unacceptable to many test developers.

## Effect on Classification Decision Consistency

The Form B test variants, for each examination used, produced lower observed estimates of classification decision consistency, as expressed by coefficient $\kappa$, than did the

Form A variants. These observed values are included in Table 5.4. The lack of a test of significant differences between coefficients $\kappa$ based on dependent samples derived from a single test administration precluded the possibility of hypothesis testing. Analysis of differences, therefore, was limited to comparisons of 95\% confidence intervals. Whereas 95\% confidence intervals that do not overlap indicate differences that are significantly greater than zero at the $\alpha=.05$ level of significance, similar conclusions cannot be reached if the confidence intervals do overlap. For each of the examinations used, the coefficient $\kappa$ 95\% confidence intervals overlapped. It is plausible, therefore, that the observed differences are not significantly greater than zero. Statistically speaking, however, such a conclusion cannot be made with certainty.

Within the context of using restricted item discrimination values as a criterion for item selection, the interpretation of classification decision consistency coefficients is similar to that of the examination reliability estimates. Although the observed coefficients associated with the Form B test variants were lower than their Form A counterparts, the differences were relatively minor. The observed difference between Examination 1 test variants, for example, was 0.02 . The differences between Examinations 2 and 3 variants were 0.03 and 0.07 respectively. Coefficient $\kappa$ is interpreted as the increase in decision consistency over chance as a proportion of the maximum possible increase over chance (Crocker \& Algina, 2008). For the Form A variant associated with Examination 1, therefore, $69 \%$ of the total possible increase over chance consistency was observed. This figure was $67 \%$ for the Form B variant.

The primary question, therefore, may be the acceptability of slightly lower classification decision consistency coefficients when restricted item discrimination values are
used. Again, the extent to which the differences between classification decision consistency coefficients produced using unrestricted and restricted discrimination values are greater than zero is not known. However, it seems unlikely that the relatively minor observed differences would lead to the elimination of restricted values as a consideration when conducting item analysis.

## Recommendations for Future Research

While conducting the research, several recommendations for future research were identified. These recommendations include:

1. Conduct research with items that have not previously gone through the process of item analysis.
2. Conduct research using test specifications, or blueprints, to guide the selection of items.
3. Develop longer test variants to assess the difference between tests with items selected using restricted discrimination values and tests with items selected using unrestricted values.
4. Examine the degree to which non-classical test theory approaches, such as item response theory, support the findings of the current study.
5. Study how using restricted discrimination values affects the standard setting process.

Each recommendation is discussed in further detail in the sections that follow.

## Conduct Research with Less Refined Items

The recommendation to conduct similar research with less refined items is based on a limitation discussed earlier. Each of the three examinations used in this study included items that had previously undergone the process of item analysis. As such, the items used had, presumably, already met certain thresholds that validated their inclusion in the examination. It is also possible that based on this analysis, the items maybe have been modified in certain aspects, such as item or selected response wording.

Future research might examine the degree to which using less refined items, such as those in actual banks of field-tested items, supports the results of the current study. Using less refined items would more accurately replicate the process of item analysis, which, in many ways, was the intent of the procedures conducted to answer Research Question 2.

## Use Test Specifications to Guide Item Selection

Like the previous recommendation, this suggestion is tied to one of the study limitations. In the process of creating the test variants used to answer Research Question 2, item discrimination was the only criterion considered in selecting items. As mentioned previously, in reality, it is much more likely that additional factors would be considered when determining the suitability of items to be included in the final version of the examination.

Test specifications, or blueprints, stipulate the degree to which content standards should be covered by examination items. The specifications are important because they tie content standards to the knowledge, skills, and attributes measured by the examination. The content standards are also important because, if developed correctly, they support the purpose of the examination, and, consequently, serve as sources of test validity. Test specifications
were not available for the examinations used in the study. Future research might compare the results of this study with procedures that incorporate test specifications. Had test blueprints been available, the number of items unique to the Form B variants might have been different. A comparison of this nature would provide more clarity to the findings of this study.

Develop Longer Tests to Assess Effects of Restricted Discrimination Values
Another recommendation for possible future research is to increase the scope of the comparison between examinations developed using restricted item discrimination values and those developed using unrestricted values. Whereas the procedures associated with the current study included the creation of two 50-item examinations, the development of larger examinations might provide more insight into the differences in examination reliability and classification decision consistency estimates.

According to Crocker and Algina (2008), test length is one of several factors that affect examination reliability estimates, with longer tests generally producing larger reliability estimates. The differences in examination reliability between the 50 -item test variants considered in this study were found to be significantly greater than zero for each examination. Further research, however, might explore the degree to which these findings were consistent with much larger examinations. For example, test variants with between 100 and 200 items might produce reliability estimates that are much closer, and, potentially, whose differences are not found to be significantly greater than zero. This may also be true for classification decision consistency.

## Use of Non-Classical Test Theory Approaches

The context under which this study was conducted purposefully attempted to replicate the realistic conditions faced by many developers of credentialing examinations. That is, it utilized examinations with relatively small sample sizes, which necessitated the use of classical test theory. By doing so, the results of this study may be more generalizable to those involved with the development of credentialing examinations.

Research aimed at other approaches used to calculate item discrimination values, however, might help to expand the understanding of restricted discrimination values. One such approach, which is frequently used in larger-scale testing programs, is item response theory. As discussed in Chapter 2, item response theory is a general statistical theory that relates performance on test items to the abilities the test is intended to measure (Hambleton \& Jones, 1993). Two- and three-parameter logistic item response models may be used to estimate item discrimination. These procedures, however, generally require much larger sample sizes than classical test theory approaches require. Some have argued that a minimum of 500 cases is required to produce dependable estimates (Reise \& Yu, 1990). Comparing the results of this study with a similar study conducted using item response theory may help in gaining a better understanding of restricted item discrimination values.

## Effects of Restricted Discrimination Values on Standard Setting

A final recommendation for future research involves studying the degree to which using restricted item discrimination values in the selection of items affects the location of cut scores, as determined by the standard setting process. Using restricted discrimination values to select items, as documented in this study, resulted in the selection of items that were
different than those selected had unrestricted values been used. A key component in many standard-setting procedures is the requirement for panelists to make judgments regarding the likelihood that a hypothetical examinee would respond correctly to examination items. Whereas the use of restricted discrimination values results in the selection of a different set of items, it is quite possible that this may also affect the location of the examination cut score. Further research might examine how using restricted discrimination values in the test development process affects the placement of examination cut scores.

## Practical Implications for Test Developers

The results of this study present several practical implications for licensure and certification agencies, test developers, and other entities responsible for administering credentialing examinations. First, limiting the calculation of item discrimination values to examinee scores near cut scores will likely result in lower point-biserial values. As discussed previously, this decrease, in large part, is due to the lower variance associated with scores from much more homogeneous groups of examinees. Test developers should consider this implication when making a determination to use either restricted or unrestricted values. In many cases, the lower point-biserials produced by restricted values may not be considered problematic when viewed within the context of the purpose of the examination.

Second, and possibly most important, test developers must understand that using restricted point-biserials as a criterion for item selection, as opposed to the traditionally used unrestricted values, will likely change the content of the examination. Even when examination specifications are considered in the item selection process, using restricted point-biserials as a criterion of selection will likely result in the selection of items that would
not have been included using unrestricted values. Test developers must evaluate how these differences affect the examination's coverage of content standards. Again, in many cases, particularly those in which the examination is unidimensional in nature, the content change affected by using restricted item discrimination values may not be problematic. In cases where test specifications require the inclusion of items dedicated to two or more content domains, however, the changes may require a closer evaluation.

The degree to which using restricted item discrimination values as a criterion for item selection affects test content may serve as an important element of evidence when test developers make a case for examination validity. For credentialing examinations, the interpretation of test scores, the primary concern of validity, is that those who pass the examination are qualified to receive the credential in question, whereas those who do not pass are not. Selecting items using restricted item discrimination values results in the inclusion of items that are more discriminating for those candidates with total score near the examination cut score. As such, the items selected support the interpretation of the scores and, therefore, strengthen the case for examination validity.

Third, test developers should also consider the effect using restricted item discrimination values has on examination reliability and classification decision consistency. The results of the study suggested that using restricted point-biserials as the criterion by which items are selected resulted in lower levels of reliability and decision consistency. As discussed, however, the observed differences were relatively minor. In many cases, the importance placed on obtaining high levels of reliability and decision consistency may determine the decision to use restricted or unrestricted discrimination values when considering items for inclusion. As the Standards reiterate however, scoring precision for
credentialing examinations may be focused on scores near the cut score. Including items that discriminate better among those examinees with scores near the cut score may outweigh the relatively lower reliability and classification decision estimates observed in this study.

Ultimately, therefore, developers of credentialing examinations will need to weigh the advantages associated with using restricted discrimination values, mainly the potential for increased validity evidence, with the perceived disadvantages, primarily lower reliability and decision consistency estimates, when determining which type of discrimination value to use. The context of the examination, to include its overriding purpose and the role it plays in the credentialing process, will likely play a significant role in this determination. If governing bodies view the interpretation of test scores as a high priority, using restricted discrimination values may be appropriate. If, however, greater emphasis is placed on test statistics such as reliability and decision consistency, unrestricted discrimination values may be more suitable.

## Conclusion

This study examined the degree to which limiting the scores upon which item discrimination values are calculated to those at or near anticipated cut scores affected item discrimination values, referred to here as restricted item discrimination values; item selection; examination reliability; and classification decision consistency. For each examination used in this study, 15 sets of restricted discrimination values were calculated. The sets of values were based on scores within $0.50 S D, 0.75 S D$, and $1.00 S D$ of five unique cut score locations. For Examinations 1 and 2, the difference between each set of restricted values and the examination's unrestricted set was found to be significantly greater than zero. For Examination 3, the difference between 11 of the 15 sets of restricted values and the
unrestricted set were found be significantly greater than zero. In each case in which a significant difference was identified, the mean restricted discrimination value was smaller than the mean unrestricted value.

An evaluation of the effect restricted item discrimination values had on item selection, examination reliability, and classification decision consistency was conducted through the creation of Form A and Form B test variants for each examination. Form A test variants included the 50 most discriminating items using unrestricted discrimination values as the criterion for item selection. Form B variants included the 50 most discriminating items using restricted values (based on scores within $1.00 S D$ of the actual cut score, $\mathrm{C}_{\mathrm{X} 1}$ ). Using restricted values to select items resulted in Form B test variants including many items that were not included in Form A variants. The selection of unique items directly relates to the validity of credentialing examinations, as it places increased emphasis on examinees with scores closest to the cut score.

In terms of examination reliability, the study indicated that using restricted discrimination values to select items resulted in scores that produced lower examination reliability than scores derived from items selected using unrestricted discrimination values. For each examination, the difference between test variant reliability estimates were found to be significantly greater than zero, with Form B variants producing lower estimates. Although the differences were statistically significant, in practice, they were relatively small. Similar outcomes were observed with respect to classification decision consistency. Although the degree to which differences in test variant coefficients $\kappa$ were significantly greater than zero was not determined, their respective $95 \%$ confidence intervals overlapped.

The observed Form B variant estimates for each examination were slightly smaller than the Form A variants.

In conclusion, the research found that the use of restricted item discrimination values resulted in the selection of different items than those that would have been selected had unrestricted values been used. The validity-based benefits of using restricted values, namely that doing so increases focus on scores nearest the cut score, appeared to come at the cost of slight decreases in examination reliability and classification decision consistency. When considering the use of restricted item discrimination values, therefore, those who develop credentialing examinations must consider and prioritize these factors. This decision will most likely be tied to the purpose of the examination.

## APPENDIX A: ITEM DISCRIMINATION VALUES - EXAMINATION 1

Table A. 1
Item Discrimination Values for Examination $1-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.1114 | 0.2241 |
| 145 | 0.1624 | 0.0987 | 0.0788 | 0.0117 |
| 825 | 0.2558 | 0.0557 | 0.1366 | 0.0797 |
| 890 | 0.2829 | 0.1469 | 0.0874 | 0.1350 |
| 1115 | 0.2312 | 0.0308 | 0.1013 | 0.0928 |
| 1530 | 0.2356 | 0.1130 | 0.0093 | 0.0623 |
| 1660 | 0.1695 | 0.0645 | 0.1428 | 0.1996 |
| 1870 | 0.3104 | 0.1614 | 0.0158 | 0.0132 |
| 2030 | 0.1547 | 0.0718 | 0.1565 | 0.1111 |
| 2175 | 0.3985 | 0.2273 | 0.1339 | 0.0752 |
| 2250 | 0.3256 | 0.2055 | 0.0043 | -0.0387 |
| 2385 | 0.2968 | 0.1232 | 0.2753 | 0.2452 |
| 2545 | 0.3718 | 0.2382 | 0.0911 | 0.1040 |
| 2725 | 0.1192 | 0.1403 | 0.1721 | 0.0714 |
| 2830 | 0.3376 | 0.1948 | 0.0897 | 0.0479 |
| 2900 | 0.2990 | 0.1925 | 0.0787 | 0.0491 |
| 3115 | 0.2302 | 0.0406 | 0.1061 | -0.0048 |
| 3520 | 0.2361 | 0.1682 | 0.2291 | 0.1028 |
| 3535 | 0.3454 | 0.2758 | 0.2444 | 0.2120 |
| 3570 | 0.2944 | 0.1975 | 0.0584 | -0.0889 |
| 4300 | 0.2130 | 0.0329 | 0.1897 | 0.1170 |
| 5925 | 0.3991 | 0.1452 | 0.1220 | 0.0139 |
| 6620 | 0.2502 | 0.2016 | 0.1154 | 0.0537 |
| 6665 | 0.2674 | 0.1608 | 0.1235 | 0.1029 |
| 6710 | 0.3363 | 0.1617 | 0.3085 | 0.1837 |
| 6785 | 0.4877 | 0.3462 | 0.0926 | 0.0735 |
| 7185 | 0.3084 | 0.1510 | 0.1279 | 0.0724 |
| 7240 | 0.4227 | 0.2551 | 0.0348 | 0.1292 |
| 7420 | 0.1212 | 0.1202 | 0.2751 |  |
| 8120 | 0.4633 | 0.2837 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 1 (continued)
Item Discrimination Values for Examination $1-C_{X I}$

| Item | Unrestricted | +/-1.00 SD | +/-0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 10005 | 0.3567 | 0.3014 | 0.1700 | 0.0044 |
| 10560 | 0.3234 | 0.2235 | 0.2336 | 0.0802 |
| 10645 | 0.4002 | 0.3023 | 0.2222 | 0.2182 |
| 10690 | 0.1914 | 0.0896 | 0.0643 | -0.0014 |
| 11210 | 0.1892 | 0.0702 | -0.0028 | -0.0327 |
| 11490 | 0.2190 | 0.1661 | 0.1468 | 0.1318 |
| 12530 | 0.2701 | 0.1369 | 0.1170 | 0.0516 |
| 12585 | 0.2238 | 0.0730 | 0.0701 | 0.0895 |
| 12770 | 0.2145 | 0.0937 | 0.0601 | 0.0265 |
| 12895 | 0.3307 | 0.1609 | 0.2066 | 0.1886 |
| 12945 | 0.2775 | 0.0422 | 0.0815 | 0.1097 |
| 12960 | 0.3174 | 0.1018 | 0.0983 | 0.2026 |
| 12965 | 0.2288 | 0.1921 | 0.1806 | 0.2454 |
| 13020 | 0.2327 | 0.1244 | 0.1120 | 0.0650 |
| 13050 | 0.2285 | 0.0959 | 0.0840 | 0.1339 |
| 13245 | 0.3465 | 0.2137 | 0.1712 | 0.1869 |
| 13500 | 0.2145 | 0.0542 | 0.0405 | 0.0581 |
| 13695 | 0.2334 | 0.1855 | 0.1177 | 0.0510 |
| 14185 | 0.3771 | 0.2350 | 0.2052 | 0.1392 |
| 14275 | 0.3031 | 0.0783 | 0.0509 | 0.0572 |
| 15105 | 0.2166 | 0.0530 | 0.0231 | -0.0291 |
| 15145 | 0.1613 | 0.1444 | 0.1073 | 0.1553 |
| 15355 | 0.2608 | 0.1836 | 0.1558 | -0.0056 |
| 15395 | 0.2588 | 0.1855 | 0.1756 | 0.1678 |
| 15430 | 0.2713 | 0.2808 | 0.2170 | 0.1692 |
| 15450 | 0.4525 | 0.2385 | 0.2337 | 0.1822 |
| 15455 | 0.1563 | 0.1015 | 0.0911 | -0.0069 |
| 15500 | 0.1028 | 0.0205 | -0.0142 | 0.0659 |
| 15505 | 0.2527 | 0.0844 | 0.0506 | 0.1611 |
| 15535 | 0.1162 | 0.1225 | 0.1340 | 0.1591 |
| 15540 | 0.2146 | 0.1234 | 0.0619 | -0.0370 |

Table A. 1 (continued)
Item Discrimination Values for Examination $1-C_{X I}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 15545 | 0.1930 | 0.0318 | 0.0479 | 0.0822 |
| 15565 | 0.1988 | 0.1554 | 0.0278 | -0.0313 |
| 15570 | 0.2229 | 0.0549 | 0.1125 | 0.0738 |
| 15585 | 0.3128 | 0.1657 | 0.2323 | 0.1815 |
| 15600 | 0.3630 | 0.2735 | 0.0192 | 0.0106 |
| 15615 | 0.2029 | 0.0409 | 0.0670 | 0.0645 |
| 15630 | 0.2938 | 0.1218 | 0.2409 | 0.0931 |
| 15635 | 0.3756 | 0.2626 | 0.0688 | 0.0360 |
| 15645 | 0.2468 | 0.1369 | -0.0016 | 0.0440 |
| 15660 | 0.2246 | 0.0500 | -0.0330 | -0.0564 |
| 15665 | 0.1791 | 0.0233 | 0.0641 | 0.0909 |
| 15670 | 0.1945 | 0.0606 | 0.0942 | 0.0798 |
| 15680 | 0.1612 | 0.0810 | 0.0873 | 0.0710 |
| 15690 | 0.1971 | 0.0492 | 0.1600 | 0.1181 |
| 15705 | 0.3067 | 0.1425 | 0.1034 | 0.0850 |
| 15710 | 0.1927 | 0.0905 | 0.0494 | 0.0322 |
| 15715 | 0.1508 | 0.0410 | 0.1444 | 0.0858 |
| 15720 | 0.2855 | 0.1174 | 0.0670 | -0.0002 |
| 15725 | 0.2224 | 0.0687 | 0.0571 | 0.0140 |
| 15745 | 0.2033 | 0.0813 | 0.1914 | 0.0355 |
| 15755 | 0.2927 | 0.2352 | 0.1149 | 0.1097 |
| 15770 | 0.1920 | 0.1451 | 0.0666 |  |
| 15790 | 0.1749 | 0.0586 | 0.0165 | -0.0431 |
| 15795 | 0.0801 | 0.1052 | 0.0645 | -0.0088 |
| 15800 | 0.1837 | 0.0335 | 0.0331 | 0.0060 |
| 15810 | 0.3582 | 0.2594 | 0.1571 | 0.0006 |
| 15890 | 0.3484 | 0.2522 | 0.1991 | 0.1121 |
| 15905 | 0.4017 | 0.1613 | 0.1622 | 0.0135 |
| 15930 | 0.1898 | 0.0139 | 0.0777 | 0.0456 |
| 16005 | 0.3986 | 0.2250 | 0.1129 | -0.0491 |
| 16060 | 0.2868 | 0.2038 | 0.1613 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 1 (continued)
Item Discrimination Values for Examination $1-C_{X I}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16110 | 0.4243 | 0.2523 | 0.2227 | 0.2340 |
| 16135 | 0.2716 | 0.1932 | 0.2243 | 0.0902 |
| 16210 | 0.4188 | 0.2700 | 0.0322 | -0.0244 |
| 16255 | 0.1220 | 0.0461 | 0.0611 | 0.0115 |
| 16275 | 0.1715 | 0.0954 | 0.0993 | 0.1103 |
| 16280 | 0.2719 | 0.1153 | 0.1745 | 0.2014 |
| 16305 | 0.1803 | 0.1728 | 0.0292 | 0.0244 |
| 16375 | 0.3525 | 0.1046 | 0.0740 | 0.0672 |
| 16390 | 0.2068 | 0.0851 | 0.1186 | 0.1587 |
| 16415 | 0.2210 | 0.0733 | 0.1035 | 0.1435 |
| 16425 | 0.2634 | 0.1103 | 0.0682 | 0.0925 |
| 16435 | 0.1672 | 0.0830 | 0.1239 | 0.1025 |
| 16445 | 0.2965 | 0.1263 | 0.0311 | -0.0761 |
| 16470 | 0.2230 | 0.0479 | 0.0865 | 0.1695 |
| 16475 | 0.2791 | 0.1397 | 0.1309 | 0.1902 |
| 16505 | 0.1442 | 0.0806 | 0.0801 | 0.0210 |
| 16515 | 0.1771 | 0.0348 | -0.0137 | -0.0763 |
| 16525 | 0.0557 | -0.0528 | -0.0256 | -0.0366 |
| 16560 | 0.2017 | 0.0204 | 0.0993 | 0.0055 |
| 16605 | 0.0227 | 0.0216 | 0.2161 | 0.1762 |
| 16615 | 0.2484 | 0.1855 | 0.0513 | 0.0257 |
| 16635 | 0.2581 | 0.1910 | 0.1366 | 0.0717 |
| 16690 | 0.3027 | 0.1250 | 0.2027 | 0.0886 |
| 16710 | 0.4226 | 0.2092 | 0.0990 | 0.0825 |
| 16715 | 0.2471 | 0.1254 | 0.2374 |  |
| 16720 | 0.4695 | 0.2338 | 0.2002 | 0.1314 |
| 16725 | 0.3359 | 0.2670 | 0.2170 | 0.1545 |
| 16740 | 0.3407 | 0.1694 | 0.0780 |  |
| 16745 | 0.2644 | 0.1405 | 0.0569 |  |
| 16755 | 0.3060 |  | 0.0471 |  |
| 16760 |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 1 (continued)
Item Discrimination Values for Examination $1-C_{X I}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16765 | 0.3531 | 0.1780 | 0.0903 | 0.2293 |
| 16775 | 0.2067 | 0.1331 | 0.1151 | 0.0654 |
| 16800 | 0.2626 | 0.0996 | 0.0142 |  |
| 16840 | 0.3550 | 0.1438 | 0.1190 | 0.1118 |
| 16845 | 0.4353 | 0.2710 | 0.2696 | 0.2051 |
| 16895 | 0.1322 | 0.0451 | 0.1059 | 0.0709 |
| 16910 | 0.3609 | 0.1933 | 0.1470 | 0.1142 |
| 16925 | 0.1134 | 0.0506 | 0.0208 | 0.0570 |
| 16940 | 0.2813 | 0.1988 | 0.1574 | 0.1041 |
| 16945 | 0.2289 | 0.0703 | 0.0768 | 0.0048 |
| 16955 | 0.1375 | 0.1015 | 0.0932 | 0.0121 |
| 17010 | 0.2586 | 0.2537 | 0.2250 | 0.0631 |
| 17020 | 0.2144 | 0.1154 | 0.1457 | 0.1492 |
| 17050 | 0.1092 | 0.0399 | 0.0387 | 0.0510 |
| 17105 | 0.2338 | -0.0138 | -0.0721 | 0.0181 |
| 17115 | 0.2431 | 0.1210 | 0.0714 | 0.0423 |
| 17125 | 0.1825 | 0.1004 | 0.0416 | 0.0955 |
| 17130 | 0.1915 | 0.1213 | 0.1441 | 0.0204 |
| 17160 | 0.0628 | 0.1059 | 0.0027 | -0.1089 |
| 17240 | 0.3277 | 0.2031 | 0.1609 | -0.0280 |
| 17245 | 0.2662 | 0.0308 | 0.0539 | 0.0068 |
| 17265 | 0.2291 | 0.1284 | 0.0698 | 0.1021 |
| 17275 | -0.0524 | 0.0225 | 0.0134 | -0.0028 |
| 17285 | 0.2758 | 0.1224 | 0.1213 | -0.0014 |
| 17300 | 0.1741 | 0.0614 | 0.0880 | 0.0059 |
| 17335 | 0.2070 | 0.1040 | 0.0056 | 0.0752 |
| 17340 | 0.1773 | 0.1084 | 0.0627 | 0.0249 |
| 17345 | 0.2649 | 0.1496 | 0.1161 | 0.1183 |
| 17370 | 0.3665 | 0.1429 | 0.1507 | 0.1042 |
| 17380 | 0.1734 | 0.0003 | -0.0097 | -0.1235 |
| 17385 | 0.2695 | 0.1863 | 0.1517 | 0.1290 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 1 (continued)
Item Discrimination Values for Examination $1-C_{X I}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.0378 | -0.0394 |
| 17425 | 0.2585 | 0.1166 | 0.1369 | 0.1246 |
| 17430 | 0.3883 | 0.1941 | 0.0757 | 0.1310 |
| 17460 | 0.2122 | 0.0855 | 0.0432 | 0.0844 |
| 17480 | 0.2261 | 0.1164 | 0.0546 | 0.0551 |
| 17495 | 0.2081 | 0.0939 | 0.0487 | 0.0477 |
| 17500 | 0.3347 | 0.1333 | 0.0497 | 0.0761 |
| 17510 | 0.1919 | 0.0776 | 0.0710 | 0.0564 |
| 17520 | 0.1498 | 0.1070 | 0.0226 | -0.0314 |
| 17545 | 0.1543 | 0.0686 | 0.0334 | 0.0302 |
| 17550 | 0.0717 | -0.0323 | 0.0707 | 0.1405 |
| 17565 | 0.1601 | 0.0476 | 0.1530 | 0.0664 |
| 17580 | 0.3078 | 0.1574 | 0.0325 | 0.0002 |
| 17585 | 0.2006 | 0.0743 | 0.1310 | -0.0137 |
| 17595 | 0.2681 | 0.1883 | 0.0045 | 0.1166 |
| 17605 | 0.1571 | 0.0292 | 0.0012 | 0.0259 |
| 17635 | 0.1269 | 0.0262 | 0.0598 | -0.0144 |
| 17640 | 0.2316 | 0.1293 | 0.1449 | 0.0712 |
| 17660 | 0.1847 | 0.1130 | 0.1156 | 0.1005 |
| 17665 | 0.2484 | 0.1319 | 0.0999 | 0.0449 |
| 17670 | 0.2394 | 0.1224 | -0.0415 |  |
| 17690 | 0.1535 | 0.0662 | 0.0695 | 0.0614 |
| 17695 | 0.1150 | 0.0998 | 0.1010 |  |
|  |  |  |  |  |

Table A. 2
Item Discrimination Values for Examination 1 - $C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 145 | 0.1624 | 0.0569 | 0.0586 | 0.0681 |
| 825 | 0.2558 | 0.0822 | 0.0859 | 0.0807 |
| 890 | 0.2829 | 0.1564 | 0.1181 | 0.0248 |
| 1115 | 0.2312 | 0.1276 | 0.0469 | -0.01387 |
| 1530 | 0.2356 | 0.1388 | 0.0822 | 0.1085 |
| 1660 | 0.1695 | 0.0943 | 0.1646 | 0.1682 |
| 1870 | 0.3104 | 0.1468 | 0.1434 | 0.2124 |
| 2030 | 0.1547 | 0.1012 | 0.1709 | 0.1540 |
| 2175 | 0.3985 | 0.2511 | 0.1220 | 0.0807 |
| 2250 | 0.3256 | 0.1841 | 0.0154 | -0.0248 |
| 2385 | 0.2968 | 0.0059 | 0.1439 | -0.0007 |
| 2545 | 0.3718 | 0.2301 | 0.0957 | 0.0601 |
| 2725 | 0.1192 | 0.0742 | 0.1284 | 0.0579 |
| 2830 | 0.3376 | 0.1591 | 0.1909 | 0.0857 |
| 2900 | 0.2990 | 0.2075 | 0.0377 | 0.0076 |
| 3115 | 0.2302 | 0.1105 | 0.1039 | 0.0459 |
| 3520 | 0.2361 | 0.1253 | 0.2222 | 0.2421 |
| 3535 | 0.3454 | 0.2579 | 0.2014 | 0.1934 |
| 3570 | 0.2944 | 0.2004 | -0.0417 | -0.0049 |
| 4300 | 0.2130 | 0.0226 | 0.1403 | 0.0437 |
| 5925 | 0.3991 | 0.2188 | 0.1020 | 0.1386 |
| 6620 | 0.2502 | 0.1597 | 0.1081 | 0.0113 |
| 6665 | 0.2674 | 0.1359 | 0.1846 | 0.0719 |
| 6710 | 0.3363 | 0.2728 | 0.2740 | 0.2486 |
| 6785 | 0.4877 | 0.3398 | 0.1107 | 0.0056 |
| 7185 | 0.3084 | 0.1153 | 0.1826 | 0.1593 |
| 7240 | 0.4227 | 0.2315 | -0.0299 | -0.0680 |
| 7420 | 0.1212 | 0.0088 | 0.1564 | 0.1503 |
| 8120 | 0.4633 | 0.2529 | 0.1869 | 0.1665 |
| 10005 | 0.3567 | 0.2596 | 0.1681 | 0.1668 |
| 10560 | 0.3234 | 0.2183 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 2 (continued)
Item Discrimination Values for Examination $1-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 10645 | 0.4002 | 0.2914 | 0.2172 | 0.0108 |
| 10690 | 0.1914 | 0.0941 | 0.0383 | 0.0750 |
| 11210 | 0.1892 | 0.0629 | 0.1413 | 0.0565 |
| 11490 | 0.2190 | 0.1542 | 0.0592 | 0.0532 |
| 12530 | 0.2701 | 0.1301 | 0.1164 | -0.0743 |
| 12585 | 0.2238 | 0.1316 | 0.0291 | 0.0369 |
| 12770 | 0.2145 | 0.0801 | 0.2280 | 0.1422 |
| 12895 | 0.3307 | 0.2014 | 0.1001 | 0.0092 |
| 12945 | 0.2775 | 0.1254 | 0.1069 | 0.0533 |
| 12960 | 0.3174 | 0.1090 | 0.1924 | 0.1524 |
| 12965 | 0.2288 | 0.2522 | 0.0819 | 0.0958 |
| 13020 | 0.2327 | 0.1459 | 0.1401 | 0.0717 |
| 13050 | 0.2285 | 0.1223 | 0.1354 | 0.1052 |
| 13245 | 0.3465 | 0.2274 | 0.0811 | 0.0352 |
| 13500 | 0.2145 | 0.0724 | 0.1201 | 0.0508 |
| 13695 | 0.2334 | 0.1949 | 0.1793 | 0.1722 |
| 14185 | 0.3771 | 0.2278 | 0.0126 | -0.0441 |
| 14275 | 0.3031 | 0.0295 | 0.0495 | -0.0201 |
| 15105 | 0.2166 | 0.0993 | 0.1489 | 0.0329 |
| 15145 | 0.1613 | 0.1112 | 0.1590 | 0.1580 |
| 15355 | 0.2608 | 0.2480 | 0.2425 | 0.1499 |
| 15395 | 0.2588 | 0.1981 | 0.1328 |  |
| 15430 | 0.2713 | 0.2893 | 0.2248 | 0.0433 |
| 15450 | 0.4525 | 0.2462 | 0.1829 | 0.1499 |
| 15455 | 0.1563 | 0.1021 | 0.1320 | -0.0211 |
| 15500 | 0.1028 | 0.0213 | 0.0910 | 0.0670 |
| 15505 | 0.2527 | 0.0833 | 0.0952 | -0.0628 |
| 15535 | 0.1162 | 0.0991 | 0.0344 | 0.0585 |
| 15540 | 0.2146 | 0.0242 | 0.0287 | 0.0988 |
| 15545 | 0.1930 | 0.0512 | 0.0653 |  |
| 15565 | 0.1988 | 0.1891 | 0.1771 | 0.0905 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 2 (continued)
Item Discrimination Values for Examination $1-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 15570 | 0.2229 | -0.0446 | 0.0034 | -0.0546 |
| 15585 | 0.3128 | 0.0578 | 0.079 | 0.2395 |
| 15600 | 0.3630 | 0.3022 | 0.0114 | -0.0455 |
| 15615 | 0.2029 | 0.1034 | 0.0302 | -0.0324 |
| 15630 | 0.2938 | 0.0846 | 0.1477 | 0.1207 |
| 15635 | 0.3756 | 0.2189 | 0.1892 | 0.2132 |
| 15645 | 0.2468 | 0.1214 | 0.1000 | 0.0757 |
| 15660 | 0.2246 | 0.0278 | 0.0046 | 0.0236 |
| 15665 | 0.1791 | 0.0058 | 0.0645 | 0.0878 |
| 15670 | 0.1945 | 0.0650 | 0.0886 | 0.0594 |
| 15680 | 0.1612 | 0.1002 | -0.0250 | -0.0096 |
| 15690 | 0.1971 | 0.0591 | 0.1257 | 0.1744 |
| 15705 | 0.3067 | 0.1569 | 0.1249 | 0.0673 |
| 15710 | 0.1927 | 0.1083 | 0.1004 | 0.1365 |
| 15715 | 0.1508 | 0.0890 | 0.1075 | 0.0625 |
| 15720 | 0.2855 | 0.1034 | 0.0686 | 0.0553 |
| 15725 | 0.2224 | 0.0782 | 0.0585 | 0.0898 |
| 15745 | 0.2033 | 0.0769 | 0.1874 | 0.0951 |
| 15755 | 0.2927 | 0.2212 | 0.1015 | 0.0662 |
| 15770 | 0.1920 | 0.1273 | 0.1205 | -0.0049 |
| 15790 | 0.1749 | 0.0689 | 0.0342 | 0.0676 |
| 15795 | 0.0801 | 0.0364 | -0.0038 | -0.0354 |
| 15800 | 0.1837 | 0.0583 | 0.0915 | 0.1439 |
| 15810 | 0.3582 | 0.1832 | 0.0299 | 0.0852 |
| 15890 | 0.3484 | 0.1533 | 0.0986 | 0.0819 |
| 15905 | 0.4017 | 0.1554 | -0.0702 | 0.0586 |
| 15930 | 0.1898 | -0.0035 | 0.2313 | 0.1158 |
| 16005 | 0.3986 | 0.1720 | 0.2483 | 0.1306 |
| 16060 | 0.2868 | 0.1734 | 0.0524 | 0.0951 |
| 16110 | 0.4243 | 0.2461 |  |  |
| 16135 | 0.2716 | 0.1314 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 2 (continued)
Item Discrimination Values for Examination $1-C_{X 2}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 16210 | 0.4188 | 0.2236 | 0.1434 | 0.0675 |
| 16255 | 0.1220 | 0.0794 | 0.0657 | -0.0292 |
| 16275 | 0.1715 | 0.0896 | 0.0159 | 0.0851 |
| 16280 | 0.2719 | 0.1383 | 0.1093 | 0.0597 |
| 16305 | 0.1803 | 0.1924 | 0.1786 | 0.1146 |
| 16375 | 0.3525 | 0.0394 | 0.0232 | 0.0078 |
| 16390 | 0.2068 | 0.1290 | 0.2056 | 0.1613 |
| 16415 | 0.2210 | 0.0584 | 0.0624 | 0.0207 |
| 16425 | 0.2634 | 0.1260 | 0.0707 | 0.0361 |
| 16435 | 0.1672 | 0.1614 | 0.1378 | 0.0567 |
| 16445 | 0.2965 | 0.0898 | 0.1567 | 0.1322 |
| 16470 | 0.2230 | 0.0764 | 0.0099 | -0.0402 |
| 16475 | 0.2791 | 0.1014 | 0.0883 | -0.0315 |
| 16505 | 0.1442 | 0.0937 | 0.1442 | 0.1762 |
| 16515 | 0.1771 | 0.0177 | 0.0448 | 0.0582 |
| 16525 | 0.0557 | -0.0036 | 0.0038 | -0.0414 |
| 16560 | 0.2017 | 0.0466 | -0.0084 | -0.0655 |
| 16605 | 0.0227 | 0.0125 | -0.0608 | -0.0332 |
| 16615 | 0.2484 | 0.2078 | 0.1635 | 0.2102 |
| 16635 | 0.2581 | 0.0308 | 0.0928 | 0.1026 |
| 16690 | 0.3027 | 0.1090 | 0.0563 | -0.0667 |
| 16710 | 0.4226 | 0.2040 | 0.1681 | 0.1408 |
| 16715 | 0.2471 | 0.1422 | 0.1150 | 0.1434 |
| 16720 | 0.4695 | 0.2503 | 0.1414 | 0.1803 |
| 16725 | 0.3359 | 0.2414 | 0.2626 | 0.3158 |
| 16740 | 0.3407 | 0.2174 | 0.1859 | 0.1624 |
| 16745 | 0.2644 | 0.1664 | 0.1095 | 0.0237 |
| 16755 | 0.2363 | 0.1324 | 0.0804 | 0.0205 |
| 16760 | 0.3060 | 0.1280 | 0.0616 | 0.0772 |
| 16765 | 0.3531 | 0.1070 | 0.1658 | 0.0452 |
| 16775 | 0.2067 | 0.1953 | 0.1598 | 0.1111 |

Table A. 2 (continued)
Item Discrimination Values for Examination $1-C_{X 2}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 16800 | 0.2626 | 0.1244 | 0.0672 | 0.0886 |
| 16840 | 0.3550 | 0.1205 | 0.0740 | 0.1393 |
| 16845 | 0.4353 | 0.2553 | 0.2359 | 0.1161 |
| 16895 | 0.1322 | 0.0840 | 0.1307 | 0.0133 |
| 16910 | 0.3609 | 0.1927 | 0.1711 | 0.1150 |
| 16925 | 0.1134 | 0.0861 | 0.0643 | -0.0075 |
| 16940 | 0.2813 | 0.1650 | 0.1874 | 0.1787 |
| 16945 | 0.2289 | 0.0630 | 0.0561 | 0.0568 |
| 16955 | 0.1375 | 0.1054 | 0.0855 | 0.1842 |
| 17010 | 0.2586 | 0.1811 | 0.0927 | 0.1929 |
| 17020 | 0.2144 | 0.1732 | 0.2085 | 0.1049 |
| 17050 | 0.1092 | 0.0615 | 0.0402 | 0.0376 |
| 17105 | 0.2338 | 0.0115 | -0.0374 | -0.1434 |
| 17115 | 0.2431 | 0.0537 | 0.0021 | -0.0409 |
| 17125 | 0.1825 | 0.0764 | 0.0543 | 0.1113 |
| 17130 | 0.1915 | 0.1151 | 0.0356 | 0.1070 |
| 17160 | 0.0628 | 0.0838 | 0.0327 | 0.1226 |
| 17240 | 0.3277 | 0.1606 | 0.1184 | 0.1620 |
| 17245 | 0.2662 | 0.1052 | 0.0500 | 0.1546 |
| 17265 | 0.2291 | 0.1357 | 0.0975 | -0.0071 |
| 17275 | -0.0524 | 0.0646 | 0.0011 | 0.0113 |
| 17285 | 0.2758 | 0.1682 | 0.0747 | 0.1061 |
| 17300 | 0.1741 | 0.0590 | 0.0163 | 0.0228 |
| 17335 | 0.2070 | 0.0273 | 0.0675 | 0.0237 |
| 17340 | 0.1773 | 0.1121 | 0.1105 | 0.0678 |
| 17345 | 0.2649 | 0.1634 | 0.1149 | 0.1454 |
| 17370 | 0.3665 | 0.1842 | 0.1675 | 0.1755 |
| 17380 | 0.1734 | -0.0353 | -0.1138 | -0.0456 |
| 17385 | 0.2695 | 0.1344 | 0.1355 | 0.1609 |
| 17425 | 0.2585 | 0.0531 | 0.0496 | 0.0563 |
| 17430 | 0.3883 | 0.1584 | 0.0991 | 0.0306 |

Table A. 2 (continued)
Item Discrimination Values for Examination 1 - $C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 17460 | 0.2122 | 0.1120 | 0.0464 | -0.0266 |
| 17480 | 0.2261 | 0.1542 | 0.1835 | 0.0457 |
| 17495 | 0.2081 | 0.0552 | 0.0452 | 0.0304 |
| 17500 | 0.3347 | 0.0832 | -0.0147 | -0.0760 |
| 17510 | 0.1919 | 0.1132 | 0.1074 | 0.1114 |
| 17520 | 0.1498 | 0.0951 | 0.1224 | 0.0837 |
| 17545 | 0.1543 | 0.0670 | 0.0068 | 0.0042 |
| 17550 | 0.0717 | 0.0530 | 0.0438 | -0.0259 |
| 17565 | 0.1601 | 0.0825 | 0.0376 | -0.0162 |
| 17580 | 0.3078 | 0.1566 | 0.1538 | 0.0810 |
| 17585 | 0.2006 | 0.0210 | 0.0300 | 0.0568 |
| 17595 | 0.2681 | 0.0957 | 0.0331 | 0.0731 |
| 17605 | 0.1571 | 0.0630 | 0.1039 | 0.0287 |
| 17635 | 0.1269 | 0.0557 | 0.0104 | -0.0463 |
| 17640 | 0.2316 | 0.0753 | 0.0795 | 0.0011 |
| 17660 | 0.1847 | 0.1462 | 0.1284 | 0.1978 |
| 17665 | 0.2484 | 0.0991 | 0.0800 | 0.0288 |
| 17670 | 0.2394 | 0.1490 | 0.1667 | 0.2151 |
| 17690 | 0.1535 | 0.0732 | 0.1231 | 0.1294 |
| 17695 | 0.1150 | 0.0312 | 0.0880 | 0.0903 |
|  |  |  |  |  |

Table A. 3
Item Discrimination Values for Examination $1-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 145 | 0.1624 | 0.1346 | 0.1505 | 0.0211 |
| 825 | 0.2558 | 0.1676 | 0.0084 | -0.0661 |
| 890 | 0.2829 | 0.2156 | 0.1649 | 0.0511 |
| 1115 | 0.2312 | 0.1294 | 0.0277 | 0.0325 |
| 1530 | 0.2356 | 0.1646 | 0.0935 | 0.0973 |
| 1660 | 0.1695 | 0.0472 | 0.0200 | -0.0537 |
| 1870 | 0.3104 | 0.1813 | 0.1239 | 0.0447 |
| 2030 | 0.1547 | 0.0516 | 0.0056 | -0.1473 |
| 2175 | 0.3985 | 0.2342 | 0.1727 | -0.0021 |
| 2250 | 0.3256 | 0.1518 | 0.1947 | 0.1394 |
| 2385 | 0.2968 | 0.1861 | 0.1564 | 0.1539 |
| 2545 | 0.3718 | 0.2975 | 0.2994 | 0.1919 |
| 2725 | 0.1192 | 0.0615 | 0.0868 | 0.0896 |
| 2830 | 0.3376 | 0.1953 | 0.1940 | 0.1470 |
| 2900 | 0.2990 | 0.1693 | 0.1139 | 0.0549 |
| 3115 | 0.2302 | 0.0744 | -0.0051 | -0.0122 |
| 3520 | 0.2361 | 0.1035 | 0.1386 | 0.1388 |
| 3535 | 0.3454 | 0.2713 | 0.1267 | 0.0693 |
| 3570 | 0.2944 | 0.2465 | 0.1322 | 0.1161 |
| 4300 | 0.2130 | 0.1419 | -0.0048 | 0.1544 |
| 5925 | 0.3991 | 0.1771 | 0.1445 | 0.1788 |
| 6620 | 0.2502 | 0.1309 | 0.0914 | 0.1193 |
| 6665 | 0.2674 | 0.1078 | 0.1169 | 0.1347 |
| 6710 | 0.3363 | 0.1457 | 0.1328 | 0.0503 |
| 6785 | 0.4877 | 0.3075 | 0.2273 | 0.2383 |
| 7185 | 0.3084 | 0.2279 | 0.1615 | 0.0948 |
| 7240 | 0.4227 | 0.2909 | 0.1940 | 0.0071 |
| 7420 | 0.1212 | 0.1199 | 0.1840 | 0.0655 |
| 8120 | 0.4633 | 0.2710 | 0.2475 | 0.2077 |
| 10005 | 0.3567 | 0.2196 | 0.0855 | 0.1319 |
| 10560 | 0.3234 | 0.2005 | 0.1310 | 0.1753 |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 3 (continued)
Item Discrimination Values for Examination $1-C_{X 3}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 10645 | 0.4002 | 0.2879 | 0.2730 | 0.3180 |
| 10690 | 0.1914 | 0.1130 | 0.0548 | 0.0196 |
| 11210 | 0.1892 | 0.0875 | 0.0586 | -0.0204 |
| 11490 | 0.2190 | 0.0735 | 0.1106 | 0.1890 |
| 12530 | 0.2701 | 0.2103 | 0.1477 | 0.1876 |
| 12585 | 0.2238 | 0.0834 | 0.0909 | 0.1323 |
| 12770 | 0.2145 | 0.1828 | 0.1563 | -0.0030 |
| 12895 | 0.3307 | 0.1737 | 0.0571 | 0.0795 |
| 12945 | 0.2775 | 0.0157 | 0.0199 | 0.0732 |
| 12960 | 0.3174 | 0.1701 | 0.1278 | 0.0726 |
| 12965 | 0.2288 | 0.1492 | 0.1777 | 0.0322 |
| 13020 | 0.2327 | 0.1367 | 0.0755 | 0.0275 |
| 13050 | 0.2285 | 0.1239 | 0.0878 | -0.0187 |
| 13245 | 0.3465 | 0.2067 | 0.1887 | 0.1244 |
| 13500 | 0.2145 | 0.0503 | 0.0432 | -0.0426 |
| 13695 | 0.2334 | 0.0990 | 0.0919 | 0.1156 |
| 14185 | 0.3771 | 0.2586 | 0.1969 | 0.1702 |
| 14275 | 0.3031 | 0.1733 | 0.1463 | 0.1158 |
| 15105 | 0.2166 | 0.0379 | 0.0021 | 0.0373 |
| 15145 | 0.1613 | 0.0918 | 0.0532 | 0.0959 |
| 15355 | 0.2608 | 0.1359 | 0.0582 | 0.0515 |
| 15395 | 0.2588 | 0.1882 | 0.1319 | 0.1348 |
| 15430 | 0.2713 | 0.2252 | 0.2117 | 0.1501 |
| 15450 | 0.4525 | 0.3004 | 0.2573 | 0.2423 |
| 15455 | 0.1563 | 0.0836 | -0.0106 | -0.0032 |
| 15500 | 0.1028 | 0.0428 | -0.0104 | -0.0050 |
| 15505 | 0.2527 | 0.1030 | 0.1357 | -0.0001 |
| 15535 | 0.1162 | 0.1322 | 0.2029 | 0.2135 |
| 15540 | 0.2146 | 0.1791 | 0.1356 | 0.1005 |
| 15545 | 0.1930 | 0.0725 | 0.0836 | -0.0788 |
| 15565 | 0.1988 | 0.1182 | 0.1608 | 0.0618 |

Table A. 3 (continued)
Item Discrimination Values for Examination $1-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 15570 | 0.2229 | 0.1422 | 0.0526 | -0.0022 |
| 15585 | 0.3128 | 0.1837 | 0.1821 | 0.1162 |
| 15600 | 0.3630 | 0.2401 | 0.0031 | 0.0074 |
| 15615 | 0.2029 | 0.0485 | 0.1562 | 0.1057 |
| 15630 | 0.2938 | 0.1776 | 0.2081 | 0.2122 |
| 15635 | 0.3756 | 0.2724 | -0.0492 | -0.2087 |
| 15645 | 0.2468 | 0.0884 | 0.0203 | -0.0444 |
| 15660 | 0.2246 | 0.0785 | 0.0139 | -0.0371 |
| 15665 | 0.1791 | 0.1022 | 0.0226 | 0.0533 |
| 15670 | 0.1945 | 0.1314 | 0.0824 | 0.0363 |
| 15680 | 0.1612 | 0.0453 | 0.1108 | 0.1018 |
| 15690 | 0.1971 | 0.0686 | 0.0816 | 0.0651 |
| 15705 | 0.3067 | 0.1547 | 0.0521 | 0.0785 |
| 15710 | 0.1927 | 0.0966 | -0.0966 | -0.1205 |
| 15715 | 0.1508 | -0.0171 | 0.0953 | 0.0943 |
| 15720 | 0.2855 | 0.1579 | 0.0199 | 0.0656 |
| 15725 | 0.2224 | -0.0248 | 0.0430 | 0.0023 |
| 15745 | 0.2033 | 0.1265 | 0.1694 | 0.0958 |
| 15755 | 0.2927 | 0.1682 | 0.0877 | 0.1448 |
| 15770 | 0.1920 | 0.1560 | 0.0494 | 0.0211 |
| 15790 | 0.1749 | 0.0913 | 0.0681 | 0.0131 |
| 15795 | 0.0801 | 0.0511 | 0.0506 | 0.0471 |
| 15800 | 0.1837 | 0.1283 | 0.2258 | 0.1557 |
| 15810 | 0.3582 | 0.2133 | 0.2478 | 0.1932 |
| 15890 | 0.3484 | 0.2658 | 0.1149 | 0.0867 |
| 15905 | 0.4017 | 0.1553 | 0.0869 | 0.0590 |
| 15930 | 0.1898 | 0.0585 | 0.1114 | 0.0222 |
| 16005 | 0.3986 | 0.2270 | 0.1302 | 0.0563 |
| 16060 | 0.2868 | 0.1831 | 0.2037 | 0.1184 |
| 16110 | 0.4243 | 0.2460 | 0.1827 | 0.1982 |
| 16135 | 0.2716 | 0.1941 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 3 (continued)
Item Discrimination Values for Examination $1-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16210 | 0.4188 | 0.2911 | 0.2438 | 0.2194 |
| 16255 | 0.1220 | 0.0358 | 0.0782 | 0.0725 |
| 16275 | 0.1715 | 0.1281 | 0.1258 | -0.0041 |
| 16280 | 0.2719 | 0.1301 | 0.0780 | 0.0783 |
| 16305 | 0.1803 | 0.1263 | 0.1301 | 0.0545 |
| 16375 | 0.3525 | 0.1110 | 0.0825 | 0.0262 |
| 16390 | 0.2068 | 0.0390 | -0.0137 | -0.0583 |
| 16415 | 0.2210 | 0.0901 | 0.0923 | 0.0524 |
| 16425 | 0.2634 | 0.0865 | 0.1162 | 0.0985 |
| 16435 | 0.1672 | 0.0272 | -0.0071 | 0.0118 |
| 16445 | 0.2965 | 0.1442 | 0.0705 | -0.0212 |
| 16470 | 0.2230 | 0.0794 | 0.0017 | 0.0833 |
| 16475 | 0.2791 | 0.1861 | 0.1934 | 0.1253 |
| 16505 | 0.1442 | 0.0919 | 0.0539 | -0.0235 |
| 16515 | 0.1771 | 0.0268 | 0.0264 | 0.0213 |
| 16525 | 0.0557 | -0.0013 | -0.0689 | -0.0472 |
| 16560 | 0.2017 | 0.0391 | 0.0512 | 0.0897 |
| 16605 | 0.0227 | 0.0615 | 0.0638 | 0.1439 |
| 16615 | 0.2484 | 0.1888 | 0.1556 | 0.0546 |
| 16635 | 0.2581 | 0.1022 | 0.1570 | 0.0523 |
| 16690 | 0.3027 | 0.2042 | 0.1332 | 0.2600 |
| 16710 | 0.4226 | 0.2303 | 0.1388 | 0.1003 |
| 16715 | 0.2471 | 0.0953 | 0.0719 | 0.0292 |
| 16720 | 0.4695 | 0.2956 | 0.2353 | 0.1531 |
| 16725 | 0.3359 | 0.2369 | 0.1568 | -0.0343 |
| 16740 | 0.3407 | 0.2379 | 0.1548 | 0.0328 |
| 16745 | 0.2644 | 0.1473 | 0.1372 | 0.1648 |
| 16755 | 0.2363 | 0.1481 | 0.1060 | 0.1622 |
| 16760 | 0.3060 | 0.2510 | 0.1756 | 0.1428 |
| 16765 | 0.3531 | 0.2120 | 0.2155 | 0.0561 |
| 16775 | 0.2067 | 0.1021 | 0.0257 | 0.0668 |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 3 (continued)
Item Discrimination Values for Examination $1-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16800 | 0.2626 | 0.0907 | 0.0811 | -0.0197 |
| 16840 | 0.3550 | 0.1852 | 0.2602 | 0.0552 |
| 16845 | 0.4353 | 0.2708 | 0.0198 | 0.0306 |
| 16895 | 0.1322 | 0.0050 | 0.1732 | 0.0960 |
| 16910 | 0.3609 | 0.1964 | 0.0117 | 0.1129 |
| 16925 | 0.1134 | 0.0528 | 0.1071 | 0.0342 |
| 16940 | 0.2813 | 0.2190 | 0.0282 | 0.0515 |
| 16945 | 0.2289 | 0.1666 | 0.0314 | -0.0328 |
| 16955 | 0.1375 | 0.1039 | 0.1964 | 0.1457 |
| 17010 | 0.2586 | 0.2377 | 0.0304 | 0.0128 |
| 17020 | 0.2144 | 0.1577 | 0.0229 | 0.0292 |
| 17050 | 0.1092 | 0.0639 | 0.0982 | 0.0896 |
| 17105 | 0.2338 | -0.0110 | 0.1411 | 0.1783 |
| 17115 | 0.2431 | 0.1348 | 0.0777 | -0.0787 |
| 17125 | 0.1825 | 0.0728 | 0.1104 | 0.1207 |
| 17130 | 0.1915 | 0.1616 | 0.0363 | 0.0343 |
| 17160 | 0.0628 | 0.0652 | 0.1909 | 0.1432 |
| 17240 | 0.3277 | 0.2360 | -0.0071 | -0.0844 |
| 17245 | 0.2662 | 0.0479 | 0.1320 | 0.0842 |
| 17265 | 0.2291 | 0.0967 | 0.0423 | 0.0198 |
| 17275 | -0.0524 | 0.0094 | 0.0236 | 0.1954 |
| 17285 | 0.2758 | 0.1434 | 0.0979 | 0.1365 |
| 17300 | 0.1741 | 0.1414 | 0.0308 |  |
| 17335 | 0.2070 | 0.1178 | 0.1094 | 0.0062 |
| 17340 | 0.1773 | 0.1064 | 0.0692 | 0.0084 |
| 17345 | 0.2649 | 0.0987 | 0.0553 | -0.0362 |
| 17370 | 0.3665 | 0.1916 | 0.1137 | 0.0936 |
| 17380 | 0.1734 | 0.0482 | 0.0789 | 0.0807 |
| 17385 | 0.2695 | 0.2846 | 0.1601 | 0.0006 |
| 17425 | 0.2585 | 0.1367 | 0.1224 | 0.2077 |
| 17430 | 0.3883 | 0.1833 | 0.1730 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 3 (continued)
Item Discrimination Values for Examination $1-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 17460 | 0.2122 | 0.1606 | 0.1132 | -0.0897 |
| 17480 | 0.2261 | 0.1541 | 0.0544 | 0.0492 |
| 17495 | 0.2081 | 0.0680 | 0.1067 | 0.2167 |
| 17500 | 0.3347 | 0.1437 | 0.1639 | -0.1057 |
| 17510 | 0.1919 | 0.0698 | -0.0472 | -0.0752 |
| 17520 | 0.1498 | 0.0689 | 0.0336 | 0.0217 |
| 17545 | 0.1543 | 0.0397 | 0.0478 | -0.0574 |
| 17550 | 0.0717 | -0.0646 | -0.0919 | 0.0886 |
| 17565 | 0.1601 | 0.0836 | 0.0820 | 0.0991 |
| 17580 | 0.3078 | 0.1969 | 0.1041 | 0.0294 |
| 17585 | 0.2006 | 0.1920 | 0.0394 | 0.1293 |
| 17595 | 0.2681 | 0.2471 | 0.1450 | -0.1017 |
| 17605 | 0.1571 | 0.0444 | -0.0097 | 0.0147 |
| 17635 | 0.1269 | -0.0072 | 0.0417 | 0.1233 |
| 17640 | 0.2316 | 0.0929 | 0.1163 | -0.0463 |
| 17660 | 0.1847 | 0.1219 | -0.0055 | 0.1286 |
| 17665 | 0.2484 | 0.2351 | 0.1436 | -0.0490 |
| 17670 | 0.2394 | 0.0993 | -0.0118 | -0.0038 |
| 17690 | 0.1535 | 0.0787 | -0.0531 | 0.0677 |
| 17695 | 0.1150 | 0.0921 | 0.0798 |  |

Table A. 4
Item Discrimination Values for Examination $1-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  | -0.0148 |
| 45 | 0.1624 | 0.0326 | 0.0995 | 0.0223 |
| 825 | 0.2558 | 0.0989 | 0.0656 | -0.0168 |
| 1115 | 0.2829 | 0.1826 | 0.1577 | -0.0331 |
| 1530 | 0.2312 | 0.1566 | 0.0857 | 0.0361 |
| 1660 | 0.2356 | 0.1230 | 0.1432 | 0.0252 |
| 1870 | 0.1695 | 0.1765 | 0.1340 | 0.0406 |
| 2030 | 0.3104 | 0.1913 | 0.2141 | 0.0772 |
| 2175 | 0.1547 | 0.2095 | 0.2183 | 0.1125 |
| 2250 | 0.3985 | 0.2215 | 0.1438 | 0.0924 |
| 2385 | 0.3256 | 0.2215 | 0.0457 | 0.0653 |
| 2545 | 0.2968 | 0.0199 | 0.0301 | 0.0339 |
| 2725 | 0.3718 | 0.1682 | 0.0794 | 0.0483 |
| 2830 | 0.1192 | 0.1523 | 0.1248 | 0.1792 |
| 2900 | 0.3376 | 0.1354 | 0.2401 | 0.1677 |
| 3115 | 0.2990 | 0.2134 | 0.0868 | 0.0360 |
| 3520 | 0.2302 | 0.1251 | 0.0516 | 0.1141 |
| 3535 | 0.2361 | 0.1098 | 0.1892 | 0.2181 |
| 3570 | 0.3454 | 0.1755 | 0.1468 | 0.0716 |
| 4300 | 0.2944 | 0.2204 | -0.0164 | 0.0485 |
| 5925 | 0.2130 | -0.0008 | 0.1332 | 0.0659 |
| 6620 | 0.3991 | 0.1950 | 0.2307 | 0.1934 |
| 6665 | 0.2502 | 0.1430 | 0.0968 | 0.1117 |
| 6710 | 0.2674 | 0.1736 | 0.2709 | 0.1515 |
| 6785 | 0.3363 | 0.2705 | 0.3131 | 0.2464 |
| 7185 | 0.4877 | 0.3084 | 0.0359 | 0.0461 |
| 7240 | 0.3084 | 0.0519 | 0.2483 | 0.1913 |
| 7420 | 0.4227 | 0.2522 | -0.0616 | -0.0992 |
| 8120 | 0.1212 | 0.0309 | 0.1759 | 0.1598 |
| 10005 | 0.4633 | 0.2052 | 0.2550 | 0.3006 |
| 10560 | 0.3567 | 0.2582 | 0.1661 | 0.1171 |
|  | 0.3234 | 0.1828 |  |  |
|  |  |  |  |  |

Table A. 4 (continued)
Item Discrimination Values for Examination 1 - $C_{X 4}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 10645 | 0.4002 | 0.2247 | 0.1144 | 0.1737 |
| 10690 | 0.1914 | 0.0865 | 0.0868 | 0.1312 |
| 11210 | 0.1892 | 0.0950 | 0.1158 | 0.0745 |
| 11490 | 0.2190 | 0.1291 | 0.0534 | 0.0314 |
| 12530 | 0.2701 | 0.1152 | 0.1211 | 0.0226 |
| 12585 | 0.2238 | 0.1629 | 0.0876 | 0.0702 |
| 12770 | 0.2145 | 0.0470 | 0.0530 | -0.0357 |
| 12895 | 0.3307 | 0.1939 | 0.1061 | 0.1345 |
| 12945 | 0.2775 | 0.1289 | 0.0393 | 0.0235 |
| 12960 | 0.3174 | 0.1474 | 0.0656 | 0.0221 |
| 12965 | 0.2288 | 0.2315 | 0.1989 | 0.0258 |
| 13020 | 0.2327 | 0.1748 | 0.1422 | 0.1041 |
| 13050 | 0.2285 | 0.1346 | 0.0928 | 0.0084 |
| 13245 | 0.3465 | 0.2104 | 0.1490 | 0.0086 |
| 13500 | 0.2145 | 0.0895 | 0.0723 | 0.1320 |
| 13695 | 0.2334 | 0.1971 | 0.1363 | 0.1890 |
| 14185 | 0.3771 | 0.2071 | 0.1970 | 0.0404 |
| 14275 | 0.3031 | 0.0758 | 0.0207 | -0.0546 |
| 15105 | 0.2166 | 0.0867 | 0.0624 | 0.2087 |
| 15145 | 0.1613 | 0.0802 | -0.0216 | 0.0456 |
| 15355 | 0.2608 | 0.2392 | 0.2147 | 0.2004 |
| 15395 | 0.2588 | 0.1788 | 0.1737 | 0.0870 |
| 15430 | 0.2713 | 0.1985 | 0.2224 | 0.1284 |
| 15450 | 0.4525 | 0.1934 | 0.1283 | 0.0415 |
| 15455 | 0.1563 | 0.1192 | 0.1431 | 0.1636 |
| 15500 | 0.1028 | 0.0939 | 0.0097 | 0.0435 |
| 15505 | 0.2527 | 0.0795 | 0.0139 | -0.0478 |
| 15535 | 0.1162 | 0.0401 | -0.0130 | -0.1408 |
| 15540 | 0.2146 | 0.0461 | -0.0226 | 0.0182 |
| 15545 | 0.1930 | 0.1047 | 0.0600 | -0.1046 |
| 15565 | 0.1988 | 0.2201 | 0.1693 | 0.0799 |

Table A. 4 (continued)
Item Discrimination Values for Examination 1 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 15570 | 0.2229 | 0.0145 | -0.0041 | -0.0543 |
| 15585 | 0.3128 | 0.0735 | 0.2843 | 0.0000 |
| 15600 | 0.3630 | 0.3132 | 0.0835 | 0.0365 |
| 15615 | 0.2029 | 0.0596 | 0.0571 | 0.0444 |
| 15630 | 0.2938 | 0.1816 | 0.1690 | 0.1239 |
| 15635 | 0.3756 | 0.1656 | 0.1781 | 0.2585 |
| 15645 | 0.2468 | 0.1308 | 0.0734 | 0.1176 |
| 15660 | 0.2246 | 0.0899 | 0.0245 | 0.0969 |
| 15665 | 0.1791 | 0.0173 | 0.0758 | -0.0950 |
| 15670 | 0.1945 | 0.0503 | 0.0427 | 0.0655 |
| 15680 | 0.1612 | 0.0898 | -0.0294 | -0.0746 |
| 15690 | 0.1971 | 0.0378 | 0.1126 | 0.0072 |
| 15705 | 0.3067 | 0.1783 | 0.0625 | 0.0787 |
| 15710 | 0.1927 | 0.0318 | 0.1589 | 0.1383 |
| 15715 | 0.1508 | 0.1717 | 0.0532 | -0.0533 |
| 15720 | 0.2855 | 0.1135 | 0.1259 | 0.0676 |
| 15725 | 0.2224 | 0.1699 | 0.0930 | 0.1045 |
| 15745 | 0.2033 | 0.1261 | 0.1541 | 0.0844 |
| 15755 | 0.2927 | 0.2283 | 0.0698 | 0.0827 |
| 15770 | 0.1920 | 0.1359 | 0.0410 | 0.0854 |
| 15790 | 0.1749 | 0.0884 | 0.0725 | 0.0490 |
| 15795 | 0.0801 | 0.0866 | 0.0457 | 0.0342 |
| 15800 | 0.1837 | 0.0451 | 0.1768 | 0.1241 |
| 15810 | 0.3582 | 0.1834 | 0.0717 | 0.1043 |
| 15890 | 0.3484 | 0.1181 | 0.1263 | -0.0357 |
| 15905 | 0.4017 | 0.1452 | 0.0408 | 0.1966 |
| 15930 | 0.1898 | 0.0118 | 0.1664 | 0.1563 |
| 16005 | 0.3986 | 0.2598 | 0.1786 | 0.0028 |
| 16060 | 0.2868 | 0.1327 | 0.2512 |  |
| 16110 | 0.4243 | 0.0961 |  |  |
| 16135 | 0.2716 |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 4 (continued)
Item Discrimination Values for Examination $1-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16210 | 0.4188 | 0.2581 | 0.1577 | 0.0892 |
| 16255 | 0.1220 | 0.0755 | 0.0294 | 0.0982 |
| 16275 | 0.1715 | 0.0469 | 0.1043 | 0.0954 |
| 16280 | 0.2719 | 0.1762 | 0.1021 | 0.0779 |
| 16305 | 0.1803 | 0.1073 | 0.1295 | 0.1347 |
| 16375 | 0.3525 | 0.0567 | 0.0484 | 0.1594 |
| 16390 | 0.2068 | 0.1706 | 0.1907 | -0.1474 |
| 16415 | 0.2210 | 0.0267 | -0.0342 | -0.0686 |
| 16425 | 0.2634 | 0.1207 | 0.0242 | 0.1668 |
| 16435 | 0.1672 | 0.1233 | 0.1469 | 0.1901 |
| 16445 | 0.2965 | 0.1072 | 0.1003 | 0.1140 |
| 16470 | 0.2230 | 0.0575 | 0.0673 | 0.0101 |
| 16475 | 0.2791 | 0.0801 | -0.0211 | -0.0990 |
| 16505 | 0.1442 | 0.0696 | 0.0840 | -0.0221 |
| 16515 | 0.1771 | 0.1149 | 0.0150 | 0.1407 |
| 16525 | 0.0557 | -0.0293 | 0.0111 | 0.0970 |
| 16560 | 0.2017 | 0.0707 | 0.0525 | -0.1339 |
| 16605 | 0.0227 | -0.0482 | -0.0813 | 0.0330 |
| 16615 | 0.2484 | 0.2189 | 0.1551 | 0.1260 |
| 16635 | 0.2581 | 0.1025 | 0.0448 | 0.0294 |
| 16690 | 0.3027 | 0.0897 | 0.0054 | 0.1514 |
| 16710 | 0.4226 | 0.2242 | 0.1518 | 0.0755 |
| 16715 | 0.2471 | 0.1497 | 0.1661 | -0.0017 |
| 16720 | 0.4695 | 0.2330 | 0.1807 | 0.1544 |
| 16725 | 0.3359 | 0.3001 | 0.2878 | 0.1767 |
| 16740 | 0.3407 | 0.2449 | 0.2534 | 0.1000 |
| 16745 | 0.2644 | 0.0762 | 0.0717 | 0.1414 |
| 16755 | 0.2363 | 0.1424 | 0.0762 | 0.0266 |
| 16760 | 0.3060 | 0.1330 | 0.0617 | 0.1678 |
| 16765 | 0.3531 | 0.1375 | 0.0923 |  |
| 16775 | 0.2067 | 0.2230 | 0.2116 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 4 (continued)
Item Discrimination Values for Examination 1 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16800 | 0.2626 | 0.0970 | 0.1628 | 0.1449 |
| 16840 | 0.3550 | 0.1149 | 0.1788 | 0.0094 |
| 16845 | 0.4353 | 0.2818 | 0.0587 | 0.0552 |
| 16895 | 0.1322 | 0.1844 | 0.1848 | 0.1186 |
| 16910 | 0.3609 | 0.2179 | 0.0549 | 0.0370 |
| 16925 | 0.1134 | 0.0928 | 0.1416 | 0.2022 |
| 16940 | 0.2813 | 0.1501 | 0.0924 | 0.0836 |
| 16945 | 0.2289 | 0.0514 | 0.1383 | 0.1235 |
| 16955 | 0.1375 | 0.1421 | 0.1263 | 0.1055 |
| 17010 | 0.2586 | 0.1345 | 0.0890 | 0.1114 |
| 17020 | 0.2144 | 0.2127 | 0.0510 | 0.0330 |
| 17050 | 0.1092 | 0.0616 | 0.0050 | -0.0116 |
| 17105 | 0.2338 | 0.0377 | -0.0115 | 0.0248 |
| 17115 | 0.2431 | 0.0153 | 0.1069 | 0.0574 |
| 17125 | 0.1825 | 0.0681 | 0.1020 | -0.0490 |
| 17130 | 0.1915 | 0.0582 | 0.1400 | 0.0906 |
| 17160 | 0.0628 | 0.0169 | 0.1492 | 0.1000 |
| 17240 | 0.3277 | 0.1730 | 0.1844 | 0.0125 |
| 17245 | 0.2662 | 0.1353 | 0.1439 | 0.0659 |
| 17265 | 0.2291 | 0.1474 | 0.0543 | -0.0377 |
| 17275 | -0.0524 | 0.0648 | 0.1270 |  |
| 17285 | 0.2758 | 0.2104 | 0.064706 | 0.0649 |
| 17300 | 0.1741 | 0.0287 | 0.0128 | 0.0053 |
| 17335 | 0.2070 | 0.0732 | 0.0123 | 0.1771 |
| 17340 | 0.1773 | 0.0769 | 0.0959 | 0.1065 |
| 17345 | 0.2649 | 0.1441 | 0.1590 | 0.1023 |
| 17370 | 0.3665 | 0.1406 | 0.1705 | -0.0600 |
| 17380 | 0.1734 | 0.0088 | -0.0375 | 0.1138 |
| 17385 | 0.2695 | 0.1336 | 0.0970 | 0.0529 |
| 17425 | 0.2585 | 0.0750 | 0.1044 |  |
| 17430 | 0.3883 | 0.1796 | 0.1173 |  |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 4 (continued)
Item Discrimination Values for Examination 1 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.1071 | 0.0076 |
| 17460 | 0.2122 | 0.1257 | 0.1922 | 0.1261 |
| 17480 | 0.2261 | 0.2221 | 0.0691 | 0.0534 |
| 17495 | 0.2081 | 0.0932 | 0.0204 | 0.0081 |
| 17500 | 0.3347 | 0.0083 | 0.1084 | 0.1424 |
| 17520 | 0.1919 | 0.0858 | 0.0842 | 0.0735 |
| 17545 | 0.1498 | 0.1190 | 0.0717 | 0.0159 |
| 17550 | 0.1543 | 0.0778 | 0.0444 | 0.0280 |
| 17565 | 0.0717 | 0.1143 | 0.0149 | -0.1293 |
| 17580 | 0.1601 | 0.1187 | 0.1386 | 0.2217 |
| 17585 | 0.3078 | 0.1719 | 0.0245 | 0.0793 |
| 17595 | 0.2006 | -0.0229 | 0.0720 | 0.1373 |
| 17605 | 0.2681 | 0.0486 | 0.0639 | 0.0630 |
| 17635 | 0.1571 | 0.0292 | 0.0311 | -0.0093 |
| 17640 | 0.1269 | 0.1331 | 0.0423 | 0.1402 |
| 17660 | 0.2316 | 0.0818 | 0.1461 | 0.0449 |
| 17665 | 0.1847 | 0.0623 | 0.0502 | 0.0175 |
| 17670 | 0.2484 | 0.1176 | 0.1935 | 0.1705 |
| 17690 | 0.2394 | 0.2124 | 0.0693 | 0.2254 |
| 17695 | 0.1535 | 0.0896 | -0.0173 | 0.0181 |
|  | 0.1150 | 0.0741 |  |  |

Table A. 5
Item Discrimination Values for Examination 1 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 145 | 0.1624 | 0.0937 | 0.0477 | 0.0043 |
| 825 | 0.2558 | 0.1070 | 0.1200 | 0.0083 |
| 890 | 0.2829 | 0.1821 | 0.1828 | 0.1282 |
| 1115 | 0.2312 | 0.1220 | 0.0643 | -0.2205 |
| 1530 | 0.2356 | 0.1158 | 0.1591 | 0.1251 |
| 1660 | 0.1695 | 0.1065 | 0.0543 | -0.0052 |
| 1870 | 0.3104 | 0.1340 | 0.1337 | 0.0747 |
| 2030 | 0.1547 | 0.0205 | -0.0362 | -0.0914 |
| 2175 | 0.3985 | 0.2319 | 0.1610 | 0.1886 |
| 2250 | 0.3256 | 0.1481 | 0.0907 | 0.0973 |
| 2385 | 0.2968 | 0.2932 | 0.3219 | 0.2173 |
| 2545 | 0.3718 | 0.3704 | 0.2108 | 0.1134 |
| 2725 | 0.1192 | -0.0093 | 0.0169 | 0.1191 |
| 2830 | 0.3376 | 0.1837 | 0.1528 | 0.2363 |
| 2900 | 0.2990 | 0.1169 | 0.1306 | 0.0451 |
| 3115 | 0.2302 | 0.0790 | 0.0377 | -0.0131 |
| 3520 | 0.2361 | 0.1907 | 0.1368 | 0.0656 |
| 3535 | 0.3454 | 0.1666 | 0.1609 | 0.1436 |
| 3570 | 0.2944 | 0.1961 | 0.1219 | 0.0430 |
| 4300 | 0.2130 | 0.1235 | 0.1846 | 0.2335 |
| 5925 | 0.3991 | 0.2060 | 0.1819 | 0.0139 |
| 6620 | 0.2502 | 0.0796 | 0.1340 | 0.1647 |
| 6665 | 0.2674 | 0.1083 | 0.0794 | 0.0906 |
| 6710 | 0.3363 | 0.1231 | 0.1042 | 0.0503 |
| 6785 | 0.4877 | 0.2860 | 0.1998 | 0.1008 |
| 7185 | 0.3084 | 0.2583 | 0.2431 | 0.1661 |
| 7240 | 0.4227 | 0.2620 | 0.2692 | 0.1740 |
| 7420 | 0.1212 | 0.1857 | 0.1421 | 0.2106 |
| 8120 | 0.4633 | 0.3016 | 0.1914 | 0.1339 |
| 10005 | 0.3567 | 0.1271 | 0.1402 | 0.2053 |
| 10560 | 0.3234 | 0.1794 | 0.1085 | -0.0038 |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 5 (continued)
Item Discrimination Values for Examination $1-C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 10645 | 0.4002 | 0.3018 | 0.3006 | 0.1972 |
| 10690 | 0.1914 | 0.0847 | 0.0927 | 0.1487 |
| 11210 | 0.1892 | 0.0948 | 0.1104 | 0.1272 |
| 11490 | 0.2190 | 0.1160 | 0.0599 | 0.0020 |
| 12530 | 0.2701 | 0.2109 | 0.2589 | 0.1546 |
| 12585 | 0.2238 | 0.1639 | 0.0974 | 0.1202 |
| 12770 | 0.2145 | 0.1574 | 0.1738 | 0.1527 |
| 12895 | 0.3307 | 0.1083 | 0.0747 | -0.0929 |
| 12945 | 0.2775 | 0.0790 | 0.0390 | -0.0930 |
| 12960 | 0.3174 | 0.2051 | 0.1578 | 0.0107 |
| 12965 | 0.2288 | 0.1017 | -0.0009 | -0.0182 |
| 13020 | 0.2327 | 0.1016 | 0.0467 | 0.0452 |
| 13050 | 0.2285 | 0.1553 | 0.0751 | -0.0046 |
| 13245 | 0.3465 | 0.2104 | 0.1377 | 0.0856 |
| 13500 | 0.2145 | 0.1285 | 0.0469 | 0.0859 |
| 13695 | 0.2334 | 0.0653 | 0.0760 | 0.1371 |
| 14185 | 0.3771 | 0.2524 | 0.2504 | 0.1490 |
| 14275 | 0.3031 | 0.2569 | 0.2617 | 0.1725 |
| 15105 | 0.2166 | 0.0078 | 0.1056 | 0.0568 |
| 15145 | 0.1613 | 0.0372 | 0.0484 | -0.0455 |
| 15355 | 0.2608 | 0.0644 | 0.0536 | 0.1261 |
| 15395 | 0.2588 | 0.1652 | 0.1397 | -0.0429 |
| 15430 | 0.2713 | 0.1773 | 0.0997 | 0.1382 |
| 15450 | 0.4525 | 0.3259 | 0.2216 | 0.1384 |
| 15455 | 0.1563 | 0.0191 | 0.0267 | -0.0509 |
| 15500 | 0.1028 | 0.0125 | 0.0460 | -0.0121 |
| 15505 | 0.2527 | 0.1576 | 0.0632 | -0.0192 |
| 15535 | 0.1162 | 0.1978 | 0.1776 | 0.0291 |
| 15540 | 0.2146 | 0.1759 | 0.2321 | 0.2190 |
| 15545 | 0.1930 | 0.1097 | 0.0122 | -0.0424 |
| 15565 | 0.1988 | 0.0411 | -0.0041 | 0.0208 |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 5 (continued)
Item Discrimination Values for Examination 1 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 15570 | 0.2229 | 0.1718 | 0.1751 | 0.0885 |
| 15585 | 0.3128 | 0.2454 | 0.1973 | 0.2066 |
| 15600 | 0.3630 | 0.1857 | 0.0534 | -0.0131 |
| 15615 | 0.2029 | 0.0816 | 0.2168 | 0.1541 |
| 15630 | 0.2938 | 0.2319 | 0.1811 | 0.2114 |
| 15635 | 0.3756 | 0.2302 | -0.0395 | 0.0333 |
| 15645 | 0.2468 | -0.0099 | 0.0773 | -0.0293 |
| 15660 | 0.2246 | 0.0640 | 0.1279 | 0.0858 |
| 15665 | 0.1791 | 0.0625 | 0.1587 | -0.0291 |
| 15670 | 0.1945 | 0.1537 | 0.0233 | -0.0080 |
| 15680 | 0.1612 | 0.0802 | 0.0750 | 0.0606 |
| 15690 | 0.1971 | 0.1508 | 0.0505 | -0.0148 |
| 15705 | 0.3067 | 0.1803 | 0.0389 | -0.0058 |
| 15710 | 0.1927 | 0.0747 | -0.0572 | -0.1244 |
| 15715 | 0.1508 | -0.0444 | 0.1394 | -0.0770 |
| 15720 | 0.2855 | 0.1780 | 0.0129 | 0.0175 |
| 15725 | 0.2224 | 0.0661 | 0.0823 | 0.0384 |
| 15745 | 0.2033 | 0.0376 | 0.0672 | 0.1272 |
| 15755 | 0.2927 | 0.1555 | 0.1453 | 0.0623 |
| 15770 | 0.1920 | 0.0887 | 0.1407 | -0.0016 |
| 15790 | 0.1749 | 0.1648 | 0.0089 | 0.1319 |
| 15795 | 0.0801 | 0.0082 | 0.1324 | 0.1549 |
| 15800 | 0.1837 | 0.1092 | 0.1995 | 0.3237 |
| 15810 | 0.3582 | 0.1756 | 0.2419 | 0.3322 |
| 15890 | 0.3484 | 0.2303 | 0.1115 | 0.0534 |
| 15905 | 0.4017 | 0.1764 | 0.0123 | 0.1021 |
| 15930 | 0.1898 | 0.0497 | 0.1781 | 0.1658 |
| 16005 | 0.3986 | 0.2210 | 0.1513 | 0.1900 |
| 16060 | 0.2868 | 0.1568 | 0.1397 | 0.1144 |
| 16110 | 0.4243 | 0.2521 | 0.2559 |  |
| 16135 | 0.2716 | 0.2047 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 5 (continued)
Item Discrimination Values for Examination $1-C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16210 | 0.4188 | 0.2665 | 0.2382 | 0.2047 |
| 16255 | 0.1220 | 0.0334 | 0.0502 | 0.0686 |
| 16275 | 0.1715 | 0.1114 | 0.1274 | 0.2231 |
| 16280 | 0.2719 | 0.1954 | -0.0017 | -0.0005 |
| 16305 | 0.1803 | 0.0598 | 0.1123 | 0.1211 |
| 16375 | 0.3525 | 0.1963 | -0.0340 | -0.0618 |
| 16390 | 0.2068 | 0.0928 | 0.0576 | 0.0015 |
| 16415 | 0.2210 | 0.1390 | 0.0572 | 0.0461 |
| 16425 | 0.2634 | 0.1116 | -0.0110 | -0.0653 |
| 16435 | 0.1672 | 0.0178 | 0.0176 | 0.1032 |
| 16445 | 0.2965 | 0.1245 | 0.0810 | 0.0579 |
| 16470 | 0.2230 | 0.0618 | 0.1847 | 0.1152 |
| 16475 | 0.2791 | 0.2140 | -0.0365 | -0.1760 |
| 16505 | 0.1442 | 0.0462 | -0.0570 | -0.1010 |
| 16515 | 0.1771 | 0.0099 | -0.0165 | 0.0163 |
| 16525 | 0.0557 | -0.0421 | 0.1399 | 0.1473 |
| 16560 | 0.2017 | 0.1201 | 0.0658 | 0.0329 |
| 16605 | 0.0227 | 0.1076 | 0.0420 | -0.0575 |
| 16615 | 0.2484 | 0.1331 | 0.0532 | 0.2335 |
| 16635 | 0.2581 | 0.0779 | 0.2417 | 0.1423 |
| 16690 | 0.3027 | 0.2056 | 0.0423 |  |
| 16710 | 0.4226 | 0.2604 | 0.0341 |  |
| 16715 | 0.2471 | 0.1103 | 0.0889 | 0.0469 |
| 16720 | 0.4695 | 0.3140 | 0.2672 | 0.0522 |
| 16725 | 0.3359 | 0.1690 | 0.0549 | 0.0885 |
| 16740 | 0.3407 | 0.2339 | 0.1416 | 0.0156 |
| 16745 | 0.2644 | 0.1573 | 0.1125 | 0.0950 |
| 16755 | 0.2363 | 0.1455 | 0.1299 | 0.0681 |
| 16760 | 0.3060 | 0.2325 | 0.2058 | 0.1175 |
| 16765 | 0.3531 | 0.2844 | 0.1999 | 0.0113 |
| 16775 | 0.2067 | 0.1370 | 0.0884 |  |
|  |  |  |  |  |
|  |  |  |  | 0 |
|  |  |  |  |  |

Table A. 5 (continued)
Item Discrimination Values for Examination 1 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 16800 | 0.2626 | 0.0802 | 0.0606 | 0.1544 |
| 16840 | 0.3550 | 0.2071 | 0.1176 | 0.0611 |
| 16845 | 0.4353 | 0.2977 | 0.1702 | 0.0713 |
| 16895 | 0.1322 | 0.0347 | -0.0619 | -0.0933 |
| 16910 | 0.3609 | 0.2187 | 0.2064 | 0.0382 |
| 16925 | 0.1134 | 0.0449 | 0.0748 | -0.0591 |
| 16940 | 0.2813 | 0.1512 | 0.1427 | 0.0582 |
| 16945 | 0.2289 | 0.0978 | 0.1447 | 0.0745 |
| 16955 | 0.1375 | 0.0253 | 0.0119 | 0.0192 |
| 17010 | 0.2586 | 0.1302 | 0.1414 | 0.1248 |
| 17020 | 0.2144 | 0.0608 | 0.0562 | -0.0242 |
| 17050 | 0.1092 | 0.0501 | 0.0510 | -0.0246 |
| 17105 | 0.2338 | 0.1698 | 0.1251 | 0.1069 |
| 17115 | 0.2431 | 0.1839 | 0.1494 | 0.1351 |
| 17125 | 0.1825 | 0.0658 | 0.0233 | 0.0365 |
| 17130 | 0.1915 | 0.1276 | 0.1095 | 0.0660 |
| 17160 | 0.0628 | 0.1016 | 0.1340 | 0.1322 |
| 17240 | 0.3277 | 0.2671 | 0.2135 | 0.1926 |
| 17245 | 0.2662 | 0.1124 | -0.0107 | -0.1466 |
| 17265 | 0.2291 | 0.1063 | 0.0837 | 0.1190 |
| 17275 | -0.0524 | -0.0694 | -0.0494 | 0.0851 |
| 17285 | 0.2758 | 0.0835 | 0.1730 | -0.0374 |
| 17300 | 0.1741 | 0.1673 | 0.1295 | 0.0307 |
| 17335 | 0.2070 | 0.1304 | 0.1133 | 0.1119 |
| 17340 | 0.1773 | 0.1290 | 0.1003 | 0.1004 |
| 17345 | 0.2649 | 0.0939 | 0.0689 | -0.0474 |
| 17370 | 0.3665 | 0.1695 | 0.0389 | 0.1256 |
| 17380 | 0.1734 | 0.0917 | 0.1203 | 0.1460 |
| 17385 | 0.2695 | 0.2140 | 0.2053 | 0.1199 |
| 17425 | 0.2585 | 0.1435 | 0.1396 | 0.2140 |
| 17430 | 0.3883 | 0.2379 | 0.2515 | 0.0908 |
|  |  |  |  |  |
|  |  |  |  |  |

Table A. 5 (continued)
Item Discrimination Values for Examination 1 - $C_{X 5}$

| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  |  |  |
| 17460 | 0.2122 | 0.0983 | 0.0861 | 0.1096 |
| 17480 | 0.2261 | 0.1772 | 0.0491 | 0.0521 |
| 17495 | 0.2081 | 0.0856 | 0.0704 | 0.1124 |
| 17500 | 0.3347 | 0.2137 | 0.1789 | 0.1560 |
| 17510 | 0.1919 | 0.0519 | 0.0465 | -0.0357 |
| 17520 | 0.1498 | 0.0039 | -0.0028 | -0.0779 |
| 17545 | 0.1543 | 0.0382 | 0.0315 | 0.1497 |
| 17550 | 0.0717 | -0.0578 | -0.1204 | -0.0823 |
| 17565 | 0.1601 | 0.1594 | 0.0552 | -0.1419 |
| 17580 | 0.3078 | 0.1659 | 0.2171 | 0.0338 |
| 17585 | 0.2006 | 0.1467 | 0.2387 | 0.1421 |
| 17595 | 0.2681 | 0.1879 | 0.2464 | 0.2786 |
| 17605 | 0.1571 | 0.0392 | 0.0392 | -0.0645 |
| 17635 | 0.1269 | 0.0419 | -0.0191 | 0.0643 |
| 17640 | 0.2316 | 0.1347 | 0.1030 | 0.1278 |
| 17660 | 0.1847 | 0.0337 | 0.0200 | -0.0305 |
| 17665 | 0.2484 | 0.2262 | 0.2389 | 0.0733 |
| 17670 | 0.2394 | 0.0600 | 0.0414 | -0.1113 |
| 17690 | 0.1535 | 0.0912 | 0.1065 | 0.0696 |
| 17695 | 0.1150 | 0.0550 | 0.0756 | 0.0639 |
|  |  |  |  |  |

## APPENDIX B: ITEM DISCRIMINATION VALUES - EXAMINATION 2

Table B. 1

Item Discrimination Values for Examination $2-C_{X I}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 32497 | 0.2411 | 0.1088 | 0.1045 | 0.3119 |
| 40655 | 0.3516 | 0.2677 | 0.2728 | 0.1453 |
| 40657 | 0.3651 | 0.1948 | -0.1051 | -0.0620 |
| 40668 | 0.3077 | 0.1417 | 0.1592 | -0.0461 |
| 40885 | 0.4001 | 0.2015 | -0.0435 | -0.0805 |
| 40903 | 0.1012 | 0.0204 | 0.1016 | 0.0440 |
| 40907 | 0.0642 | 0.1555 | 0.1137 | 0.0939 |
| 40916 | 0.0227 | 0.1408 | 0.0667 | 0.0485 |
| 40944 | 0.2750 | 0.3059 | 0.2568 | 0.1897 |
| 40969 | 0.3391 | 0.3193 | 0.3707 | 0.3796 |
| 40992 | 0.3734 | 0.2826 | 0.1050 | 0.0509 |
| 41054 | 0.0845 | -0.2002 | -0.1794 | -0.1965 |
| 41059 | 0.1457 | 0.1697 | 0.0926 | 0.1185 |
| 41066 | 0.0993 | 0.0626 | 0.1019 | -0.1139 |
| 41093 | 0.2935 | 0.1386 | -0.0086 | 0.0120 |
| 41098 | 0.1669 | 0.0264 | 0.1606 | 0.1545 |
| 41101 | 0.1742 | 0.1560 | 0.1077 | 0.0704 |
| 41181 | 0.3619 | 0.2595 | 0.1404 | 0.0876 |
| 41190 | 0.2607 | 0.1098 | 0.0807 | 0.0434 |
| 41337 | 0.1841 | 0.0721 | -0.0447 | 0.1130 |
| 41407 | 0.2033 | 0.0772 | 0.1626 | 0.0524 |
| 41632 | 0.2642 | 0.1538 | 0.1429 | 0.2577 |
| 41645 | 0.2524 | 0.2612 | 0.0294 | 0.0234 |
| 41653 | -0.0460 | -0.0435 | -0.0026 | 0.1456 |
| 41669 | -0.0728 | 0.0051 | 0.0408 | -0.0599 |
| 41690 | 0.1151 | 0.0477 | -0.0077 | 0.0809 |
| 42503 | 0.2553 | 0.1759 | 0.0161 | 0.0337 |
| 42504 | 0.1333 | -0.0408 | -0.0308 | -0.0690 |
| 51693 | 0.1997 | 0.1550 | -0.0197 | -0.1513 |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 51717 | 0.1998 | 0.1537 | 0.2050 | 0.1044 |
| 51731 | 0.2534 | 0.0747 | 0.0125 | 0.0761 |
| 51732 | 0.1875 | 0.0388 | 0.0258 | 0.0063 |
| 51742 | 0.2057 | -0.0932 | 0.0101 | 0.0495 |
| 51781 | 0.2811 | 0.2617 | 0.0883 | 0.1210 |
| 51786 | 0.1640 | -0.0188 | -0.1100 | -0.0919 |
| 51787 | 0.3500 | 0.2684 | 0.1367 | 0.0703 |
| 51793 | 0.3382 | 0.1907 | 0.0770 | 0.0337 |
| 51806 | 0.3013 | 0.0783 | 0.1679 | 0.0776 |
| 51817 | 0.3639 | 0.1137 | 0.2227 | 0.1256 |
| 51825 | 0.2683 | 0.1417 | 0.0685 | -0.0656 |
| 51839 | 0.0601 | -0.1208 | -0.0579 | 0.0741 |
| 51843 | -0.0361 | 0.0095 | -0.0785 | -0.0844 |
| 59127 | 0.2984 | 0.0516 | 0.0607 | 0.0363 |
| 59128 | 0.2980 | 0.2535 | 0.1781 | 0.0248 |
| 59137 | 0.0786 | 0.1356 | 0.1849 | 0.0948 |
| 59138 | 0.1432 | 0.2705 | 0.1524 | 0.2133 |
| 59149 | 0.3560 | 0.1372 | 0.1294 | -0.2314 |
| 59155 | 0.2243 | 0.1424 | 0.0573 | -0.0345 |
| 59184 | 0.3020 | 0.1213 | 0.0784 | 0.0472 |
| 59185 | 0.0802 | -0.0351 | 0.1374 | 0.1896 |
| 59217 | 0.1105 | 0.0538 | 0.0347 | 0.0000 |
| 59231 | 0.3649 | 0.1660 | 0.1124 | 0.1784 |
| 59239 | 0.2305 | 0.1372 | 0.2357 | 0.1786 |
| 59267 | 0.0877 | 0.1473 | 0.1972 | 0.2640 |
| 59489 | 0.2272 | 0.0358 | 0.0383 | -0.0484 |
| 60230 | 0.2109 | 0.0000 | 0.0000 | 0.0000 |
| 60240 | -0.0221 | 0.0692 | 0.1464 | 0.0743 |
| 60241 | 0.2474 | -0.0835 | -0.0642 | -0.1499 |
| 60242 | 0.2856 | 0.3245 | 0.3093 | 0.0980 |
| 60250 | 0.1273 | -0.0456 | -0.0978 | 0.0058 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | ---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 60258 | 0.3100 | 0.3535 | 0.2324 | -0.1019 |
| 62125 | 0.1018 | -0.1046 | 0.1886 | -0.1755 |
| 62129 | 0.2786 | 0.1223 | 0.2131 | 0.0713 |
| 62132 | 0.2708 | 0.2167 | 0.2044 | 0.1020 |
| 62135 | 0.2983 | 0.2064 | 0.1068 | 0.1849 |
| 62136 | 0.3512 | 0.2545 | 0.3059 | 0.3243 |
| 62140 | 0.3812 | 0.3175 | 0.1779 | -0.0328 |
| 84442 | 0.3069 | 0.3427 | 0.0799 | 0.0899 |
| 90381 | 0.2819 | 0.1300 | -0.0459 | -0.0293 |
| 90382 | 0.0569 | -0.0875 | 0.1892 | 0.0765 |
| 90384 | 0.1036 | 0.0280 | 0.1522 | 0.0637 |
| 90385 | 0.1792 | 0.0809 | 0.1265 | -0.0874 |
| 90386 | 0.2554 | 0.1521 | 0.1648 | 0.2544 |
| 90390 | 0.4318 | 0.2788 | -0.0912 | 0.0178 |
| 90396 | 0.0718 | -0.0149 | -0.0399 | -0.0980 |
| 90397 | 0.1268 | 0.0301 | 0.0011 | 0.1049 |
| 90398 | 0.2795 | 0.2136 | 0.0123 | 0.1101 |
| 90400 | 0.1985 | 0.2581 | -0.1390 | -0.2133 |
| 94445 | 0.2217 | 0.0226 | 0.1340 | 0.0115 |
| 94505 | 0.0600 | 0.0874 | 0.1603 | 0.1091 |
| 94510 | 0.1683 | 0.1346 | 0.1626 | -0.0174 |
| 94566 | 0.3157 | 0.2505 | 0.1438 |  |
| 94582 | 0.3808 | 0.2083 | 0.1338 | 0.1837 |
| 94586 | 0.3209 | 0.1196 | 0.2029 | -0.0696 |
| 108131 | 0.1744 | 0.0693 | -0.0800 | 0.0891 |
| 108132 | 0.2818 | 0.1501 | -0.0250 | 0.2447 |
| 108133 | 0.1783 | 0.1505 | 0.0895 | 0.1733 |
| 108134 | 0.2241 | 0.1205 | -0.0060 | -0.0590 |
| 108135 | 0.0995 | 0.0123 | 0.0845 | 0.1717 |
| 108136 | 0.2619 | 0.1431 | 0.1045 | 0.0000 |
| 108137 | 0.1924 | 0.1771 | 0.0000 |  |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108138 | 0.3920 | 0.2342 | 0.2480 | 0.0839 |
| 108139 | 0.2255 | 0.0218 | -0.1041 | -0.1280 |
| 108140 | 0.1381 | 0.0705 | 0.0461 | 0.1769 |
| 108141 | 0.0964 | 0.0340 | 0.0788 | -0.0748 |
| 108142 | 0.2794 | 0.1865 | 0.2009 | 0.3159 |
| 108143 | 0.4001 | 0.2367 | 0.3507 | 0.3147 |
| 108144 | 0.2980 | 0.2610 | 0.1608 | 0.0062 |
| 108145 | 0.1484 | 0.0735 | 0.1589 | 0.0191 |
| 108146 | 0.3033 | 0.1912 | 0.1364 | 0.2508 |
| 108147 | 0.1831 | 0.1061 | 0.0302 | 0.1849 |
| 108148 | 0.1280 | -0.0378 | -0.0804 | -0.0306 |
| 108149 | -0.0003 | -0.0045 | 0.0284 | 0.0653 |
| 108150 | 0.1513 | 0.0746 | -0.0432 | 0.0410 |
| 108151 | 0.1869 | 0.1356 | 0.0359 | -0.0693 |
| 108152 | 0.3092 | 0.0667 | 0.0031 | -0.0805 |
| 108153 | 0.4983 | 0.3633 | 0.3799 | 0.4009 |
| 108154 | 0.3769 | 0.2998 | 0.1842 | 0.3276 |
| 108155 | 0.3370 | 0.0562 | -0.0211 | -0.0402 |
| 108156 | 0.3039 | 0.1754 | 0.1458 | 0.0122 |
| 108157 | 0.2223 | 0.0730 | 0.1115 | 0.0094 |
| 108158 | 0.1994 | 0.2088 | 0.1896 | 0.1006 |
| 108159 | 0.0479 | -0.0548 | -0.1484 | -0.2965 |
| 108160 | -0.0783 | -0.0242 | 0.0621 | 0.0586 |
| 108161 | -0.0186 | 0.1302 | 0.2062 | 0.2870 |
| 108162 | 0.3128 | 0.2080 | 0.2559 | 0.2191 |
| 108163 | 0.2605 | 0.1741 | 0.2342 | 0.1667 |
| 108164 | 0.2834 | 0.0740 | 0.1866 | 0.1513 |
| 108165 | -0.0191 | -0.0032 | -0.0881 | -0.2013 |
| 108166 | 0.3038 | 0.1698 | 0.1245 | 0.0899 |
| 108167 | 0.3061 | 0.1468 | 0.2038 | 0.0708 |
| 108168 | 0.2278 | 0.0633 | 0.0661 | 0.0556 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108169 | -0.0768 | -0.1679 | -0.2659 | -0.2630 |
| 108170 | 0.2778 | 0.1033 | 0.0900 | -0.0708 |
| 108171 | 0.1082 | 0.1786 | 0.0758 | -0.0466 |
| 108172 | 0.0626 | -0.0366 | -0.1741 | -0.1177 |
| 108173 | 0.2470 | 0.1272 | 0.0804 | 0.1750 |
| 108174 | 0.1734 | 0.1075 | 0.0961 | 0.0436 |
| 108175 | 0.2533 | 0.1593 | 0.1923 | 0.1063 |
| 108176 | 0.3077 | 0.2271 | 0.0810 | -0.0843 |
| 108177 | 0.3764 | 0.1448 | 0.0299 | -0.1053 |
| 108178 | 0.4233 | 0.2727 | 0.2012 | 0.3035 |
| 108179 | 0.0484 | 0.0543 | -0.0744 | -0.0849 |
| 108180 | 0.3023 | 0.2710 | 0.1600 | 0.0839 |
| 108181 | 0.2335 | 0.0185 | -0.0181 | 0.0219 |
| 108182 | 0.3435 | 0.1362 | -0.0165 | 0.1017 |
| 108183 | 0.2177 | 0.0914 | 0.0733 | 0.0413 |
| 108184 | 0.2636 | 0.0996 | 0.1018 | 0.2895 |
| 108185 | -0.2086 | -0.1260 | -0.1461 | -0.0351 |
| 108186 | -0.1053 | 0.0000 | 0.0000 | 0.0000 |
| 108187 | 0.2966 | 0.2967 | 0.1270 | -0.0990 |
| 108188 | 0.2519 | -0.0497 | -0.1329 | -0.1258 |
| 108189 | 0.3518 | 0.2523 | 0.3109 | 0.3695 |
| 108190 | 0.2102 | 0.1429 | 0.2029 | 0.0749 |
| 108191 | 0.2098 | 0.0961 | 0.1530 | 0.2026 |
| 108192 | 0.2053 | -0.0515 | -0.0028 | 0.0564 |
| 108193 | 0.2459 | 0.2312 | 0.2570 | 0.3014 |
| 108194 | 0.1967 | 0.1117 | 0.2363 | 0.1381 |
| 108195 | 0.1906 | 0.1138 | 0.0027 | -0.0350 |
| 108196 | 0.2824 | 0.0755 | 0.0942 | 0.1752 |
| 108197 | 0.2556 | 0.2702 | 0.3184 | 0.2874 |
| 108198 | 0.2492 | 0.2751 | 0.0787 | 0.0761 |
| 108199 | 0.2099 | 0.0687 | -0.0039 | -0.0428 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108200 | 0.1870 | 0.1260 | -0.0356 | -0.0658 |
| 108201 | 0.3032 | 0.1365 | 0.1140 | -0.1432 |
| 108202 | -0.0915 | -0.1807 | -0.1019 | 0.0059 |
| 108203 | -0.0080 | -0.0564 | 0.1211 | 0.1383 |
| 108204 | 0.3355 | 0.1726 | 0.1002 | 0.0644 |
| 108477 | 0.3690 | 0.3482 | 0.4438 | 0.2723 |
| 108478 | 0.2129 | 0.3464 | 0.2730 | 0.2179 |
| 108479 | 0.2899 | 0.2657 | 0.1506 | 0.1274 |
| 108480 | 0.1822 | 0.0239 | -0.0250 | -0.1625 |
| 108481 | 0.2634 | 0.2052 | 0.1571 | 0.0942 |
| 108482 | 0.2746 | 0.0001 | 0.2489 | 0.1444 |
| 108483 | 0.3192 | -0.0554 | 0.0399 | -0.0519 |
| 108484 | 0.2636 | 0.1011 | 0.0268 | -0.0127 |
| 108485 | 0.1653 | 0.0840 | 0.1294 | 0.1474 |
| 108486 | 0.4052 | 0.3405 | 0.2698 | 0.2398 |
| 108487 | 0.3372 | 0.2292 | 0.1781 | 0.0484 |
| 108488 | 0.1810 | 0.2683 | 0.2115 | 0.1441 |
| 108489 | 0.3494 | 0.0874 | -0.0482 | 0.0980 |
| 108490 | 0.1444 | 0.0790 | 0.0177 | 0.1201 |
| 108491 | 0.0475 | 0.1464 | -0.0475 | -0.0061 |
| 108492 | 0.4010 | 0.3024 | 0.2600 | 0.2214 |
| 108493 | 0.1714 | 0.0297 | 0.0171 | -0.1620 |
| 108494 | 0.3177 | 0.2564 | 0.1706 | 0.0551 |
| 108495 | 0.1151 | -0.0390 | -0.0312 | -0.1451 |
| 108497 | 0.2921 | 0.1930 | 0.0910 | 0.1558 |
| 108498 | 0.2975 | 0.0949 | 0.0201 | 0.1049 |
| 108499 | 0.2863 | 0.1820 | 0.0229 | -0.0977 |
| 108500 | 0.1575 | 0.0487 | 0.1197 | 0.0636 |
| 108501 | 0.0743 | 0.1239 | -0.0333 | 0.0121 |
| 108502 | 0.1566 | 0.2135 | 0.2387 | 0.2550 |
| 108503 | 0.1387 | 0.0435 | 0.1541 | 0.1523 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 1 (continued)
Item Discrimination Values for Examination $2-C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108504 | 0.1001 | 0.1834 | 0.1075 | 0.1422 |
| 108505 | 0.2677 | 0.1163 | 0.0931 | 0.0825 |
| 108506 | 0.3136 | 0.1253 | 0.1246 | 0.3025 |
| 108507 | 0.3458 | 0.1827 | 0.1238 | 0.0908 |
| 108508 | 0.2697 | 0.1583 | 0.0758 | 0.1447 |
| 108509 | 0.3294 | 0.1003 | 0.1494 | -0.0464 |
| 108510 | 0.1560 | -0.0513 | 0.0152 | 0.0121 |
| 108511 | 0.0278 | -0.1043 | -0.0483 | 0.0000 |
| 108512 | 0.2234 | 0.2715 | 0.0935 | -0.0300 |
| 108513 | 0.1451 | 0.0620 | 0.0312 | 0.0000 |
| 108514 | 0.2470 | 0.0962 | 0.0856 | 0.0532 |
| 108515 | 0.3178 | 0.2680 | 0.2600 | 0.1618 |
| 108523 | 0.1711 | -0.0361 | 0.1494 | 0.0194 |
| 108524 | 0.1976 | 0.0928 | -0.0181 | 0.0328 |
| 108525 | 0.0897 | -0.0013 | -0.0400 | -0.1213 |
| 950671494 | 0.1338 | 0.0737 | 0.0980 | 0.1684 |
|  |  |  |  |  |

Table B. 2
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.0356 | -0.0685 |
| 32497 | 0.2411 | -0.0409 | 0.2038 | 0.2693 |
| 40655 | 0.3516 | 0.2498 | 0.1451 | -0.0118 |
| 40657 | 0.3651 | 0.1337 | 0.1325 | 0.2305 |
| 40668 | 0.3077 | 0.2367 | 0.2103 | 0.1320 |
| 40885 | 0.4001 | 0.2311 | 0.0340 | 0.2328 |
| 40903 | 0.1012 | -0.0477 | 0.1475 | 0.2394 |
| 40907 | 0.0642 | 0.0797 | 0.1755 | 0.0237 |
| 40916 | 0.0227 | 0.1476 | 0.2527 | 0.1175 |
| 40944 | 0.2750 | 0.3149 | 0.2221 | 0.1288 |
| 40969 | 0.3391 | 0.2484 | 0.2357 | 0.1442 |
| 40992 | 0.3734 | 0.2332 | 0.0627 | 0.0166 |
| 41054 | 0.0845 | 0.0599 | 0.1369 | 0.0000 |
| 41059 | 0.1457 | 0.1110 | -0.0411 | 0.2237 |
| 41066 | 0.0993 | -0.0005 | 0.1270 | -0.0072 |
| 41093 | 0.2935 | 0.0780 | 0.1325 | 0.3057 |
| 41098 | 0.1669 | 0.1128 | 0.1219 | 0.2197 |
| 41101 | 0.1742 | 0.0943 | 0.1940 | 0.1671 |
| 41181 | 0.3619 | 0.1092 | 0.0972 | 0.1656 |
| 41190 | 0.2607 | 0.1729 | 0.1123 | -0.1718 |
| 41337 | 0.1841 | 0.1587 | -0.0203 | 0.1618 |
| 41407 | 0.2033 | 0.1377 | 0.2057 | 0.0311 |
| 41632 | 0.2642 | 0.1378 | 0.2203 | 0.1220 |
| 41645 | 0.2524 | 0.1910 | -0.0151 | -0.1590 |
| 41653 | -0.0460 | -0.0469 | 0.0186 | 0.0985 |
| 41669 | -0.0728 | 0.0416 | 0.0636 | -0.0829 |
| 41690 | 0.1151 | 0.0899 | 0.1459 | 0.0578 |
| 42503 | 0.2553 | 0.1381 | 0.0296 | 0.2842 |
| 42504 | 0.1333 | 0.0049 | 0.0556 |  |
| 51693 | 0.1997 | 0.1037 | 0.0203 | 0.0625 |
| 51717 | 0.1998 | 0.1898 |  | 0.1393 |
| 51731 | 0.2534 |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 51732 | 0.1875 | 0.1015 | 0.1013 | -0.0023 |
| 51742 | 0.2057 | 0.1257 | 0.0608 | 0.1165 |
| 51781 | 0.2811 | 0.2147 | 0.2574 | 0.0447 |
| 51786 | 0.1640 | -0.0322 | -0.0112 | -0.1302 |
| 51787 | 0.3500 | 0.3092 | 0.3465 | 0.3056 |
| 51793 | 0.3382 | 0.2101 | 0.1935 | 0.2670 |
| 51806 | 0.3013 | 0.1144 | 0.0106 | 0.0204 |
| 51817 | 0.3639 | 0.1467 | 0.0585 | 0.1444 |
| 51825 | 0.2683 | 0.2142 | 0.1666 | 0.1742 |
| 51839 | 0.0601 | -0.0765 | -0.1698 | -0.2116 |
| 51843 | -0.0361 | -0.0606 | -0.0627 | 0.0258 |
| 59127 | 0.2984 | 0.0633 | 0.0295 | -0.1260 |
| 59128 | 0.2980 | 0.1150 | 0.1521 | 0.2419 |
| 59137 | 0.0786 | 0.1915 | 0.1243 | 0.0000 |
| 59138 | 0.1432 | 0.1491 | 0.0000 | 0.0000 |
| 59149 | 0.3560 | 0.1003 | 0.1248 | 0.3233 |
| 59155 | 0.2243 | 0.0655 | 0.0000 | 0.1517 |
| 59184 | 0.3020 | 0.1718 | 0.1475 | 0.0595 |
| 59185 | 0.0802 | 0.0221 | -0.1538 | 0.0000 |
| 59217 | 0.1105 | 0.0763 | 0.0395 | -0.0459 |
| 59231 | 0.3649 | 0.2244 | 0.2134 | 0.0129 |
| 59239 | 0.2305 | 0.2019 | 0.2079 | 0.2386 |
| 59267 | 0.0877 | 0.2161 | 0.2692 | 0.0000 |
| 59489 | 0.2272 | 0.1459 | 0.0969 | 0.1042 |
| 60230 | 0.2109 | 0.0000 | 0.0000 | 0.0000 |
| 60240 | -0.0221 | 0.1232 | 0.2100 | 0.3068 |
| 60241 | 0.2474 | -0.0498 | -0.0746 | -0.0469 |
| 60242 | 0.2856 | 0.2881 | 0.2412 | 0.3010 |
| 60250 | 0.1273 | 0.0464 | 0.0975 | -0.1050 |
| 60258 | 0.3100 | 0.1869 | 0.2612 | 0.1235 |
| 62125 | 0.1018 | -0.1275 | -0.0901 | -0.0296 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 62129 | 0.2786 | 0.0722 | 0.0919 | 0.0699 |
| 62132 | 0.2708 | 0.2335 | 0.2262 | 0.0934 |
| 62135 | 0.2983 | 0.1945 | 0.2147 | 0.2990 |
| 62136 | 0.3512 | 0.2552 | 0.2120 | -0.0235 |
| 62140 | 0.3812 | 0.2072 | 0.2383 | 0.1184 |
| 84442 | 0.3069 | 0.1904 | 0.1511 | 0.0700 |
| 90381 | 0.2819 | 0.2417 | 0.2833 | 0.1656 |
| 90382 | 0.0569 | -0.0270 | 0.0174 | 0.0790 |
| 90384 | 0.1036 | -0.0126 | 0.0568 | 0.2153 |
| 90385 | 0.1792 | 0.1035 | -0.0138 | -0.0617 |
| 90386 | 0.2554 | 0.0319 | 0.0816 | 0.1070 |
| 90390 | 0.4318 | 0.2446 | 0.1485 | -0.0311 |
| 90396 | 0.0718 | -0.0831 | -0.2507 | -0.3411 |
| 90397 | 0.1268 | 0.0316 | -0.0064 | 0.1305 |
| 90398 | 0.2795 | 0.1817 | 0.1018 | 0.0740 |
| 90400 | 0.1985 | 0.1197 | 0.1070 | 0.0861 |
| 94445 | 0.2217 | 0.0171 | -0.1217 | -0.2287 |
| 94505 | 0.0600 | 0.0143 | 0.0634 | -0.0759 |
| 94510 | 0.1683 | 0.1864 | 0.1453 | 0.0429 |
| 94566 | 0.3157 | 0.1226 | 0.0945 | 0.2292 |
| 94582 | 0.3808 | 0.1181 | 0.2256 | 0.1695 |
| 94586 | 0.3209 | 0.1981 | 0.2185 | 0.0121 |
| 108131 | 0.1744 | 0.0105 | 0.0207 | 0.0469 |
| 108132 | 0.2818 | 0.2602 | 0.2164 | -0.0865 |
| 108133 | 0.1783 | -0.0488 | 0.0032 | -0.1045 |
| 108134 | 0.2241 | 0.1481 | 0.1232 | 0.0297 |
| 108135 | 0.0995 | 0.1083 | 0.1095 | 0.0279 |
| 108136 | 0.2619 | 0.1307 | 0.1885 | 0.1314 |
| 108137 | 0.1924 | 0.0000 | 0.0000 | 0.0000 |
| 108138 | 0.3920 | 0.3379 | 0.2430 | 0.2353 |
| 108139 | 0.2255 | -0.1072 | -0.0631 | -0.0587 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108140 | 0.1381 | 0.1324 | 0.0841 | -0.0398 |
| 108141 | 0.0964 | 0.0539 | 0.0641 | -0.0041 |
| 108142 | 0.2794 | 0.3282 | 0.3306 | 0.3337 |
| 108143 | 0.4001 | 0.2228 | 0.1770 | -0.0022 |
| 108144 | 0.2980 | 0.2069 | 0.1486 | -0.1215 |
| 108145 | 0.1484 | 0.0818 | 0.0727 | 0.0487 |
| 108146 | 0.3033 | 0.1245 | -0.0142 | -0.1655 |
| 108147 | 0.1831 | 0.1334 | 0.0434 | -0.0895 |
| 108148 | 0.1280 | -0.0283 | -0.0687 | -0.1046 |
| 108149 | -0.0003 | 0.0193 | 0.0535 | 0.1319 |
| 108150 | 0.1513 | 0.1207 | 0.1502 | 0.2018 |
| 108151 | 0.1869 | 0.0852 | 0.0205 | 0.0903 |
| 108152 | 0.3092 | 0.0244 | -0.0329 | -0.0596 |
| 108153 | 0.4983 | 0.4318 | 0.3205 | 0.1534 |
| 108154 | 0.3769 | 0.3061 | 0.1561 | 0.0663 |
| 108155 | 0.3370 | 0.2068 | 0.0810 | -0.0083 |
| 108156 | 0.3039 | 0.1962 | 0.1195 | 0.1964 |
| 108157 | 0.2223 | 0.1026 | -0.0280 | -0.0803 |
| 108158 | 0.1994 | 0.1983 | 0.1742 | 0.0972 |
| 108159 | 0.0479 | -0.0009 | -0.1028 | -0.0762 |
| 108160 | -0.0783 | -0.0487 | 0.0139 | -0.2239 |
| 108161 | -0.0186 | 0.0268 | 0.0329 | 0.0000 |
| 108162 | 0.3128 | 0.2197 | 0.1607 | -0.2268 |
| 108163 | 0.2605 | 0.1773 | 0.2176 | 0.1323 |
| 108164 | 0.2834 | 0.1937 | 0.1661 | 0.1679 |
| 108165 | -0.0191 | -0.0882 | -0.1260 | -0.0140 |
| 108166 | 0.3038 | 0.1974 | 0.1397 | 0.1454 |
| 108167 | 0.3061 | 0.2556 | 0.2492 | 0.1423 |
| 108168 | 0.2278 | 0.1328 | 0.1586 | 0.1722 |
| 108169 | -0.0768 | -0.1223 | -0.0718 | -0.0618 |
| 108170 | 0.2778 | 0.0695 | 0.1498 | 0.1763 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108171 | 0.1082 | 0.0818 | 0.0882 | 0.0989 |
| 108172 | 0.0626 | -0.0499 | -0.1440 | -0.2073 |
| 108173 | 0.2470 | 0.1050 | 0.0396 | -0.1281 |
| 108174 | 0.1734 | 0.2191 | 0.2533 | 0.3011 |
| 108175 | 0.2533 | 0.2265 | 0.1590 | 0.1882 |
| 108176 | 0.3077 | 0.0356 | 0.0660 | -0.0282 |
| 108177 | 0.3764 | 0.2775 | 0.1787 | 0.1763 |
| 108178 | 0.4233 | 0.2678 | 0.2430 | 0.1639 |
| 108179 | 0.0484 | 0.0942 | 0.0159 | -0.0792 |
| 108180 | 0.3023 | 0.1829 | 0.1862 | 0.2005 |
| 108181 | 0.2335 | 0.1187 | 0.1325 | 0.0459 |
| 108182 | 0.3435 | 0.1400 | 0.0570 | 0.0145 |
| 108183 | 0.2177 | 0.0602 | 0.1471 | 0.1260 |
| 108184 | 0.2636 | 0.1966 | 0.0864 | -0.0542 |
| 108185 | -0.2086 | -0.1775 | -0.1737 | -0.1801 |
| 108186 | -0.1053 | 0.0000 | 0.0000 | 0.0000 |
| 108187 | 0.2966 | 0.0598 | 0.0906 | 0.1340 |
| 108188 | 0.2519 | 0.0627 | -0.0230 | -0.0222 |
| 108189 | 0.3518 | 0.2250 | 0.2959 | 0.0555 |
| 108190 | 0.2102 | 0.1048 | 0.0543 | 0.0845 |
| 108191 | 0.2098 | 0.0772 | 0.0848 | 0.0000 |
| 108192 | 0.2053 | 0.1652 | 0.0681 | 0.1117 |
| 108193 | 0.2459 | 0.2224 | 0.2608 | 0.0953 |
| 108194 | 0.1967 | 0.1120 | 0.0186 | 0.0079 |
| 108195 | 0.1906 | 0.1768 | 0.0788 | -0.0568 |
| 108196 | 0.2824 | 0.0525 | -0.0686 | -0.0690 |
| 108197 | 0.2556 | 0.2288 | 0.2365 | 0.1443 |
| 108198 | 0.2492 | 0.2085 | 0.1556 | 0.1362 |
| 108199 | 0.2099 | 0.0962 | -0.0188 | 0.0890 |
| 108200 | 0.1870 | 0.0545 | 0.0392 | -0.0089 |
| 108201 | 0.3032 | -0.0407 | -0.1503 | -0.0539 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.1736 | -0.3521 |
| 108202 | -0.0915 | -0.1957 | -0.0679 | 0.0960 |
| 108203 | -0.0080 | 0.0047 | 0.2585 | 0.1443 |
| 108477 | 0.3355 | 0.1792 | 0.3629 | 0.2620 |
| 108478 | 0.3690 | 0.3591 | 0.3151 | 0.1808 |
| 108479 | 0.2129 | 0.2162 | 0.2497 | 0.2681 |
| 108480 | 0.2899 | 0.1181 | -0.0510 | -0.0046 |
| 108481 | 0.1822 | 0.0954 | 0.1354 | 0.0663 |
| 108482 | 0.2634 | 0.2301 | 0.1303 | 0.1886 |
| 108483 | 0.2746 | 0.2312 | -0.0776 | 0.0314 |
| 108484 | 0.3192 | 0.1264 | 0.1125 | 0.1025 |
| 108485 | 0.2636 | 0.1800 | 0.0455 | 0.0918 |
| 108486 | 0.1653 | 0.0653 | 0.2064 | 0.0487 |
| 108487 | 0.4052 | 0.2466 | 0.1430 | 0.2869 |
| 108488 | 0.3372 | 0.1526 | 0.0996 | 0.1125 |
| 108489 | 0.1810 | 0.1106 | 0.1170 | 0.0214 |
| 108490 | 0.3494 | 0.1799 | 0.1149 | -0.0248 |
| 108491 | 0.1444 | 0.1418 | 0.1788 | -0.1105 |
| 108492 | 0.0475 | 0.0332 | 0.2435 | 0.1062 |
| 108493 | 0.4010 | 0.1720 | -0.0186 | 0.0519 |
| 108494 | 0.1714 | -0.0571 | 0.2521 | 0.0934 |
| 108495 | 0.3177 | 0.1239 | -0.0012 | 0.0754 |
| 108497 | 0.1151 | 0.0337 | 0.1631 | 0.1070 |
| 108498 | 0.2921 | 0.2791 | 0.1986 | 0.0963 |
| 108499 | 0.2975 | 0.2417 | 0.0647 | 0.1240 |
| 108500 | 0.2863 | 0.2217 | 0.1184 | 0.1517 |
| 108501 | 0.1575 | 0.1448 | -0.0156 | 0.0257 |
| 108502 | 0.0743 | 0.0221 | 0.2897 | 0.1833 |
| 108503 | 0.1566 | 0.1955 | 0.0391 | 0.0000 |
| 108504 | 0.1387 | 0.0786 | -0.0092 |  |
| 108505 | 0.1001 | 0.1205 |  |  |
|  | 0.2677 |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 2 (continued)
Item Discrimination Values for Examination $2-C_{X 2}$

| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  | 0.0521 | -0.0286 |
| 108506 | 0.3136 | 0.2320 | 0.2820 | 0.2133 |
| 108507 | 0.3458 | 0.3281 | 0.2170 | 0.0453 |
| 108509 | 0.2697 | 0.1885 | 0.1566 | 0.3253 |
| 108510 | 0.3294 | 0.1470 | -0.1199 | -0.0163 |
| 108511 | 0.1560 | 0.0060 | -0.0327 | -0.0101 |
| 108512 | 0.0278 | 0.1130 | 0.0265 | 0.0830 |
| 108513 | 0.2234 | 0.1716 | 0.0599 | -0.0522 |
| 108514 | 0.1451 | 0.0606 | 0.1871 | 0.1988 |
| 108515 | 0.2470 | 0.1048 | 0.3610 | 0.2085 |
| 108523 | 0.3178 | 0.3673 | -0.0154 | 0.0738 |
| 108524 | 0.1711 | 0.0283 | 0.0732 | 0.0578 |
| 108525 | 0.1976 | 0.0757 | 0.0856 | 0.0429 |
| 950671494 | 0.0897 | 0.0051 | 0.1739 | 0.0523 |
|  | 0.1338 | 0.1032 |  |  |

Table B. 3
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.2911 | 0.2372 |
| 32497 | 0.2411 | 0.1217 | 0.1418 | 0.0946 |
| 40655 | 0.3516 | 0.1787 | 0.1862 | -0.0338 |
| 40657 | 0.3651 | 0.1815 | 0.0199 | -0.0814 |
| 40668 | 0.3077 | 0.0998 | 0.0126 | -0.1389 |
| 40885 | 0.4001 | 0.0595 | 0.0044 | -0.0519 |
| 40903 | 0.1012 | 0.1158 | 0.1101 | 0.0747 |
| 40907 | 0.0642 | 0.0668 | 0.0259 | 0.0198 |
| 40916 | 0.0227 | 0.0482 | 0.2943 | 0.0588 |
| 40944 | 0.2750 | 0.1619 | 0.3650 | 0.3248 |
| 40969 | 0.3391 | 0.4393 | 0.2219 | 0.1033 |
| 40992 | 0.3734 | 0.1857 | -0.4227 | -0.3417 |
| 41054 | 0.0845 | -0.2944 | 0.1566 | 0.0467 |
| 41059 | 0.1457 | 0.2378 | -0.0505 | -0.0692 |
| 41066 | 0.0993 | 0.1139 | 0.0826 | 0.0153 |
| 41093 | 0.2935 | 0.1475 | -0.1313 | -0.1752 |
| 41098 | 0.1669 | -0.0142 | 0.0579 | 0.1010 |
| 41101 | 0.1742 | 0.0931 | 0.2558 | 0.0544 |
| 41181 | 0.3619 | 0.1822 | 0.1521 | -0.0783 |
| 41190 | 0.2607 | 0.1234 | 0.0357 | 0.0783 |
| 41337 | 0.1841 | 0.0156 | 0.2159 | 0.1236 |
| 41407 | 0.2033 | 0.0939 | 0.2437 | 0.1469 |
| 41632 | 0.2642 | 0.1513 | 0.1553 | 0.0205 |
| 41645 | 0.2524 | 0.0773 | -0.0819 |  |
| 41653 | -0.0460 | -0.0020 | -0.1137 | 0.0506 |
| 41669 | -0.0728 | -0.0029 | 0.0334 | 0.1119 |
| 41690 | 0.1151 | 0.1193 | 0.0190 | 0.1416 |
| 42503 | 0.2553 | 0.1045 | 0.0600 | -0.00983 |
| 42504 | 0.1333 | -0.0560 | -0.1206 | -0.1187 |
| 51693 | 0.1997 | 0.0725 | -0.0334 |  |
| 51717 | 0.1998 | 0.2391 | 0.0325 |  |
| 51731 | 0.2534 | -0.0200 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 51732 | 0.1875 | -0.0100 | -0.0683 | 0.0772 |
| 51742 | 0.2057 | -0.1380 | -0.1095 | 0.0286 |
| 51781 | 0.2811 | -0.0394 | 0.0874 | 0.1686 |
| 51786 | 0.1640 | 0.0213 | 0.1243 | -0.0812 |
| 51787 | 0.3500 | 0.2230 | 0.0849 | -0.1707 |
| 51793 | 0.3382 | 0.1669 | 0.0782 | 0.0234 |
| 51806 | 0.3013 | 0.2396 | 0.1721 | 0.2945 |
| 51817 | 0.3639 | 0.1244 | 0.0426 | -0.0276 |
| 51825 | 0.2683 | 0.0962 | -0.0146 | 0.0823 |
| 51839 | 0.0601 | 0.1050 | 0.0366 | 0.0283 |
| 51843 | -0.0361 | -0.0055 | 0.0191 | 0.0776 |
| 59127 | 0.2984 | 0.0736 | 0.0298 | 0.0617 |
| 59128 | 0.2980 | 0.3114 | 0.1593 | 0.1616 |
| 59137 | 0.0786 | 0.0871 | 0.0980 | 0.1532 |
| 59138 | 0.1432 | 0.2311 | 0.2990 | -0.1454 |
| 59149 | 0.3560 | 0.1274 | -0.1024 | 0.0567 |
| 59155 | 0.2243 | 0.1461 | 0.0779 | 0.0343 |
| 59184 | 0.3020 | 0.0842 | 0.0433 | 0.1266 |
| 59185 | 0.0802 | 0.0670 | 0.0771 | 0.0585 |
| 59217 | 0.1105 | 0.1144 | 0.1114 | 0.1412 |
| 59231 | 0.3649 | 0.3284 | 0.0592 | -0.0674 |
| 59239 | 0.2305 | 0.1432 | 0.0148 | 0.1450 |
| 59267 | 0.0877 | 0.0838 | 0.0872 | -0.0298 |
| 59489 | 0.2272 | 0.0759 | -0.0345 | 0.0000 |
| 60230 | 0.2109 | 0.2620 | 0.0000 | -0.0783 |
| 60240 | -0.0221 | 0.0301 | -0.0949 | -0.0979 |
| 60241 | 0.2474 | 0.0609 | -0.0407 | 0.1290 |
| 60242 | 0.2856 | 0.1924 | 0.1850 | 0.0428 |
| 60250 | 0.1273 | -0.1714 | -0.1301 | -0.1394 |
| 60258 | 0.3100 | 0.2502 | 0.3162 |  |
| 62125 | 0.1018 | -0.0626 | -0.0799 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 62129 | 0.2786 | 0.2183 | 0.0703 | 0.1507 |
| 62132 | 0.2708 | 0.2230 | 0.1634 | 0.1714 |
| 62135 | 0.2983 | 0.1894 | 0.0622 | -0.0243 |
| 62136 | 0.3512 | 0.0556 | 0.0980 | 0.0734 |
| 62140 | 0.3812 | 0.3997 | 0.2932 | 0.1867 |
| 84442 | 0.3069 | 0.2896 | 0.3430 | 0.2690 |
| 90381 | 0.2819 | 0.0743 | 0.0308 | 0.0544 |
| 90382 | 0.0569 | -0.0666 | -0.1631 | -0.1797 |
| 90384 | 0.1036 | 0.0291 | -0.0618 | -0.1006 |
| 90385 | 0.1792 | 0.1182 | 0.0907 | 0.1984 |
| 90386 | 0.2554 | 0.2240 | 0.1738 | 0.0131 |
| 90390 | 0.4318 | 0.2296 | 0.3076 | 0.2261 |
| 90396 | 0.0718 | 0.0581 | 0.2216 | 0.2125 |
| 90397 | 0.1268 | 0.0539 | 0.0679 | 0.0027 |
| 90398 | 0.2795 | 0.1449 | 0.1014 | 0.0765 |
| 90400 | 0.1985 | 0.1838 | 0.2377 | 0.0545 |
| 94445 | 0.2217 | 0.0902 | 0.2386 | 0.2064 |
| 94505 | 0.0600 | 0.0758 | 0.2334 | 0.1571 |
| 94510 | 0.1683 | 0.2604 | 0.0842 | 0.2566 |
| 94566 | 0.3157 | 0.1853 | 0.1088 | -0.0114 |
| 94582 | 0.3808 | 0.2727 | 0.2503 | 0.0715 |
| 94586 | 0.3209 | 0.1767 | 0.1205 | 0.0739 |
| 108131 | 0.1744 | -0.0082 | -0.0508 | -0.1536 |
| 108132 | 0.2818 | -0.0358 | 0.0473 | 0.0858 |
| 108133 | 0.1783 | 0.1990 | 0.2375 | 0.2328 |
| 108134 | 0.2241 | 0.0081 | 0.0877 | 0.0675 |
| 108135 | 0.0995 | -0.0635 | -0.0919 | 0.0243 |
| 108136 | 0.2619 | 0.0560 | 0.0854 | 0.0194 |
| 108137 | 0.1924 | 0.1588 | 0.2087 | 0.0000 |
| 108138 | 0.3920 | 0.1531 | 0.1059 | 0.1550 |
| 108139 | 0.2255 | 0.0048 | 0.1009 | 0.0857 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108140 | 0.1381 | 0.0922 | 0.0667 | 0.0519 |
| 108141 | 0.0964 | -0.1230 | -0.0641 | 0.1657 |
| 108142 | 0.2794 | 0.1372 | 0.0225 | -0.0360 |
| 108143 | 0.4001 | 0.2199 | 0.3057 | 0.3272 |
| 108144 | 0.2980 | 0.2069 | 0.3116 | 0.3302 |
| 108145 | 0.1484 | 0.1023 | 0.0881 | 0.1930 |
| 108146 | 0.3033 | 0.2080 | 0.3127 | 0.2948 |
| 108147 | 0.1831 | 0.1017 | 0.1227 | 0.0598 |
| 108148 | 0.1280 | -0.0069 | 0.0870 | 0.0808 |
| 108149 | -0.0003 | 0.0100 | -0.1872 | -0.1751 |
| 108150 | 0.1513 | 0.0463 | 0.0080 | -0.2060 |
| 108151 | 0.1869 | 0.0749 | 0.0054 | 0.0814 |
| 108152 | 0.3092 | -0.0013 | 0.0284 | 0.1362 |
| 108153 | 0.4983 | 0.3790 | 0.4198 | 0.2038 |
| 108154 | 0.3769 | 0.1944 | 0.2969 | 0.2564 |
| 108155 | 0.3370 | -0.0389 | -0.0162 | 0.0638 |
| 108156 | 0.3039 | 0.0818 | 0.1233 | 0.0519 |
| 108157 | 0.2223 | 0.1048 | 0.0668 | 0.2663 |
| 108158 | 0.1994 | 0.0958 | 0.0823 | 0.1313 |
| 108159 | 0.0479 | -0.0489 | -0.0332 | -0.0308 |
| 108160 | -0.0783 | -0.0238 | -0.0248 | 0.2878 |
| 108161 | -0.0186 | 0.1776 | 0.2220 | 0.1996 |
| 108162 | 0.3128 | 0.2078 | 0.2890 | 0.3717 |
| 108163 | 0.2605 | 0.1899 | 0.1070 | 0.1781 |
| 108164 | 0.2834 | 0.2282 | 0.0387 | -0.0213 |
| 108165 | -0.0191 | 0.0059 | 0.0416 | -0.0638 |
| 108166 | 0.3038 | 0.1530 | 0.0850 | 0.1153 |
| 108167 | 0.3061 | -0.0156 | -0.0932 | 0.1550 |
| 108168 | 0.2278 | -0.0166 | -0.0573 | -0.0898 |
| 108169 | -0.0768 | -0.2375 | -0.2086 | -0.2323 |
| 108170 | 0.2778 | 0.1576 | -0.0179 | -0.0432 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  | 0.03404 |
| 108171 | 0.1082 | 0.1928 | 0.0234 | 0.0898 |
| 108172 | 0.0626 | 0.0753 | 0.2478 | 0.0885 |
| 108174 | 0.2470 | 0.1042 | -0.0603 | -0.0777 |
| 108175 | 0.1734 | -0.0273 | 0.0518 | 0.0909 |
| 108176 | 0.2533 | 0.1767 | 0.2413 | 0.2623 |
| 108177 | 0.3077 | 0.1837 | -0.0403 | -0.1178 |
| 108178 | 0.3764 | 0.1079 | 0.1938 | 0.1592 |
| 108179 | 0.4233 | 0.2860 | 0.1089 | -0.0271 |
| 108180 | 0.0484 | 0.1223 | 0.0541 | 0.0049 |
| 108181 | 0.3023 | 0.2827 | -0.1103 | -0.1758 |
| 108182 | 0.2335 | -0.1106 | 0.1049 | 0.1034 |
| 108183 | 0.3435 | 0.1203 | 0.0199 | 0.0007 |
| 108184 | 0.2177 | 0.0238 | 0.1117 | 0.0898 |
| 108185 | 0.2636 | 0.1160 | -0.0452 | -0.1867 |
| 108186 | -0.2086 | -0.0782 | -0.2581 | 0.0000 |
| 108187 | -0.1053 | -0.1932 | 0.3405 | 0.1056 |
| 108188 | 0.2966 | 0.2050 | -0.0594 | -0.1239 |
| 108189 | 0.2519 | -0.0577 | 0.2531 | 0.2961 |
| 108190 | 0.3518 | 0.2695 | 0.0810 | 0.1066 |
| 108191 | 0.2102 | 0.1495 | 0.1738 | 0.1046 |
| 108192 | 0.2098 | 0.2354 | -0.0668 | -0.2116 |
| 108193 | 0.2053 | 0.0072 | 0.2892 | -0.0176 |
| 108194 | 0.2459 | 0.2536 | 0.1037 | 0.3159 |
| 108195 | 0.1967 | 0.1732 | 0.0877 | 0.1234 |
| 108196 | 0.1906 | 0.0752 | 0.2598 | 0.1426 |
| 108197 | 0.2824 | 0.2447 | 0.3477 | 0.3605 |
| 108198 | 0.2556 | 0.3347 | 0.1668 | 0.1388 |
| 108199 | 0.2492 | 0.0750 | 0.0254 | 0.0854 |
| 108200 | 0.2099 | 0.0096 | 0.1655 | 0.3142 |
| 108201 | 0.1870 | 0.0532 |  |  |
|  | 0.3032 | 0.2991 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108202 | -0.0915 | -0.0465 | 0.0356 | 0.1872 |
| 108203 | -0.0080 | 0.1403 | 0.0820 | 0.1054 |
| 108204 | 0.3355 | 0.0802 | -0.0470 | -0.1009 |
| 108477 | 0.3690 | 0.2985 | 0.1924 | 0.3385 |
| 108478 | 0.2129 | 0.2447 | 0.2236 | 0.3502 |
| 108479 | 0.2899 | 0.1261 | 0.1693 | 0.0900 |
| 108480 | 0.1822 | -0.0022 | -0.0028 | -0.1547 |
| 108481 | 0.2634 | 0.1822 | 0.0962 | 0.2431 |
| 108482 | 0.2746 | 0.0888 | -0.0717 | -0.0044 |
| 108483 | 0.3192 | 0.0077 | -0.0422 | 0.0113 |
| 108484 | 0.2636 | 0.0241 | -0.0433 | -0.1182 |
| 108485 | 0.1653 | 0.0430 | -0.0136 | 0.0383 |
| 108486 | 0.4052 | 0.3345 | 0.4211 | 0.3361 |
| 108487 | 0.3372 | 0.2391 | 0.2514 | -0.0842 |
| 108488 | 0.1810 | 0.3028 | 0.1783 | 0.0559 |
| 108489 | 0.3494 | 0.1110 | 0.0574 | 0.0238 |
| 108490 | 0.1444 | 0.0764 | 0.1009 | 0.1119 |
| 108491 | 0.0475 | 0.0339 | 0.0567 | 0.1032 |
| 108492 | 0.4010 | 0.2650 | 0.2272 | 0.1414 |
| 108493 | 0.1714 | 0.0953 | 0.0085 | -0.0257 |
| 108494 | 0.3177 | 0.2071 | 0.1362 | 0.1086 |
| 108495 | 0.1151 | 0.0099 | -0.0931 | -0.1405 |
| 108497 | 0.2921 | 0.0552 | 0.0404 | 0.0257 |
| 108498 | 0.2975 | 0.0388 | -0.0977 | -0.1106 |
| 108499 | 0.2863 | 0.0663 | 0.0319 | 0.0551 |
| 108500 | 0.1575 | 0.0332 | 0.0608 | -0.0183 |
| 108501 | 0.0743 | 0.1760 | 0.1984 | 0.0823 |
| 108502 | 0.1566 | 0.1520 | 0.1028 | 0.2247 |
| 108503 | 0.1387 | 0.1494 | 0.0831 | -0.0423 |
| 108504 | 0.1001 | 0.1442 | 0.1800 | 0.0733 |
| 108505 | 0.2677 | 0.0876 | 0.1113 | 0.0872 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 3 (continued)
Item Discrimination Values for Examination $2-C_{X 3}$

| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  | 0.1950 | 0.1758 |
| 108506 | 0.3136 | 0.1621 | 0.0021 | -0.0423 |
| 108507 | 0.3458 | 0.1042 | 0.0883 | 0.1295 |
| 108508 | 0.2697 | 0.1154 | -0.0651 | -0.1140 |
| 108510 | 0.3294 | 0.1092 | 0.0150 | -0.0506 |
| 108511 | 0.1560 | 0.1441 | -0.1353 | -0.0928 |
| 108512 | 0.0278 | -0.0767 | 0.1299 | 0.0605 |
| 108513 | 0.2234 | 0.1574 | 0.0230 | 0.0097 |
| 108514 | 0.1451 | 0.0429 | 0.0273 | 0.0174 |
| 108515 | 0.2470 | 0.1620 | 0.0232 | 0.1550 |
| 108523 | 0.3178 | 0.1897 | -0.0043 | 0.0472 |
| 108524 | 0.1711 | 0.1200 | 0.1676 | -0.0134 |
| 108525 | 0.1976 | 0.0942 | -0.1953 | 0.0201 |
| 950671494 | 0.0897 | -0.1087 | 0.0357 | 0.0349 |
|  | 0.1338 | 0.0129 |  |  |

Table B. 4
Item Discrimination Values for Examination $2-C_{X 4}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 32497 | 0.2411 | -0.1008 | -0.1401 | -0.2282 |
| 40655 | 0.3516 | 0.2024 | 0.1856 | 0.0717 |
| 40657 | 0.3651 | 0.1839 | 0.1530 | 0.3669 |
| 40668 | 0.3077 | 0.2417 | 0.3198 | 0.1534 |
| 40885 | 0.4001 | 0.3218 | 0.4059 | 0.4472 |
| 40903 | 0.1012 | 0.0574 | 0.0744 | -0.0954 |
| 40907 | 0.0642 | 0.1053 | 0.0294 | -0.0912 |
| 40916 | 0.0227 | 0.0585 | 0.1114 | 0.2481 |
| 40944 | 0.2750 | 0.2208 | 0.2284 | 0.1060 |
| 40969 | 0.3391 | 0.1734 | 0.0080 | -0.1036 |
| 40992 | 0.3734 | 0.2252 | 0.1796 | 0.3821 |
| 41054 | 0.0845 | 0.2010 | 0.2521 | 0.2333 |
| 41059 | 0.1457 | 0.1414 | 0.0000 | 0.0000 |
| 41066 | 0.0993 | -0.0779 | 0.0920 | 0.0348 |
| 41093 | 0.2935 | 0.0969 | 0.1143 | 0.1182 |
| 41098 | 0.1669 | 0.1335 | 0.1465 | -0.2560 |
| 41101 | 0.1742 | 0.1160 | 0.2201 | 0.1946 |
| 41181 | 0.3619 | 0.1414 | 0.1649 | 0.2030 |
| 41190 | 0.2607 | 0.1517 | 0.2625 | 0.1776 |
| 41337 | 0.1841 | 0.0847 | 0.0403 | 0.1182 |
| 41407 | 0.2033 | 0.0895 | 0.1256 | -0.1482 |
| 41632 | 0.2642 | 0.1575 | -0.0573 | -0.1524 |
| 41645 | 0.2524 | 0.2136 | 0.1791 | 0.2934 |
| 41653 | -0.0460 | -0.0421 | -0.1090 | 0.0049 |
| 41669 | -0.0728 | 0.0424 | -0.0025 | 0.1698 |
| 41690 | 0.1151 | 0.0430 | 0.0488 | 0.0249 |
| 42503 | 0.2553 | 0.1879 | 0.1130 | 0.1673 |
| 42504 | 0.1333 | 0.1102 | 0.1002 | -0.1149 |
| 51693 | 0.1997 | 0.2041 | 0.1284 | 0.1210 |
| 51717 | 0.1998 | -0.0286 | -0.0828 | -0.1889 |
| 51731 | 0.2534 | 0.1455 | 0.1492 | 0.1021 |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 51732 | 0.1875 | 0.0911 | 0.1220 | 0.2345 |
| 51742 | 0.2057 | 0.1094 | 0.0485 | 0.1574 |
| 51781 | 0.2811 | 0.2842 | 0.1761 | 0.1674 |
| 51786 | 0.1640 | -0.0971 | -0.0056 | 0.0314 |
| 51787 | 0.3500 | 0.3534 | 0.3955 | 0.3739 |
| 51793 | 0.3382 | 0.1735 | 0.3432 | 0.3017 |
| 51806 | 0.3013 | 0.0307 | 0.1447 | 0.0063 |
| 51817 | 0.3639 | 0.1423 | 0.1080 | -0.0523 |
| 51825 | 0.2683 | 0.3170 | 0.2920 | 0.2039 |
| 51839 | 0.0601 | -0.1026 | -0.1554 | -0.0830 |
| 51843 | -0.0361 | -0.1070 | -0.0823 | 0.0514 |
| 59127 | 0.2984 | 0.1170 | 0.0565 | 0.0433 |
| 59128 | 0.2980 | 0.1327 | 0.2331 | 0.1400 |
| 59137 | 0.0786 | 0.1313 | 0.0000 | 0.0000 |
| 59138 | 0.1432 | 0.0000 | 0.0000 | 0.0000 |
| 59149 | 0.3560 | 0.2238 | 0.2247 | 0.1428 |
| 59155 | 0.2243 | -0.0026 | 0.0602 | 0.0473 |
| 59184 | 0.3020 | 0.1979 | 0.1827 | 0.2700 |
| 59185 | 0.0802 | -0.0891 | -0.0994 | -0.1515 |
| 59217 | 0.1105 | 0.1360 | 0.0422 | -0.0271 |
| 59231 | 0.3649 | 0.1058 | 0.1398 | 0.1596 |
| 59239 | 0.2305 | 0.1157 | 0.2070 | -0.1744 |
| 59267 | 0.0877 | 0.1414 | 0.0000 | 0.0000 |
| 59489 | 0.2272 | 0.1015 | 0.1885 | 0.1037 |
| 60230 | 0.2109 | 0.0000 | 0.0000 | 0.0000 |
| 60240 | -0.0221 | 0.1729 | 0.1333 | 0.1596 |
| 60241 | 0.2474 | 0.0201 | 0.0565 | 0.0173 |
| 60242 | 0.2856 | 0.2059 | 0.1876 | 0.1013 |
| 60250 | 0.1273 | 0.0859 | 0.0688 | 0.1171 |
| 60258 | 0.3100 | 0.0559 | 0.0115 | 0.2931 |
| 62125 | 0.1018 | -0.0646 | -0.0337 | 0.2210 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 62129 | 0.2786 | 0.0858 | 0.0622 | -0.0131 |
| 62132 | 0.2708 | 0.1623 | 0.1133 | 0.1231 |
| 62135 | 0.2983 | 0.2405 | 0.1654 | -0.0223 |
| 62136 | 0.3512 | 0.2524 | 0.1805 | 0.2300 |
| 62140 | 0.3812 | 0.1665 | 0.0158 | 0.0020 |
| 84442 | 0.3069 | 0.1943 | 0.0531 | 0.1410 |
| 90381 | 0.2819 | 0.2079 | 0.1367 | 0.1534 |
| 90382 | 0.0569 | 0.0482 | 0.0730 | -0.0112 |
| 90384 | 0.1036 | 0.0004 | 0.0212 | -0.0591 |
| 90385 | 0.1792 | 0.0780 | -0.0655 | -0.0012 |
| 90386 | 0.2554 | 0.1484 | 0.1108 | 0.0742 |
| 90390 | 0.4318 | 0.2234 | 0.1359 | 0.0163 |
| 90396 | 0.0718 | -0.1581 | -0.0904 | -0.1955 |
| 90397 | 0.1268 | 0.0713 | 0.0313 | -0.1344 |
| 90398 | 0.2795 | 0.1627 | 0.2002 | 0.1115 |
| 90400 | 0.1985 | 0.0562 | 0.1426 | 0.1210 |
| 94445 | 0.2217 | -0.0389 | -0.0144 | 0.0271 |
| 94505 | 0.0600 | -0.0071 | -0.0573 | -0.1349 |
| 94510 | 0.1683 | 0.1897 | 0.0337 | 0.1053 |
| 94566 | 0.3157 | 0.1712 | 0.1686 | 0.1255 |
| 94582 | 0.3808 | 0.2002 | 0.0257 | -0.0289 |
| 94586 | 0.3209 | 0.2932 | 0.1490 | 0.1285 |
| 108131 | 0.1744 | -0.0416 | 0.0644 | 0.1593 |
| 108132 | 0.2818 | 0.3513 | 0.2107 | 0.3229 |
| 108133 | 0.1783 | -0.0720 | -0.1758 | -0.1673 |
| 108134 | 0.2241 | 0.1065 | 0.0725 | 0.0061 |
| 108135 | 0.0995 | 0.2004 | 0.1655 | 0.1131 |
| 108136 | 0.2619 | 0.0481 | 0.0615 | -0.0040 |
| 108137 | 0.1924 | 0.0000 | 0.0000 | 0.0000 |
| 108138 | 0.3920 | 0.2367 | 0.2288 | 0.1489 |
| 108139 | 0.2255 | -0.0332 | 0.0412 | 0.1318 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108140 | 0.1381 | 0.1026 | -0.0197 | 0.0418 |
| 108141 | 0.0964 | 0.0659 | 0.1310 | 0.1493 |
| 108142 | 0.2794 | 0.3519 | 0.1847 | 0.1679 |
| 108143 | 0.4001 | 0.1842 | 0.1018 | -0.0147 |
| 108144 | 0.2980 | 0.1904 | 0.0910 | 0.1319 |
| 108145 | 0.1484 | 0.0476 | -0.0152 | 0.1043 |
| 108146 | 0.3033 | 0.0713 | 0.0348 | -0.0854 |
| 108147 | 0.1831 | 0.1117 | 0.0521 | -0.0132 |
| 108148 | 0.1280 | 0.0108 | -0.0666 | 0.1542 |
| 108149 | -0.0003 | 0.0812 | 0.0241 | 0.1013 |
| 108150 | 0.1513 | 0.1792 | 0.1603 | 0.0762 |
| 108151 | 0.1869 | -0.0003 | 0.1200 | 0.1543 |
| 108152 | 0.3092 | 0.0498 | 0.0744 | -0.0062 |
| 108153 | 0.4983 | 0.2693 | 0.2240 | 0.0823 |
| 108154 | 0.3769 | 0.1593 | 0.0820 | 0.1272 |
| 108155 | 0.3370 | 0.2547 | 0.1597 | 0.2285 |
| 108156 | 0.3039 | 0.2275 | 0.1806 | 0.0783 |
| 108157 | 0.2223 | -0.0063 | 0.0653 | 0.0528 |
| 108158 | 0.1994 | 0.2125 | 0.1772 | 0.1238 |
| 108159 | 0.0479 | 0.0914 | 0.1506 | 0.2450 |
| 108160 | -0.0783 | -0.0899 | -0.1264 | -0.0281 |
| 108161 | -0.0186 | -0.1408 | -0.1608 | -0.1014 |
| 108162 | 0.3128 | 0.2290 | 0.0064 | 0.0745 |
| 108163 | 0.2605 | 0.1653 | 0.1467 | 0.0000 |
| 108164 | 0.2834 | 0.2057 | 0.0677 | 0.0249 |
| 108165 | -0.0191 | -0.0050 | 0.0561 | 0.1370 |
| 108166 | 0.3038 | 0.1810 | 0.1079 | 0.1015 |
| 108167 | 0.3061 | 0.2761 | 0.1472 | 0.2005 |
| 108168 | 0.2278 | 0.1829 | 0.1990 | 0.0000 |
| 108169 | -0.0768 | -0.0341 | 0.0387 | 0.1474 |
| 108170 | 0.2778 | 0.1988 | 0.1933 | 0.1731 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

|  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108171 | 0.1082 | 0.1432 | -0.0242 | -0.0491 |
| 108172 | 0.0626 | -0.0108 | 0.0098 | 0.0315 |
| 108173 | 0.2470 | 0.1681 | 0.1064 | 0.0085 |
| 108174 | 0.1734 | 0.3298 | 0.2310 | 0.1517 |
| 108175 | 0.2533 | 0.1238 | 0.2669 | 0.0738 |
| 108176 | 0.3077 | 0.0872 | 0.0308 | 0.1037 |
| 108177 | 0.3764 | 0.3003 | 0.3243 | 0.3186 |
| 108178 | 0.4233 | 0.2182 | 0.1827 | 0.0244 |
| 108179 | 0.0484 | 0.0225 | 0.0137 | 0.0565 |
| 108180 | 0.3023 | 0.2007 | 0.2488 | 0.1346 |
| 108181 | 0.2335 | 0.1870 | 0.1664 | -0.0013 |
| 108182 | 0.3435 | 0.0929 | 0.0483 | 0.2210 |
| 108183 | 0.2177 | 0.1948 | 0.1532 | 0.1372 |
| 108184 | 0.2636 | 0.1513 | 0.1076 | -0.0223 |
| 108185 | -0.2086 | -0.0682 | -0.0959 | -0.0743 |
| 108186 | -0.1053 | 0.0000 | 0.0000 | 0.0000 |
| 108187 | 0.2966 | 0.0434 | 0.0980 | 0.1423 |
| 108188 | 0.2519 | 0.1026 | 0.1851 | 0.1029 |
| 108189 | 0.3518 | 0.1461 | -0.0221 | -0.0237 |
| 108190 | 0.2102 | 0.0417 | 0.0926 | 0.0042 |
| 108191 | 0.2098 | 0.1245 | -0.0994 | -0.1515 |
| 108192 | 0.2053 | 0.1081 | 0.1041 | 0.1860 |
| 108193 | 0.2459 | 0.2519 | 0.1430 | -0.0174 |
| 108194 | 0.1967 | 0.0047 | 0.0553 | -0.0362 |
| 108195 | 0.1906 | 0.0988 | 0.2072 | 0.1875 |
| 108196 | 0.2824 | -0.1096 | -0.0533 | -0.2168 |
| 108197 | 0.2556 | 0.1955 | 0.0809 | 0.0991 |
| 108198 | 0.2492 | 0.1723 | 0.1878 | 0.1237 |
| 108199 | 0.2099 | 0.0955 | 0.1458 | 0.0711 |
| 108200 | 0.1870 | 0.1393 | 0.1453 | 0.0318 |
| 108201 | 0.3032 | 0.0130 | 0.0469 | 0.1239 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 108202 | -0.0915 | -0.1888 | -0.3167 | -0.1229 |
| 108203 | -0.0080 | -0.0621 | -0.0586 | -0.2977 |
| 108204 | 0.3355 | 0.1474 | 0.1512 | 0.1803 |
| 108477 | 0.3690 | 0.3697 | 0.1923 | 0.0364 |
| 108478 | 0.2129 | 0.2502 | 0.0510 | 0.3460 |
| 108479 | 0.2899 | 0.1500 | 0.1359 | 0.1131 |
| 108480 | 0.1822 | 0.1460 | 0.2189 | 0.0798 |
| 108481 | 0.2634 | 0.1928 | 0.2158 | 0.2609 |
| 108482 | 0.2746 | 0.2497 | 0.1659 | -0.0067 |
| 108483 | 0.3192 | 0.1324 | 0.0978 | -0.1615 |
| 108484 | 0.2636 | 0.1410 | 0.2098 | 0.1931 |
| 108485 | 0.1653 | 0.1035 | 0.0721 | -0.0972 |
| 108486 | 0.4052 | 0.1651 | 0.0009 | 0.0768 |
| 108487 | 0.3372 | 0.2592 | 0.2063 | -0.0056 |
| 108488 | 0.1810 | 0.0152 | 0.1042 | 0.0509 |
| 108489 | 0.3494 | 0.2902 | 0.1583 | 0.0959 |
| 108490 | 0.1444 | 0.1550 | 0.0959 | 0.1760 |
| 108491 | 0.0475 | -0.0197 | 0.0221 | 0.1300 |
| 108492 | 0.4010 | 0.1912 | 0.0415 | 0.0000 |
| 108493 | 0.1714 | -0.0034 | 0.0138 | -0.0362 |
| 108494 | 0.3177 | 0.1256 | 0.0789 | 0.0951 |
| 108495 | 0.1151 | 0.0472 | 0.1000 | 0.0812 |
| 108497 | 0.2921 | 0.2616 | 0.1935 | 0.1212 |
| 108498 | 0.2975 | 0.3111 | 0.3042 | 0.2644 |
| 108499 | 0.2863 | 0.1936 | 0.3151 | 0.2911 |
| 108500 | 0.1575 | 0.0938 | 0.1467 | 0.0609 |
| 108501 | 0.0743 | 0.0325 | -0.0553 | -0.1424 |
| 108502 | 0.1566 | 0.2182 | 0.0242 | 0.1786 |
| 108503 | 0.1387 | 0.0240 | 0.0631 | -0.0438 |
| 108504 | 0.1001 | 0.1514 | 0.0000 | 0.0000 |
| 108505 | 0.2677 | 0.0806 | 0.0379 | -0.0597 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 4 (continued)
Item Discrimination Values for Examination $2-C_{X 4}$

| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  | 0.1717 | -0.0784 |
| 108506 | 0.3136 | 0.2060 | 0.2849 | 0.1295 |
| 108507 | 0.3458 | 0.2625 | 0.1897 | 0.1565 |
| 108508 | 0.2697 | 0.1941 | 0.1687 | 0.0116 |
| 108510 | 0.3294 | 0.1869 | -0.0027 | -0.1226 |
| 108511 | 0.1560 | -0.0326 | 0.0915 | 0.0325 |
| 108512 | 0.0278 | 0.0907 | 0.1938 | 0.0063 |
| 108513 | 0.2234 | 0.1046 | 0.0894 | 0.0517 |
| 108514 | 0.1451 | 0.0370 | 0.0654 | 0.0713 |
| 108515 | 0.2470 | 0.0134 | 0.2189 | 0.2571 |
| 108523 | 0.3178 | 0.3937 | -0.0464 | -0.2090 |
| 108524 | 0.1711 | -0.0392 | 0.0412 | -0.0318 |
| 108525 | 0.1976 | 0.1472 | 0.0481 | 0.3214 |
| 950671494 | 0.0897 | -0.0182 | 0.0220 | 0.0102 |
|  | 0.1338 | 0.1095 |  |  |

Table B. 5
Item Discrimination Values for Examination $2-C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 32497 | 0.2411 | 0.1455 | 0.0719 | 0.1950 |
| 40655 | 0.3516 | 0.0823 | -0.0032 | 0.1669 |
| 40657 | 0.3651 | 0.1211 | 0.3017 | 0.2124 |
| 40668 | 0.3077 | 0.0986 | 0.1257 | 0.1358 |
| 40885 | 0.4001 | -0.0674 | 0.0744 | 0.3156 |
| 40903 | 0.1012 | -0.0320 | -0.0443 | 0.1561 |
| 40907 | 0.0642 | 0.0034 | -0.0808 | -0.0068 |
| 40916 | 0.0227 | 0.0391 | 0.0740 | 0.0379 |
| 40944 | 0.2750 | 0.1725 | 0.0334 | 0.2342 |
| 40969 | 0.3391 | 0.3172 | 0.2639 | 0.1303 |
| 40992 | 0.3734 | 0.1655 | 0.2562 | 0.2607 |
| 41054 | 0.0845 | -0.2639 | -0.1633 | -0.3825 |
| 41059 | 0.1457 | 0.1933 | 0.1991 | 0.0305 |
| 41066 | 0.0993 | -0.0325 | 0.0260 | 0.2137 |
| 41093 | 0.2935 | 0.1652 | 0.1959 | 0.2236 |
| 41098 | 0.1669 | -0.0850 | -0.2690 | -0.3012 |
| 41101 | 0.1742 | 0.0442 | 0.0536 | 0.1867 |
| 4181 | 0.3619 | 0.1673 | 0.1751 | 0.3987 |
| 41190 | 0.2607 | 0.1477 | 0.0450 | 0.2398 |
| 41337 | 0.1841 | 0.0601 | 0.0064 | -0.2234 |
| 41407 | 0.2033 | 0.0808 | -0.0727 | 0.1544 |
| 41632 | 0.2642 | 0.1875 | 0.0074 | -0.1177 |
| 41645 | 0.2524 | 0.0516 | 0.0026 | 0.0727 |
| 41653 | -0.0460 | -0.0041 | -0.0361 | -0.2491 |
| 41669 | -0.0728 | -0.0318 | -0.0607 | 0.0519 |
| 41690 | 0.1151 | 0.1807 | 0.1765 | -0.1313 |
| 42503 | 0.2553 | 0.0908 | 0.1766 | 0.0068 |
| 42504 | 0.1333 | -0.1274 | -0.2222 | -0.0200 |
| 51693 | 0.1997 | 0.0018 | 0.0946 | 0.0696 |
| 51717 | 0.1998 | 0.2025 | 0.1489 | 0.2491 |
| 51731 | 0.2534 | 0.0265 | -0.0070 | -0.1231 |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.0112 | -0.1338 |
| 51732 | 0.1875 | -0.0135 | -0.0547 | -0.2671 |
| 51742 | 0.2057 | -0.0298 | -0.0498 | -0.1169 |
| 51781 | 0.2811 | 0.0174 | 0.1758 | 0.2887 |
| 51787 | 0.1640 | 0.1006 | 0.0514 | 0.0918 |
| 51793 | 0.3500 | 0.0307 | 0.0589 | 0.3220 |
| 51806 | 0.3382 | 0.1913 | 0.1350 | 0.2362 |
| 51817 | 0.3013 | 0.2184 | -0.0138 | -0.0799 |
| 51825 | 0.3639 | 0.1070 | 0.1415 | 0.0465 |
| 51839 | 0.2683 | 0.0376 | 0.2371 | -0.1575 |
| 51843 | 0.0601 | 0.0974 | 0.1300 | 0.1934 |
| 59127 | -0.0361 | 0.0486 | 0.2855 | 0.0724 |
| 59128 | 0.2984 | 0.2337 | 0.3143 | 0.4245 |
| 59137 | 0.2980 | 0.2475 | -0.0034 | -0.0978 |
| 59138 | 0.0786 | 0.0407 | 0.1887 | 0.2717 |
| 59149 | 0.1432 | 0.1859 | 0.0136 | 0.0514 |
| 59155 | 0.3560 | -0.0121 | 0.1666 | 0.3158 |
| 59184 | 0.2243 | 0.1062 | 0.1233 | -0.0007 |
| 59185 | 0.3020 | 0.1532 | 0.0065 | -0.0503 |
| 59217 | 0.0802 | 0.2104 | 0.0920 | 0.0503 |
| 59231 | 0.1105 | 0.1335 | 0.3572 | -0.0034 |
| 59239 | 0.3649 | 0.2531 | -0.0436 | -0.0453 |
| 59267 | 0.2305 | 0.1026 | -0.0407 | -0.1959 |
| 59489 | 0.0877 | 0.0245 | 0.1705 | -0.0683 |
| 60230 | 0.2272 | 0.1735 | 0.3083 | 0.0000 |
| 60240 | 0.2109 | 0.2482 | -0.0543 | -0.2751 |
| 60241 | -0.0221 | -0.0886 | 0.1165 | 0.0876 |
| 60242 | 0.2474 | 0.0994 | -0.0060 |  |
| 60250 | 0.2856 | 0.2061 | 0.2083 |  |
| 60258 | 0.1273 | -0.0815 | 0.363 | 0.2751 |
| 62125 | 0.3100 | 0.2682 | 0.0513 |  |
|  | 0.1018 | -0.0414 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.0594 | 0.0332 |
| 62129 | 0.2786 | 0.1163 | 0.0910 | 0.0404 |
| 62132 | 0.2708 | 0.1170 | -0.0032 | -0.0355 |
| 62135 | 0.2983 | 0.0973 | 0.1278 | 0.0089 |
| 62136 | 0.3512 | 0.1218 | 0.2699 | 0.1309 |
| 62140 | 0.3812 | 0.2872 | 0.2906 | 0.2346 |
| 84442 | 0.3069 | 0.2408 | 0.0455 | -0.1277 |
| 90381 | 0.2819 | 0.1343 | -0.1320 | -0.1776 |
| 90382 | 0.0569 | -0.0280 | -0.0726 | -0.0088 |
| 90384 | 0.1036 | -0.1136 | 0.1804 | -0.1209 |
| 90385 | 0.1792 | 0.1394 | 0.2017 | 0.3749 |
| 90386 | 0.2554 | 0.2424 | 0.1077 | 0.2671 |
| 90390 | 0.4318 | 0.2959 | 0.1217 | 0.3401 |
| 90396 | 0.0718 | 0.0983 | -0.0422 | 0.1384 |
| 90397 | 0.1268 | 0.1249 | 0.1172 | 0.2108 |
| 90398 | 0.2795 | 0.2431 | 0.2266 | 0.2519 |
| 90400 | 0.1985 | 0.1710 | 0.3960 | 0.2101 |
| 94445 | 0.2217 | 0.4427 | 0.0461 | 0.1491 |
| 94505 | 0.0600 | 0.0104 | 0.2164 | -0.1702 |
| 94510 | 0.1683 | 0.1729 | 0.0818 | 0.3680 |
| 94566 | 0.3157 | 0.1377 | 0.2818 | 0.1035 |
| 94582 | 0.3808 | 0.2140 | -0.1347 |  |
| 94586 | 0.3209 | 0.1527 | 0.2231 | 0.0936 |
| 108131 | 0.1744 | -0.0580 | 0.0130 | -0.1697 |
| 108132 | 0.2818 | 0.0460 | 0.0301 | 0.1170 |
| 108133 | 0.1783 | 0.1810 | 0.1927 | -0.1358 |
| 108134 | 0.2241 | 0.1095 | -0.0471 | -0.1798 |
| 108135 | 0.0995 | -0.1453 | -0.2013 | 0.1430 |
| 108136 | 0.2619 | 0.1840 | 0.0389 | 0.2359 |
| 108137 | 0.1924 | 0.1353 | 0.1485 | 0.1073 |
| 108138 | 0.3920 | 0.1862 | 0.0376 | 0.0558 |
| 108139 | 0.2255 | 0.0782 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 108140 | 0.1381 | 0.1397 | 0.1336 | -0.0994 |
| 108141 | 0.0964 | -0.1663 | -0.2452 | -0.0174 |
| 108142 | 0.2794 | 0.0537 | 0.0016 | -0.3701 |
| 108143 | 0.4001 | 0.1823 | 0.0062 | 0.1304 |
| 108144 | 0.2980 | 0.2150 | 0.2450 | 0.2240 |
| 108145 | 0.1484 | 0.1164 | 0.2684 | -0.0231 |
| 108146 | 0.3033 | 0.2727 | 0.2381 | 0.1617 |
| 108147 | 0.1831 | 0.1680 | 0.2177 | 0.0757 |
| 108148 | 0.1280 | 0.1068 | 0.1711 | 0.1052 |
| 108149 | -0.0003 | -0.1253 | -0.0344 | -0.2881 |
| 108150 | 0.1513 | 0.0544 | -0.0475 | 0.0453 |
| 108151 | 0.1869 | 0.0793 | 0.1865 | 0.2199 |
| 108152 | 0.3092 | 0.1057 | 0.0774 | 0.1876 |
| 108153 | 0.4983 | 0.4293 | 0.2363 | 0.0383 |
| 108154 | 0.3769 | 0.2079 | 0.2560 | 0.1708 |
| 108155 | 0.3370 | 0.0407 | 0.0878 | -0.0478 |
| 108156 | 0.3039 | 0.0594 | 0.0192 | 0.0818 |
| 108157 | 0.2223 | 0.1718 | 0.2374 | 0.1191 |
| 108158 | 0.1994 | -0.0082 | 0.0242 | 0.0505 |
| 108159 | 0.0479 | -0.0006 | 0.1748 | 0.2690 |
| 108160 | -0.0783 | -0.0678 | -0.0195 | -0.2165 |
| 108161 | -0.0186 | 0.1272 | 0.1053 | 0.1017 |
| 108162 | 0.3128 | 0.2670 | 0.1841 | -0.0058 |
| 108163 | 0.2605 | 0.2152 | 0.1984 | -0.0471 |
| 108164 | 0.2834 | 0.1596 | 0.1022 | -0.1900 |
| 108165 | -0.0191 | 0.0059 | 0.1772 | 0.2704 |
| 108166 | 0.3038 | 0.1606 | 0.0867 | -0.1034 |
| 108167 | 0.3061 | -0.0108 | -0.0587 | -0.2964 |
| 108168 | 0.2278 | 0.0519 | -0.0140 | -0.1576 |
| 108169 | -0.0768 | -0.2135 | -0.2596 | 0.0000 |
| 108170 | 0.2778 | 0.0046 | 0.1606 | 0.2236 |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 108171 | 0.1082 | 0.1585 | 0.0979 | 0.2359 |
| 108172 | 0.0626 | 0.0305 | 0.2162 | 0.0636 |
| 108173 | 0.2470 | 0.1637 | 0.0510 | 0.1172 |
| 108174 | 0.1734 | 0.0132 | -0.0386 | -0.2429 |
| 108175 | 0.2533 | 0.2820 | 0.1294 | -0.0116 |
| 108176 | 0.3077 | 0.2464 | 0.2532 | 0.3877 |
| 108177 | 0.3764 | 0.1122 | 0.1671 | 0.0762 |
| 108178 | 0.4233 | 0.2682 | 0.1575 | -0.0453 |
| 108179 | 0.0484 | 0.1551 | 0.1909 | 0.2500 |
| 108180 | 0.3023 | 0.1128 | 0.1259 | 0.2216 |
| 108181 | 0.2335 | -0.1387 | -0.1553 | -0.2517 |
| 108182 | 0.3435 | 0.2730 | 0.3627 | 0.0859 |
| 108183 | 0.2177 | -0.0041 | -0.0407 | 0.0757 |
| 108184 | 0.2636 | 0.1495 | 0.0030 | -0.0947 |
| 108185 | -0.2086 | 0.0376 | 0.0944 | 0.1268 |
| 108186 | -0.1053 | -0.1729 | -0.2018 | -0.3433 |
| 108187 | 0.2966 | 0.3021 | 0.2034 | 0.4481 |
| 108188 | 0.2519 | 0.0393 | 0.0145 | 0.1430 |
| 108189 | 0.3518 | 0.3965 | 0.2078 | -0.0877 |
| 108190 | 0.2102 | 0.0576 | 0.0461 | 0.2048 |
| 108191 | 0.2098 | 0.3791 | 0.2946 | -0.0046 |
| 108192 | 0.2053 | 0.0340 | 0.0426 | -0.1191 |
| 108193 | 0.2459 | 0.3143 | 0.0569 | -0.1034 |
| 108194 | 0.1967 | 0.1171 | 0.0581 | 0.0936 |
| 108195 | 0.1906 | 0.1102 | 0.0937 | 0.2609 |
| 108196 | 0.2824 | 0.2942 | 0.3130 | 0.2000 |
| 108197 | 0.2556 | 0.3740 | 0.2885 | 0.0603 |
| 108198 | 0.2492 | 0.0932 | 0.0930 | 0.2470 |
| 108199 | 0.2099 | 0.0141 | 0.0305 | 0.2155 |
| 108200 | 0.1870 | 0.1277 | 0.0019 | 0.1978 |
| 108201 | 0.3032 | 0.2072 | 0.4069 | 0.6162 |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 108202 | -0.0915 | 0.0677 | 0.1483 | -0.1814 |
| 108203 | -0.0080 | 0.1077 | 0.0131 | -0.0463 |
| 108204 | 0.3355 | 0.1143 | 0.0650 | -0.0364 |
| 108477 | 0.3690 | 0.1987 | 0.0695 | -0.2171 |
| 108478 | 0.2129 | 0.1263 | 0.1903 | 0.0529 |
| 108479 | 0.2899 | 0.1627 | -0.0603 | 0.1299 |
| 108480 | 0.1822 | -0.0225 | -0.0153 | 0.2430 |
| 108481 | 0.2634 | 0.1603 | 0.2368 | 0.0637 |
| 108482 | 0.2746 | 0.0613 | -0.0918 | -0.3200 |
| 108483 | 0.3192 | 0.1574 | 0.0863 | -0.1462 |
| 108484 | 0.2636 | 0.0333 | 0.0795 | 0.0660 |
| 108485 | 0.1653 | 0.0934 | -0.0257 | -0.0517 |
| 108486 | 0.4052 | 0.3243 | 0.2684 | 0.2781 |
| 108487 | 0.3372 | 0.1355 | 0.0561 | 0.3222 |
| 108488 | 0.1810 | 0.0763 | 0.1285 | 0.2950 |
| 108489 | 0.3494 | 0.1019 | 0.1679 | -0.0158 |
| 108490 | 0.1444 | 0.0800 | 0.1511 | -0.0007 |
| 108491 | 0.0475 | 0.1384 | 0.0872 | 0.0517 |
| 108492 | 0.4010 | 0.2368 | 0.2601 | 0.1209 |
| 108493 | 0.1714 | 0.0956 | -0.0062 | 0.0907 |
| 108494 | 0.3177 | 0.0829 | 0.0058 | 0.1034 |
| 108495 | 0.1151 | -0.0108 | 0.0846 | 0.0258 |
| 108497 | 0.2921 | -0.0344 | 0.0610 | -0.1084 |
| 108498 | 0.2975 | 0.0919 | 0.0699 | -0.1491 |
| 108499 | 0.2863 | 0.0330 | 0.1719 | 0.2751 |
| 108500 | 0.1575 | 0.0747 | 0.0168 | 0.0024 |
| 108501 | 0.0743 | 0.0800 | 0.0890 | 0.1821 |
| 108502 | 0.1566 | 0.0288 | 0.0499 | -0.1074 |
| 108503 | 0.1387 | 0.0232 | 0.1195 | 0.0837 |
| 108504 | 0.1001 | 0.1032 | 0.0852 | 0.0818 |
| 108505 | 0.2677 | 0.0584 | 0.0466 | 0.2470 |

Table B. 5 (continued)
Item Discrimination Values for Examination $2-C_{X 5}$

| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
| :--- | :---: | ---: | ---: | ---: |
|  |  |  |  |  |
|  |  |  | 0.0231 | 0.0046 |
| 108506 | 0.3136 | 0.2281 | -0.0836 | -0.0818 |
| 108507 | 0.3458 | 0.0471 | 0.0012 | -0.0088 |
| 108508 | 0.2697 | 0.2127 | -0.0054 | -0.0198 |
| 108510 | 0.3294 | -0.0356 | 0.0394 | 0.0884 |
| 108511 | 0.1560 | 0.1454 | -0.2088 | -0.2494 |
| 108512 | 0.0278 | -0.1605 | 0.1726 | 0.3128 |
| 108513 | 0.2234 | 0.0308 | 0.0010 | 0.1075 |
| 108514 | 0.1451 | -0.0312 | 0.0910 | -0.1871 |
| 108515 | 0.2470 | 0.1964 | 0.0407 | -0.1912 |
| 108523 | 0.3178 | 0.0297 | 0.1268 | -0.1705 |
| 108524 | 0.1711 | 0.2777 | 0.0819 | 0.0444 |
| 108525 | 0.1976 | 0.2425 | -0.0645 | -0.1934 |
| 950671494 | 0.0897 | -0.1974 | -0.1208 | -0.1023 |
|  | 0.1338 | 0.0413 |  |  |

## APPENDIX C: ITEM DISCRIMINATION VALUES - EXAMINATION 3

Table C. 1
Item Discrimination Values for Examination 3 - $C_{X 1}$

| Item | Unrestricted | +/-1.00 SD | +/- 0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 159995 | -0.1125 | -0.0849 | -0.1583 | 0.1173 |
| 159996 | -0.0017 | 0.2012 | 0.0615 | 0.3932 |
| 159997 | 0.3494 | 0.1084 | -0.0850 | -0.0437 |
| 159998 | 0.3372 | 0.1826 | -0.1561 | -0.2087 |
| 159999 | 0.2592 | 0.2028 | 0.0251 | 0.4287 |
| 160000 | 0.0715 | 0.2218 | 0.0385 | -0.2058 |
| 160001 | -0.0636 | 0.0000 | 0.0000 | 0.0000 |
| 160002 | 0.1160 | 0.0635 | 0.2383 | 0.2460 |
| 160003 | 0.1235 | -0.0551 | -0.1817 | -0.0238 |
| 160004 | -0.0436 | -0.2262 | -0.1338 | 0.2089 |
| 160005 | 0.4647 | 0.2046 | -0.2202 | -0.3709 |
| 160006 | 0.4144 | 0.2992 | -0.0491 | -0.1993 |
| 160007 | 0.1342 | -0.1181 | 0.0926 | 0.1690 |
| 160008 | 0.1732 | -0.0093 | -0.0762 | -0.1010 |
| 160009 | 0.0014 | 0.0050 | -0.1561 | -0.2087 |
| 160010 | 0.1508 | -0.0196 | -0.2227 | -0.1281 |
| 160011 | 0.2503 | 0.1195 | 0.0996 | -0.0437 |
| 160012 | -0.0345 | -0.2062 | -0.3107 | 0.0898 |
| 160013 | 0.2169 | 0.3313 | 0.4271 | 0.2807 |
| 160014 | 0.3356 | 0.3432 | 0.0909 | 0.3746 |
| 160015 | 0.1391 | 0.1384 | 0.0512 | 0.0640 |
| 160016 | 0.1500 | 0.0306 | 0.2584 | 0.1173 |
| 160017 | 0.2353 | -0.1550 | 0.0371 | -0.1400 |
| 160018 | 0.2385 | 0.0206 | 0.0150 | -0.3225 |
| 160019 | 0.1183 | -0.0831 | -0.2196 | -0.2058 |
| 160020 | 0.1056 | 0.3203 | 0.3876 | -0.2919 |
| 160021 | 0.1994 | 0.2292 | 0.0312 | -0.0437 |
| 160022 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160023 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |

Table C. 1 (continued)
Item Discrimination Values for Examination 3 - $C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160024 | -0.0251 | 0.1267 | 0.1364 | -0.1921 |
| 160025 | 0.1141 | 0.2719 | 0.4514 | 0.1852 |
| 160026 | 0.1059 | 0.1562 | 0.1538 | 0.0000 |
| 160027 | 0.0230 | 0.0781 | -0.3647 | -0.0474 |
| 160028 | 0.2308 | 0.1754 | -0.0088 | 0.0530 |
| 160029 | 0.3289 | 0.0913 | -0.0694 | 0.2460 |
| 160030 | 0.2980 | 0.4566 | 0.2573 | -0.0808 |
| 160031 | 0.2154 | 0.1985 | 0.0854 | -0.0808 |
| 160032 | 0.4353 | 0.2039 | -0.0385 | 0.5033 |
| 160033 | -0.0930 | 0.0412 | 0.3579 | 0.2682 |
| 160034 | 0.4090 | 0.2046 | 0.1568 | 0.3612 |
| 160035 | 0.2824 | 0.2668 | 0.2755 | -0.4434 |
| 160036 | 0.0739 | -0.0657 | -0.2227 | 0.0000 |
| 160037 | -0.0380 | 0.0000 | 0.0000 | 0.0615 |
| 160038 | 0.0829 | -0.0984 | 0.0766 | 0.1852 |
| 160039 | 0.1151 | 0.2090 | 0.1538 | 0.3103 |
| 160040 | -0.0036 | 0.0346 | 0.3076 | -0.2260 |
| 160041 | 0.4053 | 0.3504 | 0.0736 | -0.0474 |
| 160042 | 0.1553 | -0.0364 | -0.0921 | 0.2089 |
| 160043 | 0.1101 | 0.0376 | 0.1957 | -0.0810 |
| 160044 | 0.1223 | -0.0023 | -0.1196 | -0.0474 |
| 160045 | 0.1300 | 0.0532 | 0.2916 | -0.0913 |
| 160046 | 0.2300 | -0.1416 | -0.0814 | -0.1139 |
| 160047 | 0.1708 | 0.1643 | -0.0187 | -0.2260 |
| 160048 | 0.2581 | 0.1894 | 0.0461 | -0.0056 |
| 160049 | 0.0586 | -0.1206 | -0.2200 | 0.0000 |
| 160050 | 0.2836 | 0.4618 | 0.3757 | 0.3136 |
| 160051 | 0.3579 | 0.1745 | 0.0694 | -0.2044 |
| 160052 | 0.1241 | 0.0606 | 0.1066 | 0.4263 |
| 160053 | 0.2348 | 0.2348 | 0.1290 | 0.0640 |
| 160054 | 0.1140 | -0.0838 | -0.0442 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 1 (continued)
Item Discrimination Values for Examination 3 - $C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160055 | 0.2920 | 0.2352 | -0.0670 | -0.3746 |
| 160056 | 0.2966 | 0.1313 | 0.2107 | 0.2751 |
| 160057 | 0.3838 | 0.2967 | 0.0563 | -0.3330 |
| 160058 | 0.0012 | -0.2917 | 0.0168 | -0.0692 |
| 160059 | 0.2883 | -0.1033 | 0.1318 | 0.0640 |
| 160060 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160061 | 0.2186 | -0.1030 | 0.2107 | 0.4535 |
| 160062 | 0.1056 | -0.0701 | -0.1561 | -0.2087 |
| 160063 | 0.0365 | 0.0570 | -0.0418 | 0.0898 |
| 160064 | 0.2481 | 0.3302 | 0.1592 | -0.0441 |
| 160065 | -0.0272 | -0.2615 | -0.4212 | -0.1535 |
| 160066 | -0.1009 | 0.0125 | 0.4269 | 0.2461 |
| 160067 | 0.0136 | -0.0093 | 0.1637 | 0.0898 |
| 160068 | 0.0390 | -0.1219 | -0.2824 | 0.0000 |
| 160069 | 0.2224 | -0.0120 | 0.2108 | 0.1364 |
| 160070 | 0.0423 | -0.0533 | 0.3699 | 0.1921 |
| 160071 | 0.1943 | 0.0031 | 0.1559 | -0.0810 |
| 160072 | 0.2186 | 0.2331 | 0.2169 | 0.2262 |
| 160073 | 0.2013 | -0.0126 | 0.0217 | 0.3225 |
| 160074 | 0.3289 | 0.2292 | -0.0088 | 0.0238 |
| 160075 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160076 | 0.1443 | -0.0093 | 0.0000 | 0.0000 |
| 160077 | 0.0059 | 0.1195 | 0.0703 | -0.1281 |
| 160078 | 0.1822 | -0.1186 | -0.0921 | 0.4715 |
| 160079 | 0.3248 | 0.3558 | 0.0563 | 0.0808 |
| 160080 | 0.1056 | -0.1063 | -0.1862 | 0.4535 |
| 160081 | 0.1882 | 0.1195 | 0.1034 | -0.1066 |
| 160082 | 0.1301 | 0.0649 | 0.3123 | -0.0114 |
| 160083 | 0.1376 | 0.0823 | 0.2345 | -0.2935 |
| 160084 | 0.2800 | 0.5687 | 0.3556 | 0.1139 |
| 160085 | 0.1580 | -0.1030 | -0.2944 | 0.0615 |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 1 (continued)
Item Discrimination Values for Examination 3 - $C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | ---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 160086 | 0.1300 | 0.0572 | 0.3435 | 0.2044 |
| 160087 | 0.1955 | 0.0913 | -0.2560 | -0.1535 |
| 160088 | -0.0804 | -0.1961 | -0.2351 | -0.1998 |
| 160089 | 0.2772 | 0.0126 | 0.0619 | 0.4167 |
| 160090 | -0.0300 | 0.1905 | 0.1637 | 0.0898 |
| 160526 | 0.2836 | 0.1977 | 0.2561 | 0.4535 |
| 160527 | 0.2211 | 0.1618 | 0.4232 | 0.0114 |
| 160528 | 0.1479 | -0.0874 | 0.0038 | -0.2919 |
| 160529 | -0.0272 | -0.2582 | -0.2814 | -0.2894 |
| 160530 | 0.1396 | -0.0782 | 0.1867 | 0.2682 |
| 160531 | 0.1950 | 0.0346 | -0.1162 | 0.1087 |
| 160532 | 0.0518 | -0.0864 | -0.2226 | 0.0000 |
| 160533 | 0.2074 | 0.1905 | 0.3911 | 0.1852 |
| 160534 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160535 | 0.2356 | 0.4575 | 0.5232 | 0.3700 |
| 160536 | 0.0713 | 0.2292 | 0.1149 | 0.1377 |
| 160537 | 0.2885 | 0.1491 | 0.2862 | 0.6963 |
| 160538 | 0.1169 | 0.1491 | -0.0533 | -0.1535 |
| 160539 | 0.0901 | 0.1372 | 0.1034 | 0.2460 |
| 160540 | 0.1639 | 0.3961 | 0.3325 | -0.0761 |
| 160541 | 0.3539 | 0.2337 | -0.0217 | 0.0184 |
| 160542 | 0.1824 | -0.1040 | -0.0533 | -0.1400 |
| 160543 | 0.0649 | 0.1152 | 0.2597 | 0.4715 |
| 160544 | 0.1833 | -0.1929 | 0.0000 | 0.0000 |
| 160545 | 0.0954 | 0.1625 | 0.1385 | 0.1139 |
| 160546 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160547 | 0.2811 | 0.0166 | -0.1838 | -0.1111 |
| 160548 | 0.2629 | -0.0236 | 0.1241 | 0.1281 |
| 160549 | 0.2824 | 0.0992 | 0.0303 | 0.2807 |
| 160550 | 0.4042 | 0.1520 | 0.2144 | -0.0471 |
| 160551 | 0.2205 | 0.0731 | 0.5155 | 0.5977 |
|  |  |  |  |  |
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|  |  |  |  |  |

Table C. 1 (continued)
Item Discrimination Values for Examination 3 - $C_{X 1}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.2107 | -0.1965 |
| 160552 | 0.3101 | 0.5026 | -0.0945 | -0.3810 |
| 160553 | 0.1476 | -0.0690 | 0.2835 | -0.0692 |
| 160554 | 0.2015 | 0.2326 | 0.0000 | 0.0000 |
| 160555 | 0.3749 | 0.0000 | -0.1094 | -0.1111 |
| 160556 | 0.2813 | 0.0217 | 0.0996 | -0.0437 |
| 160557 | 0.3548 | 0.3032 | -0.1316 | 0.0238 |
| 160558 | 0.2423 | 0.0606 | -0.0088 | -0.0474 |
| 160559 | 0.2653 | 0.0290 | -0.0058 | -0.3305 |
| 160560 | 0.2133 | 0.0781 | 0.0000 | 0.0000 |
| 160561 | 0.2015 | -0.2284 | 0.0000 | 0.0000 |
| 160562 | -0.0636 | 0.0000 | -0.1162 | -0.1690 |
| 160563 | 0.1022 | -0.0107 | 0.2597 | -0.1281 |
| 160564 | 0.2586 | 0.4015 | 0.1025 | -0.1998 |
| 160565 | 0.1646 | 0.1414 | 0.1318 | 0.0898 |
| 160566 | 0.2540 | 0.0478 | -0.2558 | -0.3873 |
| 160567 | 0.0320 | -0.1033 | 0.1962 | 0.3228 |
| 160568 | 0.0281 | 0.1622 | 0.1637 | 0.1852 |
| 160569 | 0.2538 | 0.0063 | 0.0312 | -0.3330 |
| 160570 | 0.2961 | 0.1416 | -0.0736 | 0.1535 |
| 160571 | 0.1201 | -0.0584 | 0.0461 | 0.0332 |
| 160572 | 0.2302 | 0.0584 | -0.0251 | -0.4287 |
| 160573 | 0.2998 | -0.0015 | 0.0000 | 0.0000 |
| 160574 | 0.2106 | -0.1929 | 0.3550 | 0.1852 |
| 160575 | 0.0609 | 0.1534 | 0.1558 | 0.1066 |
| 160576 | 0.0652 | -0.1116 | 0.1066 | 0.2058 |
| 160577 | 0.1216 | 0.3236 | 0.2561 | 0.4535 |
| 160578 | 0.1544 | 0.1977 |  |  |
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Table C. 2
Item Discrimination Values for Examination 3 - $C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.0655 | -0.0555 |
| 159995 | -0.1125 | -0.0180 | -0.1121 | 0.1118 |
| 159996 | -0.0017 | 0.0173 | -0.0360 | -0.0488 |
| 159997 | 0.3494 | 0.0504 | 0.2302 | 0.1236 |
| 159999 | 0.3372 | 0.2062 | 0.1734 | -0.0075 |
| 160000 | 0.2592 | 0.2640 | 0.2313 | 0.2765 |
| 160001 | 0.0715 | 0.0223 | 0.0000 | 0.0000 |
| 160002 | -0.0636 | 0.0000 | 0.2922 | 0.2142 |
| 160003 | 0.1160 | 0.2518 | -0.1647 | -0.2173 |
| 160004 | 0.1235 | -0.1785 | -0.1274 | -0.3181 |
| 160005 | -0.0436 | -0.0804 | -0.0535 | 0.1118 |
| 160006 | 0.4647 | 0.0249 | 0.2643 | 0.3548 |
| 160007 | 0.4144 | 0.1844 | -0.1442 | -0.1101 |
| 160008 | 0.1342 | -0.1009 | 0.0817 | 0.0674 |
| 160009 | 0.1732 | -0.0133 | 0.2302 | 0.3368 |
| 160010 | 0.0014 | 0.2062 | 0.2298 | 0.3296 |
| 160011 | 0.1508 | 0.2103 | 0.2922 | 0.2550 |
| 160012 | 0.2503 | 0.2637 | 0.0519 | -0.2330 |
| 160013 | -0.0345 | 0.0886 | -0.0657 | 0.2722 |
| 160014 | 0.2169 | 0.1471 | 0.4008 | 0.0674 |
| 160015 | 0.3356 | 0.4352 | -0.1277 |  |
| 160016 | 0.1391 | 0.1174 | -0.1249 | -0.1159 |
| 160017 | 0.1500 | 0.0173 | -0.0664 | 0.0350 |
| 160018 | 0.2353 | 0.0488 | 0.1250 | 0.1416 |
| 160019 | 0.2385 | 0.1029 | 0.0360 | 0.1861 |
| 160020 | 0.1183 | -0.0727 | -0.2356 | 0.2828 |
| 160021 | 0.1056 | -0.1309 | 0.0841 | 0.0477 |
| 160022 | 0.1994 | -0.0539 | 0.0000 | 0.0000 |
| 160023 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160024 | 0.0000 | 0.0000 | 0.0302 | 0.0000 |
| 160025 | -0.0251 | -0.1067 | 0.1028 | 0.1697 |
|  | 0.1141 | 0.1172 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 2 (continued)
Item Discrimination Values for Examination $3-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160026 | 0.1059 | 0.2839 | 0.3496 | -0.2722 |
| 160027 | 0.0230 | 0.0005 | -0.0029 | -0.0824 |
| 160028 | 0.2308 | 0.0482 | 0.1261 | -0.0360 |
| 160029 | 0.3289 | 0.1678 | 0.0137 | 0.1423 |
| 160030 | 0.2980 | 0.2603 | 0.2569 | 0.4330 |
| 160031 | 0.2154 | 0.0522 | 0.2906 | 0.2315 |
| 160032 | 0.4353 | 0.2138 | -0.0952 | -0.2416 |
| 160033 | -0.0930 | -0.1229 | 0.0608 | 0.1976 |
| 160034 | 0.4090 | 0.2019 | 0.3144 | 0.3932 |
| 160035 | 0.2824 | 0.3741 | -0.0211 | 0.0154 |
| 160036 | 0.0739 | 0.0441 | 0.0000 | 0.0000 |
| 160037 | -0.0380 | -0.2140 | 0.0302 | 0.3487 |
| 160038 | 0.0829 | 0.0478 | -0.0018 | -0.0486 |
| 160039 | 0.1151 | 0.1861 | 0.0259 | -0.3030 |
| 160040 | -0.0036 | 0.0972 | 0.1443 | 0.1798 |
| 160041 | 0.4053 | 0.1560 | 0.0440 | 0.0098 |
| 160042 | 0.1553 | 0.0727 | 0.0753 | 0.2547 |
| 160043 | 0.1101 | 0.2098 | -0.0673 | -0.0167 |
| 160044 | 0.1223 | -0.2120 | 0.1451 | 0.1697 |
| 160045 | 0.1300 | 0.1483 | 0.0372 | -0.1179 |
| 160046 | 0.2300 | 0.1044 | 0.3210 | 0.4031 |
| 160047 | 0.1708 | 0.1468 | 0.0888 | 0.0674 |
| 160048 | 0.2581 | 0.0140 | -0.0033 | -0.1537 |
| 160049 | 0.0586 | 0.0440 | 0.0000 | 0.0000 |
| 160050 | 0.2836 | 0.0000 | 0.2585 | -0.0797 |
| 160051 | 0.3579 | 0.3204 | 0.1915 | 0.0974 |
| 160052 | 0.1241 | 0.1211 | -0.10738 | -0.4404 |
| 160053 | 0.2348 | 0.3454 | 0.2910 | 0.5385 |
| 160054 | 0.1140 | -0.1861 | 0.3000 |  |
| 160055 | 0.2920 | 0.3076 | 0.2298 | -21531 |
| 160056 | 0.2966 | 0.2103 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 2 (continued)
Item Discrimination Values for Examination 3 - $C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160057 | 0.3838 | 0.3311 | 0.2519 | 0.3635 |
| 160058 | 0.0012 | -0.0677 | -0.1423 | 0.0619 |
| 160059 | 0.2883 | 0.1162 | 0.0841 | 0.0098 |
| 160060 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160061 | 0.2186 | -0.1268 | -0.4155 | -0.4807 |
| 160062 | 0.1056 | 0.1638 | 0.1663 | 0.2173 |
| 160063 | 0.0365 | 0.2251 | 0.2625 | 0.1171 |
| 160064 | 0.2481 | 0.0229 | 0.1588 | 0.1530 |
| 160065 | -0.0272 | -0.0146 | -0.0052 | -0.1346 |
| 160066 | -0.1009 | 0.1454 | 0.1249 | 0.1159 |
| 160067 | 0.0136 | -0.1268 | 0.0605 | -0.1946 |
| 160068 | 0.0390 | -0.0161 | -0.0458 | -0.1101 |
| 160069 | 0.2224 | 0.1801 | 0.0334 | -0.0824 |
| 160070 | 0.0423 | -0.1425 | -0.2585 | -0.3836 |
| 160071 | 0.1943 | 0.0014 | 0.0465 | 0.0828 |
| 160072 | 0.2186 | 0.3454 | 0.4090 | 0.3380 |
| 160073 | 0.2013 | 0.3213 | 0.1974 | -0.0652 |
| 160074 | 0.3289 | 0.0413 | -0.0380 | 0.1976 |
| 160075 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160076 | 0.1443 | -0.2085 | -0.3414 | -0.5215 |
| 160077 | 0.0059 | -0.1271 | 0.0149 | 0.1179 |
| 160078 | 0.1822 | -0.1752 | -0.3759 | -0.3580 |
| 160079 | 0.3248 | 0.3544 | 0.2830 | 0.2582 |
| 160080 | 0.1056 | -0.0851 | -0.2121 | -0.2453 |
| 160081 | 0.1882 | 0.4100 | 0.4888 | 0.4336 |
| 160082 | 0.1301 | -0.0660 | 0.0097 | -0.1695 |
| 160083 | 0.1376 | 0.0424 | 0.1407 | 0.2765 |
| 160084 | 0.2800 | 0.1017 | -0.0957 | 0.3380 |
| 160085 | 0.1580 | -0.0851 | -0.1800 | -0.0719 |
| 160086 | 0.1300 | 0.0684 | -0.0977 | 0.0555 |
| 160087 | 0.1955 | 0.1308 | 0.0935 | 0.0154 |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 2 (continued)
Item Discrimination Values for Examination $3-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.1655 | 0.0619 |
| 160088 | -0.0804 | -0.3558 | 0.0383 | 0.1261 |
| 160089 | 0.2772 | 0.0621 | 0.2360 | 0.1171 |
| 160526 | -0.0300 | 0.0693 | 0.0000 | 0.0000 |
| 160527 | 0.2836 | 0.2808 | 0.0210 | 0.2315 |
| 160528 | 0.2211 | 0.0629 | -0.0241 | -0.1199 |
| 160529 | 0.1479 | -0.0851 | -0.0887 | -0.0809 |
| 160530 | -0.0272 | -0.1752 | 0.0082 | -0.1159 |
| 160531 | 0.1396 | 0.0272 | 0.2897 | 0.3688 |
| 160532 | 0.1950 | 0.2776 | -0.0012 | -0.0336 |
| 160533 | 0.0518 | 0.0169 | 0.0905 | 0.0619 |
| 160534 | 0.2074 | 0.1035 | 0.0000 | 0.0000 |
| 160535 | 0.0000 | 0.0000 | 0.4482 | 0.2668 |
| 160536 | 0.2356 | 0.4433 | 0.2568 | 0.3986 |
| 160537 | 0.0713 | 0.1740 | 0.3107 | -0.1537 |
| 160538 | 0.2885 | 0.4596 | 0.0841 | 0.1798 |
| 160539 | 0.1169 | 0.0318 | 0.0587 | -0.0694 |
| 160540 | 0.0901 | 0.0442 | 0.1164 | 0.2630 |
| 160541 | 0.1639 | 0.0945 | 0.0771 | 0.2198 |
| 160542 | 0.3539 | 0.2127 | -0.1274 | -0.1236 |
| 160543 | 0.1824 | 0.0070 | 0.0243 | -0.1101 |
| 160544 | 0.0649 | -0.0564 | -0.1348 | -0.2630 |
| 160545 | 0.1833 | -0.0821 | 0.3949 | 0.1957 |
| 160546 | 0.0954 | 0.1549 | 0.0000 | 0.0000 |
| 160547 | 0.0000 | 0.0000 | 0.0259 | 0.0637 |
| 160548 | 0.2811 | 0.0902 | -0.1240 | -0.3777 |
| 160549 | 0.2629 | -0.0154 | 0.3154 | -0.0336 |
| 160550 | 0.2824 | 0.2642 | 0.1571 | 0.1256 |
| 160551 | 0.4042 | 0.1331 | 0.0609 |  |
| 160552 | 0.2205 | 0.1021 | 0.2218 | 0.3380 |
| 160553 | 0.3101 | 0.1894 | 0.1309 | 0.0974 |
|  | 0.1476 | 0.0602 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 2 (continued)
Item Discrimination Values for Examination $3-C_{X 2}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160554 | 0.2015 | 0.1488 | 0.1768 | 0.2722 |
| 160555 | 0.3749 | 0.0000 | 0.0000 | 0.0000 |
| 160556 | 0.2813 | 0.1544 | 0.1369 | -0.0360 |
| 160557 | 0.3548 | 0.2098 | 0.2182 | 0.4002 |
| 160558 | 0.2423 | 0.1544 | 0.2661 | 0.0307 |
| 160559 | 0.2653 | 0.0617 | 0.1538 | -0.0450 |
| 160560 | 0.2133 | 0.0478 | 0.0302 | 0.0066 |
| 160561 | 0.2015 | -0.1151 | -0.1793 | 0.0000 |
| 160562 | -0.0636 | 0.0000 | 0.0000 | 0.0000 |
| 160563 | 0.1022 | -0.0729 | -0.0216 | 0.2615 |
| 160564 | 0.2586 | 0.2642 | 0.3154 | 0.2828 |
| 160565 | 0.1646 | -0.0091 | -0.0551 | 0.1171 |
| 160566 | 0.2540 | 0.1206 | 0.1138 | 0.0619 |
| 160567 | 0.0320 | 0.0440 | -0.0033 | -0.0926 |
| 160568 | 0.0281 | 0.2478 | 0.3104 | 0.0000 |
| 160569 | 0.2538 | 0.1172 | -0.0724 | -0.1946 |
| 160570 | 0.2961 | 0.1779 | 0.1588 | 0.1530 |
| 160571 | 0.1201 | -0.0012 | -0.2127 | -0.2959 |
| 160572 | 0.2302 | 0.0752 | 0.1930 | -0.1537 |
| 160573 | 0.2998 | 0.1286 | -0.0445 | 0.0225 |
| 160574 | 0.2106 | -0.0821 | -0.1348 | -0.2630 |
| 160575 | 0.0609 | 0.1861 | -0.0018 | -0.0486 |
| 160576 | 0.0652 | 0.0180 | -0.0542 | -0.0954 |
| 160577 | 0.1216 | 0.4867 | 0.4770 | 0.4420 |
| 160578 | 0.1544 | 0.2808 | 0.0000 | 0.0000 |
|  |  |  |  |  |

Table C. 3
Item Discrimination Values for Examination 3 - $C_{X 3}$

| Item | Unrestricted | +/-1.00 SD | +/-0.75SD | +/- 0.50 SD |
| :---: | :---: | :---: | :---: | :---: |
| 159995 | -0.1125 | 0.1628 | 0.3763 | 0.1530 |
| 159996 | -0.0017 | 0.4277 | 0.3027 | 0.1976 |
| 159997 | 0.3494 | 0.2042 | -0.0260 | 0.1848 |
| 159998 | 0.3372 | 0.1232 | 0.3508 | 0.2058 |
| 159999 | 0.2592 | 0.0069 | 0.0240 | 0.1122 |
| 160000 | 0.0715 | -0.1042 | -0.1651 | 0.1530 |
| 160001 | -0.0636 | 0.0000 | 0.0000 | 0.0000 |
| 160002 | 0.1160 | 0.0330 | -0.1430 | -0.5412 |
| 160003 | 0.1235 | 0.1750 | 0.2161 | 0.4686 |
| 160004 | -0.0436 | -0.0705 | -0.0794 | -0.0071 |
| 160005 | 0.4647 | 0.4288 | 0.2920 | 0.8554 |
| 160006 | 0.4144 | 0.0522 | -0.0825 | 0.0574 |
| 160007 | 0.1342 | 0.0166 | 0.0000 | -0.1848 |
| 160008 | 0.1732 | -0.0696 | -0.0809 | -0.1564 |
| 160009 | 0.0014 | -0.2772 | -0.3506 | 0.0000 |
| 160010 | 0.1508 | 0.0756 | -0.3121 | -0.4086 |
| 160011 | 0.2503 | -0.1523 | -0.1861 | -0.0855 |
| 160012 | -0.0345 | 0.2269 | -0.1559 | -0.1656 |
| 160013 | 0.2169 | 0.0786 | 0.1117 | 0.2058 |
| 160014 | 0.3356 | 0.0365 | 0.0674 | -0.0128 |
| 160015 | 0.1391 | 0.1850 | 0.4143 | 0.1848 |
| 160016 | 0.1500 | 0.2453 | 0.0470 | 0.1275 |
| 160017 | 0.2353 | -0.0055 | -0.1861 | -0.0855 |
| 160018 | 0.2385 | 0.0322 | -0.1101 | 0.2933 |
| 160019 | 0.1183 | -0.0236 | -0.1853 | -0.1782 |
| 160020 | 0.1056 | -0.0034 | 0.0069 | 0.3869 |
| 160021 | 0.1994 | -0.0321 | 0.1141 | 0.3300 |
| 160022 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160023 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160024 | -0.0251 | -0.0383 | -0.0442 | -0.2871 |
| 160025 | 0.1141 | 0.0021 | 0.0159 | 0.1497 |

Table C. 3 (continued)
Item Discrimination Values for Examination $3-C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | -0.1101 | -0.1656 |
| 160026 | 0.1059 | -0.0909 | 0.3345 | 0.4417 |
| 160027 | 0.0230 | 0.2555 | 0.2820 | 0.5418 |
| 160028 | 0.2308 | 0.0747 | -0.0660 | -0.1056 |
| 160039 | 0.3289 | 0.0784 | 0.3852 | 0.4782 |
| 160031 | 0.2980 | 0.2654 | -0.0393 | -0.2232 |
| 160032 | 0.2154 | -0.1445 | 0.0449 | 0.0574 |
| 160033 | 0.4353 | 0.1441 | 0.0955 | -0.4026 |
| 160034 | -0.0930 | 0.0574 | 0.0747 | 0.0396 |
| 160035 | 0.4090 | 0.2654 | 0.0841 | -0.6138 |
| 160036 | 0.2824 | 0.0524 | 0.1430 | 0.5632 |
| 160037 | 0.0739 | -0.0158 | 0.0000 | 0.0000 |
| 160038 | -0.0380 | 0.0000 | -0.1073 | -0.4086 |
| 160039 | 0.0829 | -0.0873 | 0.1503 | 0.3869 |
| 160040 | 0.1151 | 0.1083 | 0.0240 | -0.2678 |
| 160041 | -0.0036 | -0.0972 | 0.4644 | 0.5478 |
| 160042 | 0.4053 | 0.4493 | -0.1101 | -0.0442 |
| 160043 | 0.1553 | 0.1020 | -0.1995 | -0.3375 |
| 160044 | 0.1101 | -0.1615 | 0.2336 | 0.3849 |
| 160045 | 0.1223 | 0.3929 | -0.0963 | 0.0247 |
| 160046 | 0.1300 | -0.0825 | -0.3490 | -0.1848 |
| 160047 | 0.2300 | -0.1413 | -0.2820 | -0.1122 |
| 160048 | 0.1708 | -0.3253 | 0.4607 | 0.3065 |
| 160049 | 0.2581 | 0.0981 | 0.0189 | -0.1656 |
| 160050 | 0.0586 | -0.1618 | 0.5679 |  |
| 160051 | 0.2836 | 0.3344 | 0.0963 | -0.0247 |
| 160052 | 0.3579 | 0.0321 | -0.2161 | 0.0340 |
| 160053 | 0.1241 | -0.2708 | 0.2476 | 0.0396 |
| 160054 | 0.2348 | 0.0505 | 0.2065 | 0.2574 |
| 160055 | 0.1140 | 0.2718 | 0.0240 | -0.0071 |
| 160056 | 0.2920 | -0.1413 | 0.0757 |  |
|  | 0.2966 | 0.3169 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 3 (continued)
Item Discrimination Values for Examination 3 - $C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160057 | 0.3838 | -0.0214 | -0.2625 | 0.4633 |
| 160058 | 0.0012 | -0.1363 | 0.0159 | -0.4280 |
| 160059 | 0.2883 | 0.2571 | 0.0000 | 0.0000 |
| 160060 | 0.0000 | 0.0000 | 0.4970 | 0.1152 |
| 160061 | 0.2186 | 0.3752 | -0.0275 | 0.0000 |
| 160062 | 0.1056 | -0.0298 | -0.1559 | -0.1656 |
| 160063 | 0.0365 | -0.1264 | 0.2359 | 0.4080 |
| 160064 | 0.2481 | -0.0368 | -0.3714 | -0.2871 |
| 160065 | -0.0272 | -0.2933 | -0.0540 | -0.5418 |
| 160066 | -0.1009 | -0.1623 | 0.3164 | 0.0247 |
| 160067 | 0.0136 | 0.2337 | 0.0000 | 0.0000 |
| 160068 | 0.0390 | 0.0000 | 0.1741 | -0.0071 |
| 160069 | 0.2224 | 0.3586 | 0.1192 | -0.0658 |
| 160070 | 0.0423 | 0.0863 | 0.0240 | 0.3849 |
| 160071 | 0.1943 | 0.0069 | -0.2383 | -0.4776 |
| 160072 | 0.2186 | 0.0752 | 0.1011 | -0.2574 |
| 160073 | 0.2013 | 0.0522 | 0.2640 | 0.3300 |
| 160074 | 0.3289 | 0.1832 | 0.0000 | 0.0000 |
| 160075 | 0.0000 | 0.0000 | 0.5319 | 0.7490 |
| 160076 | 0.1443 | 0.4053 | 0.0159 | 0.1497 |
| 160077 | 0.0059 | 0.2571 | -0.2469 |  |
| 160078 | 0.1822 | 0.3455 | 0.0275 | -0.0128 |
| 160079 | 0.3248 | 0.1042 | 0.2959 | -0.0442 |
| 160080 | 0.1056 | 0.4407 | 0.0821 | -0.3375 |
| 160081 | 0.1882 | -0.1942 | -0.1125 | 0.0829 |
| 160082 | 0.1301 | -0.0684 | 0.0606 | -0.0574 |
| 160083 | 0.1376 | -0.2921 | -0.1189 | 0.7655 |
| 160084 | 0.2800 | 0.4020 | 0.5509 | 0.0247 |
| 160085 | 0.1580 | 0.1217 | 0.1651 | -0.1122 |
| 160086 | 0.1300 | -0.2251 | -0.1516 | 0.3849 |
| 160087 | 0.1955 | 0.3720 | 0.2048 |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 3 (continued)
Item Discrimination Values for Examination 3 - $C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 160088 | -0.0804 | -0.1852 | -0.2335 | 0.0000 |
| 160089 | 0.2772 | 0.0841 | -0.0449 | -0.3634 |
| 160090 | -0.0300 | 0.0193 | 0.0347 | 0.0247 |
| 160526 | 0.2836 | 0.3344 | 0.4402 | -0.0442 |
| 160527 | 0.2211 | 0.1623 | 0.1990 | -0.1497 |
| 160528 | 0.1479 | -0.1289 | 0.0275 | 0.1988 |
| 160529 | -0.0272 | -0.3069 | -0.3891 | 0.0000 |
| 160530 | 0.1396 | -0.0418 | -0.0359 | -0.0829 |
| 160531 | 0.1950 | -0.1032 | -0.3144 | -0.5412 |
| 160532 | 0.0518 | 0.0000 | 0.0000 | 0.0000 |
| 160533 | 0.2074 | 0.0490 | 0.0732 | 0.1152 |
| 160534 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160535 | 0.2356 | 0.3365 | 0.2935 | 0.0000 |
| 160536 | 0.0713 | -0.3692 | -0.2279 | -0.6184 |
| 160537 | 0.2885 | 0.3510 | 0.1759 | -0.4782 |
| 160538 | 0.1169 | -0.0386 | 0.1112 | 0.3849 |
| 160539 | 0.0901 | 0.2037 | 0.2959 | 0.2933 |
| 160540 | 0.1639 | -0.0078 | 0.1466 | 0.4686 |
| 160541 | 0.3539 | 0.2877 | 0.2406 | 0.4752 |
| 160542 | 0.1824 | 0.0365 | -0.1261 | 0.1497 |
| 160543 | 0.0649 | 0.1083 | 0.1503 | -0.0855 |
| 160544 | 0.1833 | 0.4024 | 0.5239 | 0.0000 |
| 160545 | 0.0954 | -0.1586 | -0.1965 | -0.4280 |
| 160546 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160547 | 0.2811 | -0.0321 | -0.0260 | 0.1848 |
| 160548 | 0.2629 | 0.2475 | 0.3121 | 0.4086 |
| 160549 | 0.2824 | 0.1973 | 0.2658 | -0.4280 |
| 160550 | 0.4042 | 0.1577 | 0.0636 | 0.0829 |
| 160551 | 0.2205 | 0.1020 | 0.1444 | -0.1782 |
| 160552 | 0.3101 | 0.1958 | 0.2700 | 0.7770 |
| 160553 | 0.1476 | 0.1413 | 0.0240 | 0.0247 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 3 (continued)
Item Discrimination Values for Examination 3 - $C_{X 3}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160554 | 0.2015 | 0.0154 | 0.0275 | 0.1988 |
| 160555 | 0.3749 | 0.6888 | 0.5239 | 0.0000 |
| 160556 | 0.2813 | 0.0330 | 0.069 | -0.0855 |
| 160557 | 0.3548 | 0.2671 | -0.006 | -0.0071 |
| 160558 | 0.2423 | -0.1068 | -0.1261 | 0.1152 |
| 160559 | 0.2653 | 0.1958 | 0.0732 | 0.1988 |
| 160560 | 0.2133 | 0.1973 | 0.5319 | 0.0000 |
| 160561 | 0.2015 | 0.5493 | 0.0000 | 0.0000 |
| 160562 | -0.0636 | 0.0000 | 0.0000 | -0.0330 |
| 160563 | 0.1022 | -0.0368 | -0.0275 | 0.3849 |
| 160564 | 0.2586 | 0.2344 | 0.4924 | 0.3203 |
| 160565 | 0.1646 | 0.0154 | 0.0275 | 0.0247 |
| 160566 | 0.2540 | 0.2864 | -0.0039 | 0.1988 |
| 160567 | 0.0320 | -0.1586 | -0.0183 | -0.1656 |
| 160568 | 0.0281 | 0.0106 | 0.0189 | 0.1152 |
| 160569 | 0.2538 | 0.3127 | 0.0347 | 0.1122 |
| 160570 | 0.2961 | 0.0400 | 0.0240 | 0.0658 |
| 160571 | 0.1201 | -0.0069 | -0.0240 | -0.1640 |
| 160572 | 0.2302 | 0.2200 | 0.1117 | 0.6138 |
| 160573 | 0.2998 | -0.1434 | 0.1430 | 0.0000 |
| 160574 | 0.2106 | 0.5493 | 0.0000 | 0.1152 |
| 160575 | 0.0609 | 0.0193 | 0.0347 | -0.3300 |
| 160576 | 0.0652 | 0.1441 | 0.1742 | -0.3379 |
| 160577 | 0.1216 | 0.0810 | -0.1853 | -0.0442 |
| 160578 | 0.1544 | 0.0596 | 0.0821 |  |
|  |  |  |  |  |

Table C. 4
Item Discrimination Values for Examination 3 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  | 0.0609 | 0.1748 |
| 159995 | -0.1125 | -0.1515 | -0.3230 | -0.2179 |
| 159996 | -0.0017 | -0.1534 | 0.1108 | 0.2144 |
| 159997 | 0.3494 | 0.2456 | 0.3618 | 0.2236 |
| 159999 | 0.3372 | 0.2655 | 0.1411 | 0.1371 |
| 160000 | 0.2592 | 0.1381 | -0.2648 | -0.3753 |
| 160001 | 0.0715 | -0.0351 | -0.2494 | 0.0000 |
| 160002 | -0.0636 | -0.1278 | -0.0161 | -0.1714 |
| 160003 | 0.1160 | 0.0220 | -0.1134 | -0.0947 |
| 160004 | 0.1235 | 0.0614 | 0.0882 | 0.2670 |
| 160005 | -0.0436 | -0.0684 | 0.2259 | 0.2488 |
| 160006 | 0.4647 | 0.3704 | 0.0537 | 0.1261 |
| 160007 | 0.4144 | 0.3406 | 0.1535 | -0.1148 |
| 160008 | 0.1342 | 0.0433 | 0.1411 | 0.1371 |
| 160009 | 0.1732 | 0.2487 | -0.0930 | 0.2236 |
| 160010 | 0.0014 | 0.1562 | 0.0572 | 0.3801 |
| 160011 | 0.1508 | 0.1332 | -0.0686 | 0.0928 |
| 160012 | 0.2503 | 0.2372 | 0.1086 | 0.4677 |
| 160013 | -0.0345 | -0.0837 | -0.0911 | -0.1746 |
| 160014 | 0.2169 | 0.1249 | 0.1688 | 0.4516 |
| 160015 | 0.3356 | 0.1489 | 0.1971 | 0.1958 |
| 160016 | 0.1391 | 0.0904 | 0.0286 | 0.1824 |
| 160017 | 0.1500 | 0.0395 | 0.3292 | 0.2018 |
| 160018 | 0.2353 | 0.3406 | 0.0699 | 0.2542 |
| 160019 | 0.2385 | 0.0519 | -0.1514 |  |
| 160020 | 0.1183 | 0.2867 | -0.2658 | -0.2191 |
| 160021 | 0.1056 | 0.0636 | 0.0671 | 0.0000 |
| 160022 | 0.1994 | 0.0351 | 0.00000 | 0.3094 |
| 160023 | 0.0000 | 0.0000 | -0000 |  |
| 160024 | 0.0000 | 0.0865 | 0.1410 |  |
| 160025 | 0.1141 |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 4 (continued)
Item Discrimination Values for Examination 3 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160026 | 0.1059 | 0.0673 | -0.2494 | 0.0000 |
| 160027 | 0.0230 | -0.1348 | 0.0887 | 0.3801 |
| 160028 | 0.2308 | 0.1164 | 0.2614 | 0.0947 |
| 160029 | 0.3289 | 0.2016 | 0.1515 | 0.1398 |
| 160030 | 0.2980 | 0.1099 | 0.0856 | 0.0522 |
| 160031 | 0.2154 | 0.2059 | -0.0889 | -0.3438 |
| 160032 | 0.4353 | 0.2836 | 0.0733 | -0.0510 |
| 160033 | -0.0930 | -0.2132 | -0.2163 | -0.2297 |
| 160034 | 0.4090 | 0.2511 | -0.0684 | -0.0419 |
| 160035 | 0.2824 | 0.1248 | -0.0384 | -0.0949 |
| 160036 | 0.0739 | 0.2149 | 0.3432 | 0.2346 |
| 160037 | -0.0380 | -0.0740 | -0.1439 | -0.2542 |
| 160038 | 0.0829 | 0.1826 | 0.0144 | -0.0153 |
| 160039 | 0.1151 | 0.0277 | 0.0671 | 0.0643 |
| 160040 | -0.0036 | -0.0875 | 0.2123 | 0.0153 |
| 160041 | 0.4053 | 0.2247 | 0.3350 | 0.2120 |
| 160042 | 0.1553 | 0.0190 | 0.2242 | 0.0203 |
| 160043 | 0.1101 | 0.2464 | 0.1201 | 0.1161 |
| 160044 | 0.1223 | -0.0710 | -0.0701 | -0.0966 |
| 160045 | 0.1300 | 0.1661 | -0.0686 | 0.0928 |
| 160046 | 0.2300 | 0.1515 | 0.1265 | 0.2166 |
| 160047 | 0.1708 | 0.3145 | -0.1248 | -0.1824 |
| 160048 | 0.2581 | 0.0806 | 0.1528 | -0.1175 |
| 160049 | 0.0586 | 0.0298 | 0.2259 | 0.2644 |
| 160050 | 0.2836 | 0.0000 | 0.0000 | 0.0000 |
| 160051 | 0.3579 | 0.2550 | 0.0066 | 0.2909 |
| 160052 | 0.1241 | 0.0965 | 0.2025 | 0.0015 |
| 160053 | 0.2348 | 0.1863 | 0.3718 | 0.2652 |
| 160054 | 0.1140 | 0.0117 | 0.2894 | 0.1072 |
| 160055 | 0.2920 | 0.3703 | 0.1343 | 0.2475 |
| 160056 | 0.2966 | 0.0555 | 0.0965 | 0.0928 |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 4 (continued)
Item Discrimination Values for Examination 3 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160057 | 0.3838 | 0.3438 | 0.4976 | 0.2236 |
| 160058 | 0.0012 | 0.1526 | 0.0887 | 0.0682 |
| 160059 | 0.2883 | 0.2140 | 0.0856 | 0.0522 |
| 160060 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160061 | 0.2186 | 0.0521 | 0.1087 | 0.0671 |
| 160062 | 0.1056 | 0.1597 | 0.2102 | 0.2652 |
| 160063 | 0.0365 | 0.0753 | 0.2254 | 0.3032 |
| 160064 | 0.2481 | 0.0790 | 0.0340 | -0.3226 |
| 160065 | -0.0272 | 0.0335 | 0.1914 | 0.3305 |
| 160066 | -0.1009 | -0.1216 | -0.0066 | -0.1134 |
| 160067 | 0.0136 | -0.0820 | -0.0413 | -0.1539 |
| 160068 | 0.0390 | 0.0875 | 0.1727 | 0.2236 |
| 160069 | 0.2224 | 0.0820 | 0.0344 | 0.2220 |
| 160070 | 0.0423 | -0.0521 | 0.0596 | 0.0015 |
| 160071 | 0.1943 | 0.1923 | 0.0853 | 0.0308 |
| 160072 | 0.2186 | 0.2338 | -0.0384 | -0.0949 |
| 160073 | 0.2013 | 0.2429 | 0.2322 | 0.5543 |
| 160074 | 0.3289 | 0.2511 | -0.0905 | -0.0783 |
| 160075 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160076 | 0.1443 | -0.0174 | 0.1353 | 0.1181 |
| 160077 | 0.0059 | 0.0351 | -0.1347 | -0.1510 |
| 160078 | 0.1822 | 0.0318 | 0.0701 | -0.0235 |
| 160079 | 0.3248 | 0.1978 | 0.3073 | 0.1641 |
| 160080 | 0.1056 | -0.1002 | 0.1101 | 0.0792 |
| 160081 | 0.1882 | 0.2174 | 0.1201 | 0.1439 |
| 160082 | 0.1301 | 0.0375 | 0.0733 | -0.2166 |
| 160083 | 0.1376 | 0.1450 | 0.0991 | -0.2166 |
| 160084 | 0.2800 | 0.1803 | -0.0930 | -0.1945 |
| 160085 | 0.1580 | 0.1909 | 0.1857 | 0.1958 |
| 160086 | 0.1300 | 0.1663 | 0.0428 | -0.1748 |
| 160087 | 0.1955 | 0.1131 | 0.1093 | 0.4669 |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 4 (continued)
Item Discrimination Values for Examination 3 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 160088 | -0.0804 | -0.1547 | -0.3549 | -0.3690 |
| 160089 | 0.2772 | 0.0912 | 0.0198 | -0.0402 |
| 160090 | -0.0300 | -0.1330 | -0.1315 | -0.0221 |
| 160526 | 0.2836 | 0.0000 | 0.0000 | 0.0000 |
| 160527 | 0.2211 | 0.0573 | 0.0477 | -0.1991 |
| 160528 | 0.1479 | 0.1909 | 0.2109 | 0.0098 |
| 160529 | -0.0272 | 0.0708 | -0.1234 | -0.0179 |
| 160530 | 0.1396 | 0.0847 | -0.1255 | 0.0827 |
| 160531 | 0.1950 | 0.2541 | -0.0250 | 0.1398 |
| 160532 | 0.0518 | 0.1144 | 0.2254 | 0.3032 |
| 160533 | 0.2074 | 0.1546 | 0.0586 | 0.0354 |
| 160534 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160535 | 0.2356 | 0.0222 | -0.0399 | 0.0654 |
| 160536 | 0.0713 | 0.0388 | -0.1038 | -0.4054 |
| 160537 | 0.2885 | -0.0677 | 0.2459 | 0.3078 |
| 160538 | 0.1169 | -0.0576 | -0.0173 | 0.0403 |
| 160539 | 0.0901 | -0.0641 | -0.2021 | -0.1174 |
| 160540 | 0.1639 | 0.1748 | -0.0576 | -0.1174 |
| 160541 | 0.3539 | 0.1281 | -0.0250 | 0.1398 |
| 160542 | 0.1824 | 0.1611 | 0.2504 | 0.0730 |
| 160543 | 0.0649 | -0.1348 | -0.1083 | -0.2448 |
| 160544 | 0.1833 | 0.0336 | 0.0671 | 0.0643 |
| 160545 | 0.0954 | -0.0515 | -0.1966 | -0.3338 |
| 160546 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160547 | 0.2811 | 0.3158 | 0.2413 | 0.2642 |
| 160548 | 0.2629 | 0.1228 | 0.2683 | 0.1235 |
| 160549 | 0.2824 | 0.1144 | 0.2254 | 0.3032 |
| 160550 | 0.4042 | 0.2292 | 0.0767 | -0.0557 |
| 160551 | 0.2205 | 0.1160 | 0.0025 | -0.0751 |
| 160552 | 0.3101 | 0.1704 | -0.2494 | 0.0000 |
| 160553 | 0.1476 | 0.1053 | 0.2494 | 0.0792 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 4 (continued)
Item Discrimination Values for Examination 3 - $C_{X 4}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160554 | 0.2015 | 0.2220 | 0.0000 | 0.0000 |
| 160555 | 0.3749 | 0.0000 | 0.0000 | 0.0000 |
| 160556 | 0.2813 | 0.2868 | 0.4625 | 0.2346 |
| 160557 | 0.3548 | 0.2664 | -0.0551 | -0.1370 |
| 160558 | 0.2423 | 0.2467 | 0.0384 | 0.1569 |
| 160559 | 0.2653 | 0.1597 | 0.1556 | 0.1371 |
| 160560 | 0.2133 | 0.0594 | 0.1201 | 0.2236 |
| 160561 | 0.2015 | 0.0067 | 0.0144 | -0.0153 |
| 160562 | -0.0636 | -0.1278 | -0.2494 | 0.0000 |
| 160563 | 0.1022 | 0.2429 | 0.0991 | -0.0095 |
| 160564 | 0.2586 | 0.1070 | 0.3309 | 0.0000 |
| 160565 | 0.1646 | 0.1704 | 0.1344 | 0.1503 |
| 160566 | 0.2540 | 0.0555 | 0.1344 | 0.1503 |
| 160567 | 0.0320 | 0.1434 | 0.3197 | 0.4342 |
| 160568 | 0.0281 | -0.2086 | 0.0000 | 0.0000 |
| 160569 | 0.2538 | 0.0784 | 0.2145 | 0.2599 |
| 160570 | 0.2961 | 0.1923 | 0.2856 | 0.1371 |
| 160571 | 0.1201 | 0.1228 | 0.1926 | 0.2024 |
| 160572 | 0.2302 | -0.0027 | 0.1556 | 0.1371 |
| 160573 | 0.2998 | 0.2811 | 0.3786 | 0.3438 |
| 160574 | 0.2106 | 0.0336 | 0.0671 | 0.0643 |
| 160575 | 0.0609 | 0.0119 | 0.0671 | 0.0643 |
| 160576 | 0.0652 | -0.1216 | -0.1151 | -0.0274 |
| 160577 | 0.1216 | 0.0708 | -0.0930 | 0.2236 |
| 160578 | 0.1544 | 0.0000 | 0.0000 | 0.0000 |
|  |  |  |  |  |

Table C. 5
Item Discrimination Values for Examination 3 - $C_{X 5}$

|  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Item | Unrestricted | $+-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 159995 | -0.1125 | 0.0526 | 0.0388 | 0.0624 |
| 159996 | -0.0017 | 0.3743 | 0.2466 | 0.7057 |
| 159997 | 0.3494 | 0.2443 | 0.5114 | 0.3119 |
| 159998 | 0.3372 | 0.2454 | 0.3812 | 0.2205 |
| 159999 | 0.2592 | -0.1505 | -0.6117 | -0.1953 |
| 160000 | 0.0715 | 0.1734 | 0.1165 | 0.1871 |
| 160001 | -0.0636 | 0.0000 | 0.0000 | 0.0000 |
| 160002 | 0.1160 | -0.1681 | 0.2018 | 0.8090 |
| 160003 | 0.1235 | 0.4983 | 0.2202 | 0.5300 |
| 160004 | -0.0436 | -0.1832 | -0.5157 | -0.5300 |
| 160005 | 0.4647 | 0.8629 | 0.7721 | 0.5300 |
| 160006 | 0.4144 | 0.0681 | 0.2719 | -0.2205 |
| 160007 | 0.1342 | -0.1206 | -0.1905 | 0.1871 |
| 160008 | 0.1732 | -0.2353 | -0.2202 | -0.5300 |
| 160009 | 0.0014 | 0.0000 | 0.0000 | 0.0000 |
| 160010 | 0.1508 | 0.2096 | 0.7193 | 0.8090 |
| 160011 | 0.2503 | -0.2308 | 0.0147 | -0.1953 |
| 160012 | -0.0345 | 0.3208 | 0.7193 | 0.8090 |
| 160013 | 0.2169 | -0.0128 | -0.4260 | -0.3627 |
| 160014 | 0.3356 | -0.3096 | -0.3710 | -0.4411 |
| 160015 | 0.1391 | 0.1206 | -0.1942 | -0.3119 |
| 160016 | 0.1500 | 0.1084 | -0.0301 | -0.2205 |
| 160017 | 0.2353 | 0.1538 | 0.1103 | 0.3529 |
| 160018 | 0.2385 | 0.2539 | 0.3495 | -0.0882 |
| 160019 | 0.1183 | 0.2539 | 0.4272 | 0.6862 |
| 160020 | 0.1056 | 0.0984 | -0.2466 | -0.7057 |
| 160021 | 0.1994 | -0.0201 | -0.2908 | -0.3088 |
| 160022 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160023 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160024 | -0.0251 | -0.2722 | 0.0000 | 0.0000 |
| 160025 | 0.1141 | -0.0879 | -0.2466 | -0.7057 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 5 (continued)
Item Discrimination Values for Examination 3 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160026 | 0.1059 | -0.1969 | 0.0000 | 0.0000 |
| 160027 | 0.0230 | 0.1796 | -0.2908 | -0.1953 |
| 160028 | 0.2308 | 0.1502 | 0.1905 | 0.4852 |
| 160029 | 0.3289 | 0.0793 | -0.4709 | -0.3529 |
| 160030 | 0.2980 | 0.1889 | 0.1103 | -0.3119 |
| 160031 | 0.2154 | -0.2136 | 0.5315 | 0.7057 |
| 160032 | 0.4353 | 0.2919 | -0.4512 | -0.5734 |
| 160033 | -0.0930 | -0.5759 | 0.1304 | 0.6862 |
| 160034 | 0.4090 | 0.3746 | -0.0673 | 0.3069 |
| 160035 | 0.2824 | -0.3331 | 0.3812 | 0.2205 |
| 160036 | 0.0739 | 0.4877 | 0.0000 | 0.0000 |
| 160037 | -0.0380 | 0.0000 | 0.0000 | 0.0000 |
| 160038 | 0.0829 | -0.3475 | -0.2466 | -0.0279 |
| 160039 | 0.1151 | 0.0984 | -0.4260 | -0.3627 |
| 160040 | -0.0036 | -0.4156 | 0.6919 | 0.3627 |
| 160041 | 0.4053 | 0.6563 | 0.2018 | 0.8090 |
| 160042 | 0.1553 | 0.3764 | -0.4550 | 0.0000 |
| 160043 | 0.1101 | -0.3465 | 0.6602 | 0.4411 |
| 160044 | 0.1223 | 0.6155 | -0.1028 | -0.3627 |
| 160045 | 0.1300 | -0.1241 | 0.1942 | 0.3119 |
| 160046 | 0.2300 | -0.1206 | -0.3510 | -0.0624 |
| 160047 | 0.1708 | -0.2030 | 0.1103 | -0.3119 |
| 160048 | 0.2581 | 0.2408 | 0.0000 | 0.0000 |
| 160049 | 0.0586 | -0.1969 | -0.0673 | -0.4411 |
| 160050 | 0.2836 | 0.2096 | -0.3812 | -0.2205 |
| 160051 | 0.3579 | -0.2931 | -0.1028 | -0.3627 |
| 160052 | 0.1241 | -0.2784 | -0.6117 | -0.1953 |
| 160053 | 0.2348 | -0.1940 | 0.1165 | 0.1871 |
| 160054 | 0.1140 | 0.2855 | 0.2719 | -0.2205 |
| 160055 | 0.2920 | 0.0681 | 0.5606 | 0.4852 |
| 160056 | 0.2966 | 0.1973 |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 5 (continued)
Item Discrimination Values for Examination 3 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  |  |
| 160057 | 0.3838 | 0.3277 | -0.1165 | -0.1871 |
| 160058 | 0.0012 | 0.0000 | 0.0000 | 0.0000 |
| 160059 | 0.2883 | 0.2842 | 0.8296 | 0.8822 |
| 160060 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160061 | 0.2186 | 0.1978 | -0.2106 | -0.1764 |
| 160062 | 0.1056 | 0.4054 | 0.3670 | 0.3069 |
| 160063 | 0.0365 | -0.1969 | 0.0000 | 0.0000 |
| 160064 | 0.2481 | -0.0681 | -0.3495 | -0.5614 |
| 160065 | -0.0272 | -0.2722 | 0.0000 | 0.0000 |
| 160066 | -0.1009 | -0.5016 | -0.5916 | -0.4366 |
| 160067 | 0.0136 | 0.1502 | 0.2018 | -0.0441 |
| 160068 | 0.0390 | 0.0000 | 0.0000 | 0.0000 |
| 160069 | 0.2224 | 0.4092 | 0.3495 | 0.5614 |
| 160070 | 0.0423 | -0.1797 | -0.2202 | -0.5300 |
| 160071 | 0.1943 | 0.0549 | 0.0224 | -0.3088 |
| 160072 | 0.2186 | -0.0635 | 0.7193 | 0.8090 |
| 160073 | 0.2013 | -0.2855 | -0.1165 | -0.1871 |
| 160074 | 0.3289 | 0.2030 | 0.0388 | 0.0624 |
| 160075 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160076 | 0.1443 | 0.3208 | 0.1121 | -0.1764 |
| 160077 | 0.0059 | 0.2842 | 0.3510 | 0.0624 |
| 160078 | 0.1822 | 0.1502 | 0.2018 | 0.8090 |
| 160079 | 0.3248 | -0.0929 | 0.0388 | 0.0624 |
| 160080 | 0.1056 | 0.3764 | 0.2018 | 0.8090 |
| 160081 | 0.1882 | -0.0403 | 0.3670 | 0.3069 |
| 160082 | 0.1301 | -0.2094 | 0.0224 | -0.3088 |
| 160083 | 0.1376 | -0.2919 | 0.0224 | -0.3088 |
| 160084 | 0.2800 | 0.4505 | -0.0301 | -0.2205 |
| 160085 | 0.1580 | -0.1241 | 0.0147 | -0.1953 |
| 160086 | 0.1300 | -0.4149 | -0.5826 | -0.3088 |
| 160087 | 0.1955 | 0.6155 | 0.6602 | 0.4411 |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 5 (continued)
Item Discrimination Values for Examination 3 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
| 160088 | -0.0804 | 0.0000 | 0.0000 | 0.0000 |
| 160089 | 0.2772 | -0.1734 | -0.1942 | 0.3529 |
| 160090 | -0.0300 | -0.1241 | -0.1028 | -0.3627 |
| 160526 | 0.2836 | 0.2096 | -0.0673 | 0.3069 |
| 160527 | 0.2211 | 0.0879 | 0.2466 | 0.0279 |
| 160528 | 0.1479 | 0.0290 | -0.2202 | -0.5300 |
| 160529 | -0.0272 | 0.0000 | 0.0000 | 0.0000 |
| 160530 | 0.1396 | -0.3498 | -0.3710 | -0.4411 |
| 160531 | 0.1950 | -0.1681 | 0.2018 | 0.8090 |
| 160532 | 0.0518 | 0.0000 | 0.0000 | 0.0000 |
| 160533 | 0.2074 | -0.0684 | -0.1028 | -0.3627 |
| 160534 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160535 | 0.2356 | 0.0000 | 0.0000 | 0.0000 |
| 160536 | 0.0713 | -0.6218 | -0.6951 | 0.0000 |
| 160537 | 0.2885 | 0.0279 | 0.1165 | 0.8822 |
| 160538 | 0.1169 | 0.0549 | -0.0673 | -0.4411 |
| 160539 | 0.0901 | 0.1331 | 0.0388 | 0.0624 |
| 160540 | 0.1639 | -0.0929 | -0.4312 | -0.8822 |
| 160541 | 0.3539 | 0.6161 | 0.1570 | 0.5734 |
| 160542 | 0.1824 | 0.2842 | 0.2707 | 0.6175 |
| 160543 | 0.0649 | -0.2308 | -0.5157 | -0.5300 |
| 160544 | 0.1833 | 0.4054 | 0.3670 | 0.3069 |
| 160545 | 0.0954 | -0.4021 | 0.0000 | 0.0000 |
| 160546 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| 160547 | 0.2811 | -0.1070 | 0.1121 | -0.1764 |
| 160548 | 0.2629 | 0.3475 | 0.0000 | 0.0000 |
| 160549 | 0.2824 | -0.0879 | 0.3670 | 0.3069 |
| 160550 | 0.4042 | 0.3498 | 0.3710 | 0.4411 |
| 160551 | 0.2205 | -0.3244 | -0.7721 | -0.5300 |
| 160552 | 0.3101 | 0.2931 | -0.0501 | -0.5614 |
| 160553 | 0.1476 | 0.5016 | 0.6719 | 0.5614 |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |

Table C. 5 (continued)
Item Discrimination Values for Examination 3 - $C_{X 5}$

|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Item | Unrestricted | $+/-1.00 S D$ | $+/-0.75 S D$ | $+/-0.50 S D$ |
|  |  |  |  |  |
|  |  |  |  | -0.5300 |
| 160554 | 0.2015 | 0.0290 | 0.8296 | 0.8822 |
| 160555 | 0.3749 | 0.7657 | 0.0301 | 0.2205 |
| 160557 | 0.2813 | 0.1538 | 0.7521 | 0.6862 |
| 160558 | 0.3548 | 0.3680 | 0.1321 | -0.0279 |
| 160559 | 0.2423 | -0.1832 | 0.4709 | 0.3529 |
| 160560 | 0.2653 | 0.3407 | 0.1121 | -0.1764 |
| 160561 | 0.2133 | 0.3208 | 0.7193 | 0.8090 |
| 160562 | 0.2015 | 0.6313 | 0.0000 | 0.0000 |
| 160563 | -0.0636 | 0.0000 | 0.1103 | -0.3119 |
| 160564 | 0.1022 | -0.0032 | 0.2707 | -0.0624 |
| 160565 | 0.2586 | 0.2842 | -0.1028 | -0.3627 |
| 160566 | 0.1646 | 0.1043 | 0.3812 | 0.2205 |
| 160567 | 0.2540 | 0.2931 | -0.2202 | -0.5300 |
| 160568 | 0.0320 | 0.0290 | 0.0000 | 0.0000 |
| 160569 | 0.0281 | -0.1969 | 0.0301 | 0.2205 |
| 160570 | 0.2538 | 0.3407 | 0.4312 | 0.1871 |
| 160571 | 0.2961 | 0.2030 | -0.1121 | 0.1764 |
| 160572 | 0.1201 | -0.1026 | 0.4709 | 0.3529 |
| 160573 | 0.2302 | 0.1104 | -0.2018 | -0.8090 |
| 160574 | 0.2998 | 0.2094 | 0.7193 | 0.8090 |
| 160575 | 0.2106 | 0.6313 | -0.5157 | -0.5300 |
| 160576 | 0.0609 | -0.0684 | 0.2908 | 0.3088 |
| 160577 | 0.0652 | 0.0201 | 0.1905 | 0.4852 |
| 160578 | 0.1216 | -0.1084 | -0.4550 | 0.0000 |
|  | 0.1544 | -0.1216 |  |  |
|  |  |  |  |  |

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